

Supplemental Appendix for
Inequality and Income Dynamics in Germany

by

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Appendices

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A Institutional and Macroeconomic Background: Details

In this section we extend the discussion of subsection 2.1 on relevant institutions and the macroeconomic situation in Germany for the period 1993 to 2018. We complement further elaborations by figures of institutional and trend patterns over this period.

A.1 Institutions

The Personal Income Tax. Germany applies a comprehensive income tax on income from all sources. The German tax law distinguishes seven different types of income: (i) income from agriculture and forestry, (ii) (non-corporate) business income (this includes dividends and capital gains from closely held corporations, i.e. with an ownership share of at least 1%), (iii) entrepreneurial income, (iv) salaries and wages from employment, (v) investment income (i.e. interest payments and dividends from “normal” stock holdings), (vi) rental income, and (vii) “miscellaneous income” (including, for example, taxable (private) pensions, annuities and capital gains).¹ For each type of income, all expenses that are necessary to obtain, maintain or preserve the income from a given source are deductible. The same holds for education costs, child care costs and donations to charity.

In contrast to most other countries, which use a bracket system with constant marginal tax rates within a bracket, Germany uses a formula (which is quadratic in income) to compute the tax liability. As a consequence, marginal tax rates increase linearly in taxable income from 14% up to 42% (for taxable income above 52,151 Euro in 2008). At the very top, an additional tax bracket with a marginal tax rate of 45% was introduced in 2007 for taxable income above 250,000 Euro.²

Between 2000 and 2005, a major reform of the German personal income tax took place. The basic tax allowance was increased in several steps from 6,902 Euro in 2000 to 7,664 Euro (2004-2008). The lowest marginal tax rate decreased from 22.9% in 2000 to 15% (2005-2008) and 14% (since 2009) – see Figure A.1 (A). The top marginal tax rate was reduced from 51% in 2000 to 42% in 2005. The threshold for application of the top marginal tax rate was reduced from 58,643 Euro in 2000 to 52,151 Euro in 2004. In 2007, an additional tax bracket (for taxable income above 250,000

¹The following types of income are tax exempt: payments from health insurance, accident insurance and insurance for disability and old age, welfare benefits and scholarships.

²The reasoning behind using such a formula instead of tax brackets was “to avoid bunching at kink points” (see, e.g., Riebesell, 1922, Chapter 5). The formula for the year 2008 (the last year of a major change) is defined as follows:

$$T = \begin{cases} 0 & \text{if } TI \leq 7,664 \\ (883.74 \frac{TI-7,664}{10,000} + 1,500) \frac{TI-7,664}{10,000} & \text{if } 7,664 < TI \leq 12,739 \\ (228.74 \frac{TI-12,739}{10,000} + 2,397) \frac{TI-12,739}{10,000} + 989 & \text{if } 12,739 < TI \leq 52,151 \\ 0.42TI - 7,914 & \text{if } 52,151 < TI \leq 250,000 \\ 0.45TI - 15,414 & \text{if } TI > 250,000. \end{cases}$$

For married taxpayers filing jointly, the tax is twice the amount of applying the formula to half of the married couple’s joint taxable income: $T_m(TI_1, TI_2) = 2 * T(\frac{TI_1 + TI_2}{2})$. In addition to the personal income tax, households pay the “Solidaritätszuschlag”, a tax supplement originally introduced to finance the German reunification. During the period of interest, 2000-2018, the supplement amounts to 5.5% of the income tax liability. See Doerrenberg et al. (2017) for an overview of the German personal income tax and its deduction possibilities.

Euro) was introduced with a top marginal tax rate of 45%. All nominal start and end points have been adjusted multiple times since 2008 to correct for inflation.

Marginal Employment (“Mini-Jobs”). Marginal employment contracts, called mini-jobs, are jobs with earnings below a time-varying threshold as pictured in Panel C of Figure A.1. The maximum income for marginal employment currently amounts to 450 Euro per month. Jobs below this threshold are exempted from social security contributions and income tax.³ The so-called mini-jobs were part of the Agenda 2010 labor market reforms (also called Hartz reforms) to lower entry barriers to the labor market. Over our sample period in each year around 4.5-5 million workers hold only a mini-job, while another 2.7 million workers use marginal employment as a form of secondary jobs. Mini-jobs are common among benefit recipients, students and pensioners to increase their monthly income. As a result of the tax incentives for married couples, that rewards unequal labor incomes in marriages, there are also many married women who take up mini-jobs. While, in principle, marginal employment is not limited to certain industries, the share of marginal employees is highest in hospitality, services, retail and agriculture (Hohendanner and Stegmaier, 2012). There were two reforms during our sample period (see Gudgeon and Trenkle 2020 for details). In April of 2003, the monthly earnings threshold for mini-jobs was raised from 325 Euro to 400 Euro. It also abolished an weekly working hours limit of 15 hours for mini-jobs, a constraint that was likely not binding. Probably most importantly, the reform also allowed workers to hold a (tax exempt) mini-job as a secondary job at a different employer. A second reform in 2013 raised the earnings threshold from 400 to 450 Euro. Note that apart from these reforms the earnings thresholds have remained constant at the nominal values and thus were gradually falling in real terms.

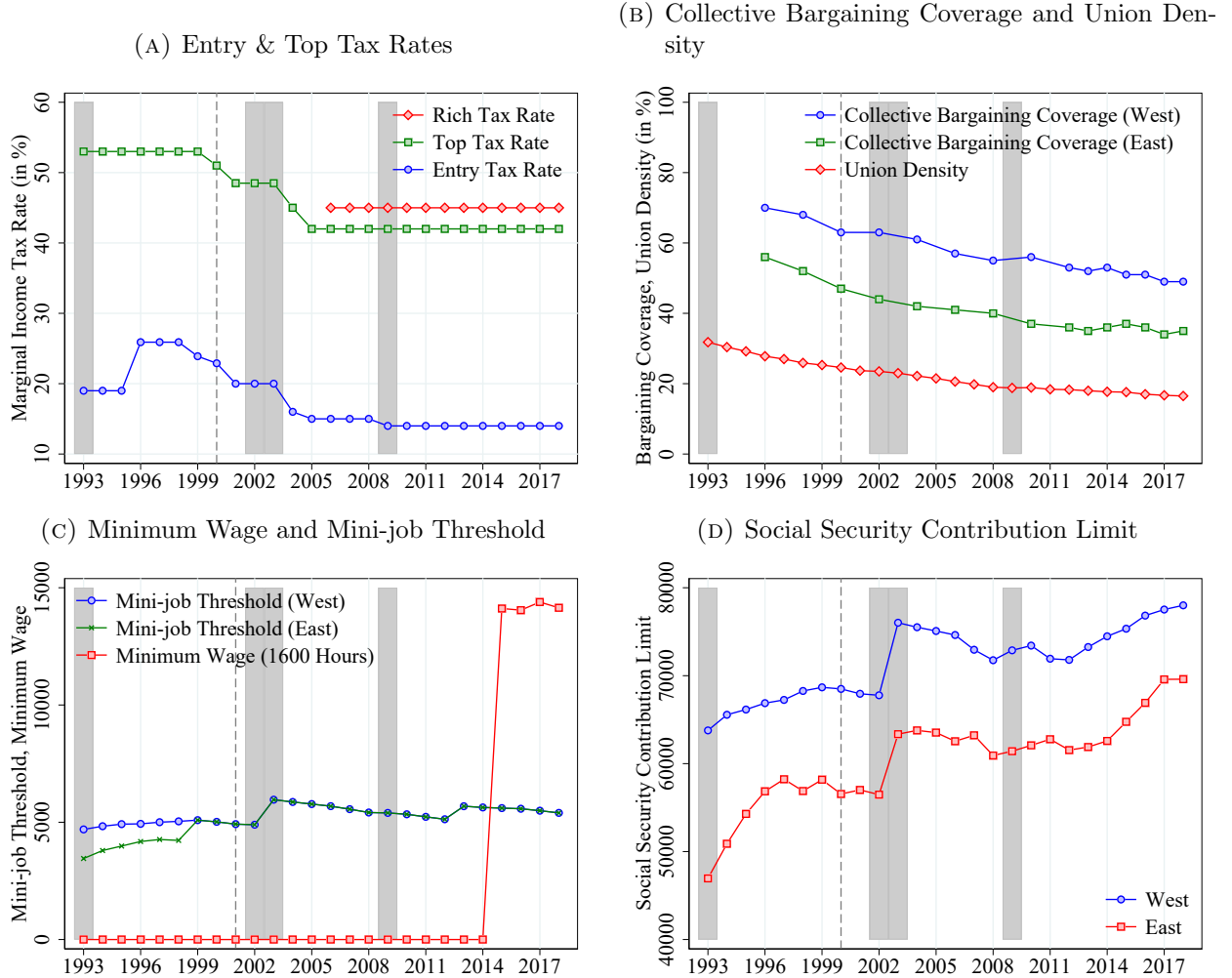
Minimum Wage. Germany introduced a statutory national minimum wage of 8.50 Euro in 2015. It was gradually increased to 8.84 Euro (January 2017), 9.19 Euro (January 2019) and after several more steps is currently at 9.82 Euro (January 2022). Real term values for our sample period are displayed in Figure A.1 (C). Before 2015 different wage floors existed in 12 industries: construction, roofing, cleaning, and nursing among others. Furthermore, some of the larger industries have binding collective agreements that set minimum wages. The impact of the wage floor on wages varied by region and affected about 15 percent of all employees (Dustmann et al., 2022).

Collective bargaining and union density. An effective instrument in Germany to set wages are tariff agreements between union and employer representatives that often have a binding character for all firms in a certain industry. The worker coverage of industry-level collective bargaining agreements varies between former West and East Germany and decreases over time (see Panel B of Figure A.1). Especially start-ups and smaller firms are less likely to be part of a collective agreement. Less common firm-level collective bargaining agreements cover an additional 2% of firms and 8% of employees in 2018 (Ellguth and Kohaut, 2019). The union density (union members out of all employees) declined steadily at the same time.

³A person can hold multiple mini-jobs but then only the first 450 Euro are tax exempt.

Social Security Contribution Limits. The contributions to the pension system and unemployment insurance are capped. The limit differs between East and West Germany and increases over time, roughly following inflation. Figure A.1 (D) shows the limits for East and West Germany from 1993 to 2018 in real terms (2018 Euro).

FIGURE A.1: INSTITUTIONAL BACKGROUND



Notes: This figure shows key institutional parameters for our period of analysis including tax rates (Panel A, source: Federal Ministry of Finance), collective bargaining (Panel B, source: [Ellguth and Kohaut \(2020\)](#) and OCED), mini-job thresholds in 2018 Euro (Panel C, source: Deutsche Rentenversicherung) and the social security contribution limit in 2018 Euro (Panel D, source: Deutsche Rentenversicherung), which is relevant for the top coding in the IAB data. Shaded areas indicate recessions.

A.2 Macroeconomic Background

The macroeconomic development in Germany from 1993-2018 can be broadly split into two periods: before and after 2005 (see Figure A.2). The first time span was characterized by low growth and high unemployment (above 10%) and Germany was often referred to as “the sick man of Europe” ([Dustmann et al., 2014](#)). This changed in the mid-2000s after a series of labor market and tax

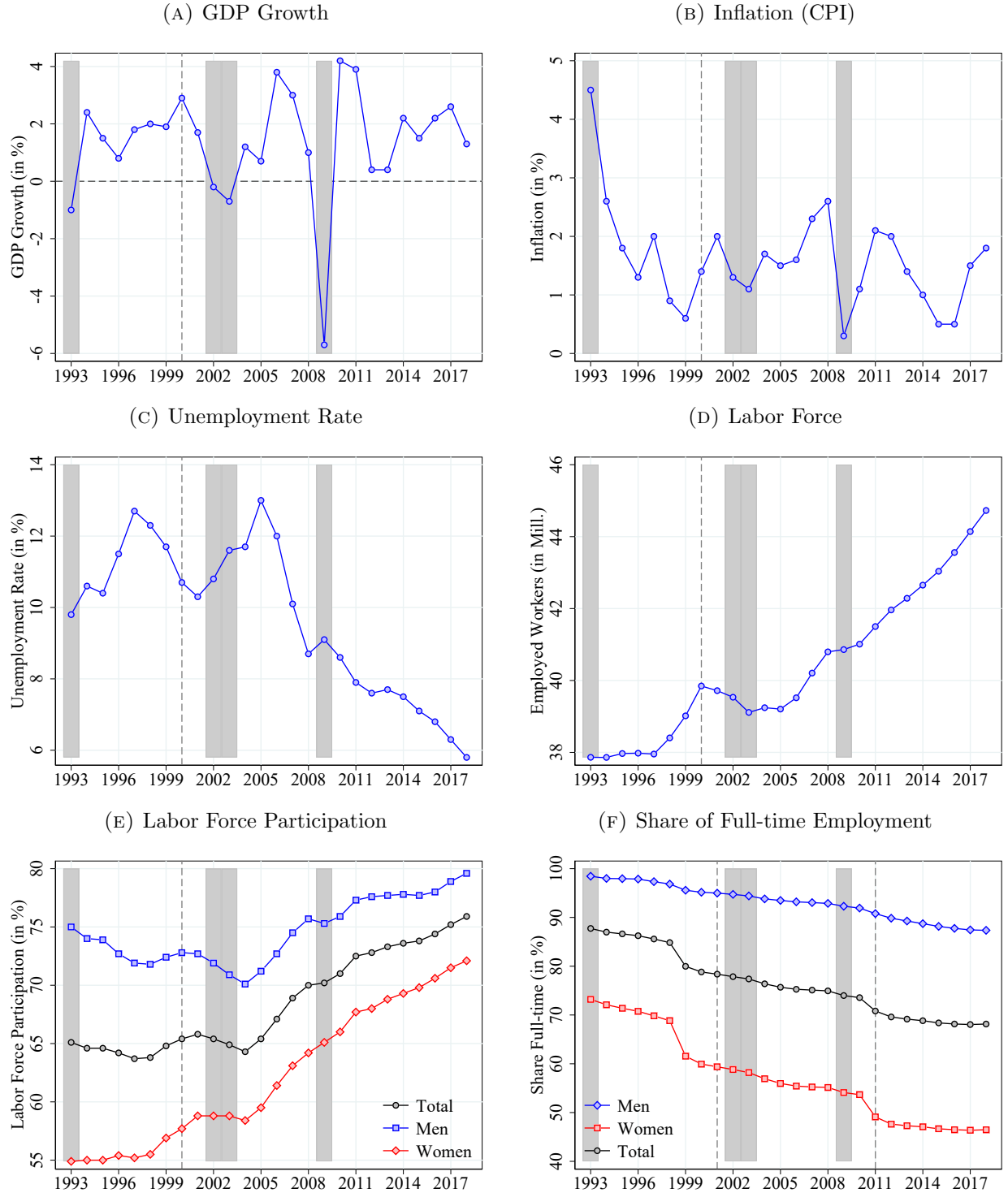
reforms were implemented. While the causal effect of these reforms (called “Hartz reforms”) on the labor market development and the exact mechanisms are still discussed in the literature, it is undisputed that these reforms “worked” - somehow. How and whether the effects were as desired is sometimes the subject of controversial debate. Critics complain, for example, that the new system is unfair and fosters the growth of the low-wage sector in Germany. Supporters of the existing system counter that the reforms have made it possible to reduce unemployment in Germany since 2005 in the first place, and that abolishing them would jeopardize this success. Critics, in turn, doubt the thesis of the positive labor market effects of the reforms and cite other reasons for the reduction in unemployment. (Macro)economic analyses of the reforms (e.g., [Krebs and Scheffel, 2013, 2017](#); [Launov and Wälde, 2013](#); [Hartung et al., 2018](#); [Bradley and Kügler, 2019](#); [Hochmuth et al., 2021](#)) show that the reforms indeed played an important role for the decline in (structural) unemployment, but they are not the only explanatory factor for the positive labor market development.

Nevertheless, neither the Great Recession nor the Euro Crisis affected the German labor market severely. In contrast to the United States and most other EU countries, Germany experienced almost no increase in unemployment, despite a sharp decline in GDP in 2008 and 2009.⁴ Moreover, labor force participation rates of both women and men increased steadily after 2004 and the unemployment rate fell below 6% in 2018.

A notable feature over this time period was a large increase in labor force participation of women, from around 55 percent to more than 70 percent as shown in Figure [A.2 \(E\)](#). However, unlike in countries such as the US, this increase was almost exclusively driven by women entering the labor market in part-time and marginal employment, so that the full-time share over this period fell from 75 to around 50 percent for women. For men, labor force participation and the part-time share also increased substantially since 2003, though nowhere near as dramatic as for women.

⁴The system of short-time work buffered the potential increase in unemployment in Germany as at the height of the economic crisis in mid-2009, the number of short-time workers peaked at 1.5 million helping to cushion the labour market impact of the crisis ([Brenke et al., 2013](#)).

FIGURE A.2: MACROECONOMIC BACKGROUND



B IAB: Social Security Data

The first source of data, which we refer to as the IAB data, is the Integrated Employment Biographies (IEB) supplied by the Institute for Employment Research (*“Institut für Arbeitsmarkt- und Berufsforschung (IAB)”* in German). The IEB are administrative data covering all individuals subject to social security contributions and marginal employment. Moreover, unemployment spells and episodes in active labor market policies are included as well. The IEB allows to follow individuals from labor market entry to retirement. We use 10% random sample of individuals that are either in employed or unemployed, i.e. we exclude persons in active labor market policies.

Employers have to file employment records at least annually or whenever information changes that impacts unemployment benefit or pension calculation. Labor earnings are reported including bonuses and extra pay but only up to the social security contribution limit, which is at an annual labor income of 78,000 Euro in West Germany and 69,600 Euro in East Germany in 2018 (see Figure A.1 for real values over time). All earnings above that limit are censored. We describe below how we impute wages for some of the analyses. Besides to the top-coding, another limitation of the IAB data is that it does not include self-employed individuals (around 4 million) and civil servants (around 1.9 million individuals).

The data contains information on the exact dates of employment and earnings as well as a variety of worker and firm characteristics such as gender, education, year of birth, occupation or industry code. The information is spell based, i.e. accurate to the date and especially with respect to earnings trustworthy. Note, however, that the education information contains some missing values which we impute (described below) using the procedure suggested by the IAB. Moreover, throughout 2011, the reporting procedure for full-time and part-time employment in the social security data changed. This results in a small fraction of workers being falsely classified as working full-time before 2012. We are able to partially correct the full-time indicator in the years prior to 2012 using a cell-wise reclassification approach (see below).

B.1 Top-Coding and Imputation of Wages

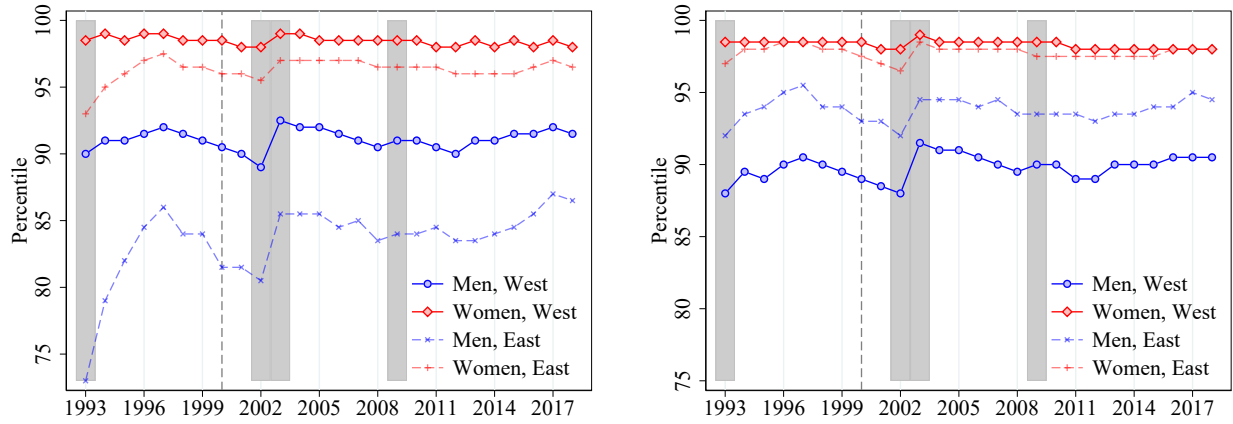
Figure B.1 (A) shows that in the overall labor income distribution for women, the West German social security contribution limit is binding for women at roughly the 99th percentile, while the East German limit is binding at around the 96th percentile. For the earnings distribution of men the limits are much more binding, with the West German threshold binding at roughly the 90th percentile and the East German threshold at roughly the 85th percentile when applied to the whole distribution.

In Figure B.1 (B) we ask the related but different question, where within the earnings distribution of East and West German workers the social security contribution limits fall in a given year by gender. Since East German incomes are still much lower than in West Germany, this pushes up at which percentile in the East/West-specific distribution the social security limit actually becomes binding. The figure highlights clearly that only the highest earning women in East and West are

affected by their respective thresholds and, thus, have censored earnings reported. West German men are the most likely to be affected (threshold around the 90th percentile), while the East German limit now lies close to the 95th percentile of the East German earnings distribution. This means that for West German men the highest 10% of earnings are subject to censoring while for East German men this is only the case for the highest 5-7%. Another interpretation of Figure B.1 (A) is hence that it shows where the same 5-7 percent of male workers top coded in the East according to Figure B.1 (B) rank in the overall male income distribution of Germany. The IAB data does not contain information on incomes above the social security contribution limit. Several imputation algorithms have been proposed for wages above the top-coding limit. We use the algorithm suggested by [Card et al. \(2013\)](#) and implemented by [Dauth and Eppelsheimer \(2020\)](#) for SIAB to impute daily wages which we then aggregate to annual incomes for our analysis.

FIGURE B.1: SHARE OF UNCENSORED OBSERVATIONS IN THE IAB DATA

(A) Share by Year and Gender below top coding for East/West (B) Share by Year, Gender and East/West below respective top coding



Notes: This figure shows the percentile of the labor earnings distribution at which the top-coding due to the social security contribution limit becomes binding. This corresponds to the share of uncensored observations. In Panel A, the percentiles are calculated by year based on the earnings distributions of men and women separately. In Panel B, four different distributions are calculated for men and women in East and West Germany separately for each year. Shaded areas indicate recessions.

B.2 Imputation of the Education Indicator

The education information in the IEB contains missing values predominantly for workers holding a mini-job. The number of missings increases over time and amounts to 22 percent for regular employees and 60 percent for marginal part-time employees in our data. To cope with the missing information we use the imputed education variable provided by the IAB, which adds missing information by forward and backward writing. The procedure is described in [Thomsen et al. \(2018\)](#).

B.3 Correction of the Full-Time Indicator

In 2011, the reporting procedure for full-time and part-time employment in the social security data changed. This results in a enhanced number of classification updates of workers that have been

misclassified as full-time beforehand, but in fact were working part-time, leading to an artificial drop in full-time share and jump in part-time share. The procedure changed throughout the whole year of 2011, which leads to a structural break between 2010 and 2012 with an intermediate update in 2011. [Fitzenberger and Seidlitz \(2020\)](#) document the consequences of this break for analyses of wage inequality and provide an reweighting procedure to correct for misclassifications before 2012.

We use a non-parametric correction approach instead of estimating weights, reclassifying full-time to part-time in 2011 and before if potentially misclassified. This allows us to use the IAB sample consistently without inducing potential bias to other (correct) variables when applying weights to the sample.

First, we restrict our sample to potentially affected individuals in the relevant time period and age group (25 to 55). We apply our correction only to full-time and part-time workers, marginal employment should be unaffected. Following [Fitzenberger and Seidlitz \(2020\)](#), we exclude individuals with wages above a certain threshold. We similarly exclude all observations in the years 2001 to 2011 from the correction when the respective real earnings are above the 80th percentile of earnings in 2012 for women and 25th percentile of earnings in 2012 for men. We further calculate a distance measure θ to the percentile threshold, normalized to 0 to 100.

Second, we use gender, an indicator for former West or East Germany, 11 age groups, educational attainment (6 groups) and days in employment (4 groups) to divide our sample into cell-groups. We then cell-wise calculate the share of full-time employment separately for 2009 to 2013. Using this full-time share, we apply a smoothed correction to cell-wise full-time shares for the years 2011 and 2010, based on the full-time share differences as well as (smoothed) pre- and post-trends. For the years 2001 to 2009, we cell-wise deduct the full-time share difference 2010 to 2012 and smoothed trend from the original full-time share. This results in (at least partially) corrected full-time shares for each cell for 2001 - 2011.

Third, we cell-wise reclassify full-time workers to part-time until the share of full-time workers is decreased to the corrected full-time share of the cell. We do not purely pick observations at random for this but sort according to θ , adding a small amount of noise to the latter. This means the probability to be reclassified increases with lower real earnings but not fully depends on those. We do this separately by year for all observations, because workers frequently change cells between years. Thus, we do not carry forward any reclassification from 2010 and 2011 to earlier years. This means workers' classification of full-time or part-time may switch repeatedly because of the correction. This provides us with more reliable (repeated) cross-section aggregates but may result in higher 1-year transition probabilities from full-time to part-time and vice versa in 2001 - 2011.

This procedure resolves most of the structural break for most of the cells in (repeated) cross-section. However, our approach does not necessary fully correct the structural break in attempt to not over-correct. This means there still occur some artifacts in the data around 2011 but to a much smaller degree than without the correction.

C TPP: Tax Data

C.1 General Description

The second source of data is the German Taxpayer Panel (TPP) (Kriete-Dodds and Vorgrimler, 2007), which is an administrative data set based on the universe of personal income tax returns in Germany.⁵

The data set covers all tax units filing tax returns in the period 2001-2016 in Germany. The 2001 to 2016 panel has a total of 58,808,899 unique records for which information is available for at least two waves of years. We work with a 25% random sample of these records. The unit of observation is the taxpayer, i.e., either a single individual or a couple filing jointly. In the latter case, income from all different sources (such as labor or business income) are measured on the individual level before the income is aggregated at the couple's level. The same is true for many deductions and allowances which are available on the individual level.

The data set contains all information necessary to calculate a taxpayer's annual income tax. This includes basic socio-demographic characteristics such as year of birth, gender, family status, number of children as well as detailed information on gross income (differentiated by seven different sources) and basic tax-specific parameters such as work-related expenses and deductions. A list of the variables - differentiated by assessment year - is included in the dataset description available for download.

The data set is not top-coded. Therefore, this data set is especially suited for the analysis of inequality in the upper tail of the income distribution. It is, however, missing the very bottom of the income distribution as incomes below the marginal employment threshold are except from the personal income tax and hence not included in the data.

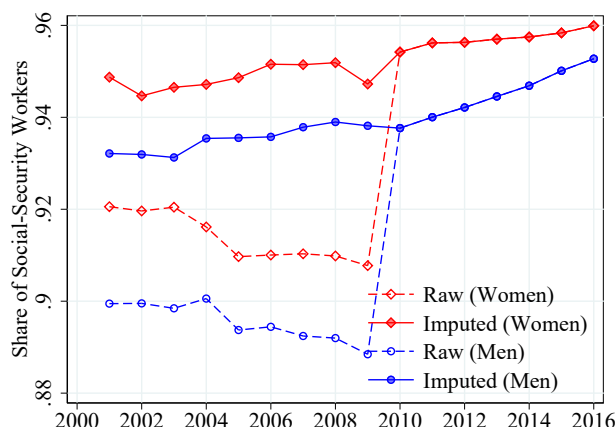
The 2001 to 2011 waves of the Taxpayer Panel (TPP) were compiled on the basis of annual income tax statistics (*Geschäftsstatistik*) of each of the 16 federal states which were then combined into one dataset for Germany. These cross-sectional data contain information from the income tax returns of around 27 million German taxpayers that filed a tax return and were linked to form a panel via the tax numbers and indirect identifiers. Starting with the 2012 assessment year, there was a change in the procedure. Instead of the annual income tax statistics, the federal wage and income tax statistics (*Bundesstatistik zur Lohn- und Einkommensteuer*), which had been collected every three years until then, was collected annually and formed the new basis for the TPP from 2013 onwards. In addition to taxpayers filing a tax return, the federal statistics also include around 12 million non-assessed taxpayers who did not file a tax return and paid the income tax withholding tax (*Lohnsteuer*). We describe how we deal with this structural break below in Appendix D.

⁵See <https://www.forschungsdatenzentrum.de/de/steuern/tpp> for more information (albeit only available in German) on the TPP data. This data has been, for example, used by Doerrenberg et al. (2017) and Dolls et al. (2018) who also provide additional information on the data. More detailed information on the construction and use of the TPP is presented in the usage concept available for download (in German only) here: <https://www.forschungsdatenzentrum.de/de/10-21242-73111-2016-00-01-1-1-0>.

C.2 Imputation of Social Security Indicator in Pre-2010 TPP Data

The definition of non-social-security workers ($C_{it} = 1$) in the TPP is imprecise prior to 2010 resulting in too many non-social-security workers (compared to official IAB data). Figure C.1 shows that the share of social-security workers is too low prior to 2010 (dashed lines). The differences is roughly 4 percentage points for both men and women. Panel A of Figure C.2 shows the share of social-security workers by year and earnings bin. Again, this share is lower at almost all income levels for all years before 2010 compared to the years after.

FIGURE C.1: SHARE OF SOCIAL-SECURITY WORKERS IN THE TPP



Notes: This figure shows the average share of social-security liable workers in the raw TPP data over time before (dashed lines) and after the structural break in 2010 for men and women separately. It also shows the corrected share after the application of the imputation procedure described in this Appendix.

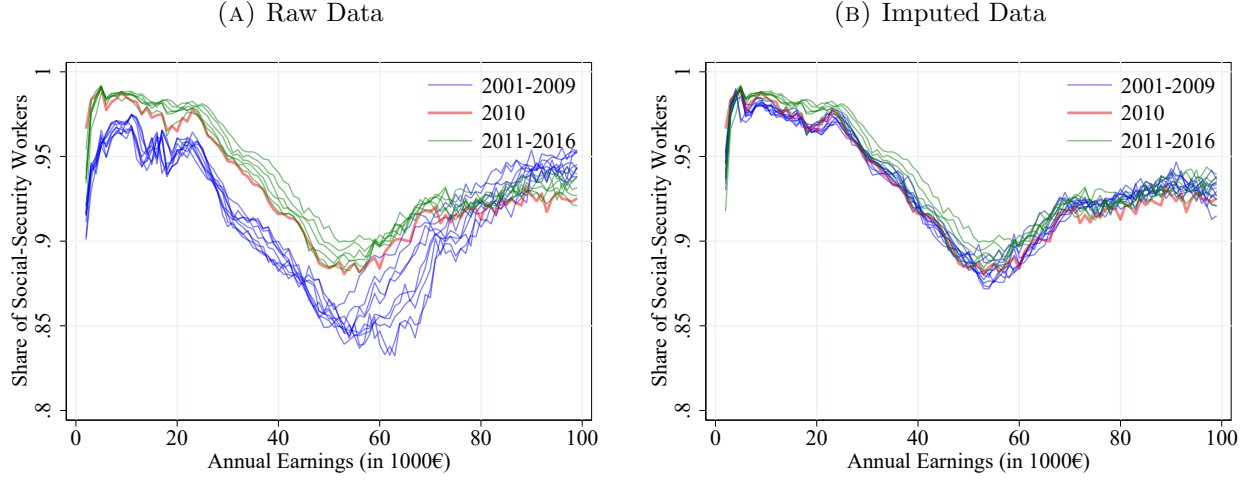
Backward-Imputation Procedure. In order to correct the data, we use the following backward imputation procedure:

- Let G capture all combinations of the observables gender, age group and binned earnings.
- Let c_{gt} be the share of social-security workers observed in the raw TPP data in group $g \in G$ and year t and c_{gt}^* the corresponding true share.
- The observed share is correct for $t \geq 2010$ ($c = c^*$) and incorrect for $t < 2010$ ($c \neq c^*$).
- We assume that the true share in each gender-age-earnings group is roughly time-invariant.

$$c_{gt}^* = c_g^* + \varepsilon_{gt} \text{ with } \varepsilon_{gt} \sim N(0, \sigma_g) \quad (\text{C.1})$$

- We use the years $t \geq 2010$ to estimate c_g^* and the standard deviation of ε_{gt} (could also be done via regression of C_{it} on a set of group dummies).
- We then approximate the true shares for $t < 2010$ using these estimates. Denote these by \hat{c}_{gt} .
- This allows us to predict the share of mis-coded observations, $\eta_{gt} = c_{gt} - \hat{c}_{gt}$ for $t < 2010$.

FIGURE C.2: SHARE OF SOCIAL-SECURITY WORKERS BY EARNINGS BINS IN THE TPP



Notes: This figure shows in Panel A the share of social-security liable workers in the raw TPP data over time before (blue lines) and after (green lines) the structural break in 2010 (red line) across the earnings distribution. Panel B shows the same information after the application of the imputation procedure described in this Appendix.

- Using data from 2010 onward, we also estimate the transition probabilities for the social-security indicator conditional on gender, age and earnings bin.

$$\pi_{gt}^0 = \Pr(C_t = 0 | C_{t+1} = 1, G_t = g) \quad (\text{C.2})$$

$$\pi_{gt}^1 = \Pr(C_t = 1 | C_{t+1} = 0, G_t = g) \quad (\text{C.3})$$

- For the years 2001 to 2009, we re-code the social-security indicator C_{it} as follows:
 - (1) Define τ as the first year where the indicator is not (yet) correctly coded (or imputed). Set $\tau = 2009$.
 - (2) Set the imputation flag F_i to zero for all workers.
 - (3) For workers who are observed in year $\tau + 1$, we impute C_{it} for $t \leq \tau$ using the transition probabilities and their value of $C_{i,\tau+1}$ as a starting point.⁶
 - (4) Re-compute $c_{g\tau}$ and update the share of mis-coded observations in year τ and group g , $\eta_{g\tau}$.⁷
 - (5) If, as expected, $\eta_{g\tau} \geq 0$, set $x = 0$, otherwise set $x = 1$.
 - (6) Randomly choose a fraction η_{gt} of the subset of workers with $G_{it} = 1$ and $C_{i,2009} = x$, and re-code their civil servant indicator $C_{i\tau}$ accordingly.
 - (7) If $\tau = 2001$, stop. Otherwise, set τ to $\tau - 1$ and return to step (3).

The results of this imputation procedure are shown in Panel B of Figure C.2. Now the share of social-security workers is similar for all years across the income distribution.

⁶In the initial step with $\tau = 2009$, if a worker is not observed in 2010 but is observed in some later period t' (starting in 2012, the TPP has full coverage), we use $C_{i,t'}$ as a starting point for the imputation for year 2009.

⁷There should still be too few social security workers as some workers who exit the data before 2010 are still mis-coded.

D Combined IAB-TPP Data

This Appendix describes how we combine the IAB and TPP data for our main analysis. As we are not allowed to directly link the micro data of IAB and TPP due to data protection legislation in Germany, we need to rely on non-parametric matching techniques to construct earnings/income distributions as well as distributions of income changes.

Before combining the data, we show descriptive statistics for the IAB and TPP data sets for the year 2008 separately for men and women who are between 25 and 55 years old in Table D.1. Unsurprisingly, the TPP has fewer observations due to missing non-filers and mini-jobs. As the TPP data contains only very limited demographic information, we can only compare both datasets in terms of age. The TPP population is slightly older which again can be attributed to missing observations who are more likely to be at the beginning of their career.

TABLE D.1: DESCRIPTIVE STATISTICS FOR EARNINGS DATA (YEAR 2008)

| | Men | | Women | |
|-------------------------------|------------|------------|------------|------------|
| | IAB (1) | TPP (2) | IAB (3) | TPP (4) |
| Observations (in mill.) | 12.430 | 9.058 | 11.228 | 7.409 |
| Mean Earnings (in 2018-Euro) | 40,562 | 46,406 | 24,010 | 28,089 |
| <i>A. Age and Nationality</i> | | | | |
| Share Age 25–34 | 0.275 | 0.230 | 0.261 | 0.233 |
| Share Age 35–44 | 0.346 | 0.352 | 0.334 | 0.331 |
| Share Age 45–55 | 0.378 | 0.418 | 0.405 | 0.436 |
| Non-German | 0.090 | – | 0.066 | – |
| <i>B. Education</i> | | | | |
| Schooling (≤ 10 years) | 0.050 | – | 0.057 | – |
| Vocational training | 0.621 | – | 0.596 | – |
| Abitur (& voc. training) | 0.111 | – | 0.157 | – |
| College Degree | 0.206 | – | 0.177 | – |
| No Education Data | 0.011 | – | 0.014 | – |
| <i>C. Employment Level</i> | | | | |
| Full-Time | 0.919 | – | 0.521 | – |
| Part-Time | 0.057 | – | 0.347 | – |
| Mini-Job | 0.024 | – | 0.132 | – |
| Days in Employment | 342.3 | – | 342.2 | – |

Notes: This table shows descriptive statistics for the IAB and TPP data sets for the year 2008 separately for men and women who are between 25 and 55 years old.

D.1 Reweighting the TPP Data to Match the IAB Data

While we have access to the 'population' version of the available taxpayer data, the TPP still does not cover the *entire* population of income taxpayers. In particular, there are two deviations. First,

the TPP only includes tax units that appear in at least two waves [D1]. Second, for the years 2001 to 2011 the TPP only includes information of taxpayers who file a tax return statement [D2]. Hence, around 12 million non-filers are missing per year. Importantly, only workers who do not receive any non-labor income (above an exemption level of roughly 400 Euro) have the option not to file a tax return.⁸

We correct these two deviations by reweighting the TPP data. Thereby, we distinguish between workers whose earnings are subject to social security contributions and who are included in the IAB, and workers whose earnings are not subject to social-security contributions (e.g. civil servants). Note that for our core analysis in Section 3 we only consider the former. The latter are only part of the total income sample in Section 4.

D.1.1 Reweighting the Pre-2012 TPP Data to Account for Missing Non-Filers

For social security workers, we use information from the IAB (headcounts by gender, age group and 1,000 Euro earnings bin) as well as post-2012 TPP data to reweight observations. The reweighted data match the joint distribution of gender, age group and earnings below the social security contribution limit and the number of workers above this limit as well as the share of non-filers by gender, age group and earnings (above the top-coding threshold in the IAB) observed in the post-2012 TPP data.

In particular, we compute from the IAB the number of workers in each (real) annual earnings bin by gender and age group (25-29, 30-34, ..., 50-55). We use bins of 1,000 Euro each up until 60,000 Euro, above which the IAB is top-coded. Hence, we only know the total number of workers above 60,000 Euro. To reweight workers above this cutoff, we additionally compute from the post 2011 TPP data the average share of non-filers in 20 time-invariant earnings vintiles above the cutoff (by gender and age group). The TPP data further allows us to distinguish between mandatory filers and voluntary filers. Loosely speaking, filing a tax return is mandatory when a worker files jointly with his/her married spouse, received non-labor income (including transfers) above 410 Euro or received other labor income for which the employer did not deduct (enough) income taxes.

In the following, we describe the reweighting procedure in more detail.

Notation:

G stratification group (combination of gender, age group and earnings bin)

N_{gt}^* target number of workers in group $g \in G$, computed using IAB data

N_{gt} observed number of workers in group g in the TPP ($N_{gt} = N_{gt}^v + N_{gt}^m$)

N_{gt}^m observed number of mandatory filers in group g in the TPP

N_{gt}^v observed number of voluntary filers in group g in the TPP

N_{gt}^n observed number of non-filers in group g in the TPP (equals zero before 2012, $N^n < N^v$ after 2012)

⁸The earnings distributions (headcounts by bins) in Figures 1 and D.1 visualize this difference.

w_{gt}^m constructed weight of mandatory filers in group g

w_{gt}^v constructed weight of voluntary filers in group g

Procedure for Workers Below the IAB Top-Coding Cutoff.

- (i) Compute the average ratio between target and observed headcounts for the years 2012 to 2016:

$$\delta_g = E_t[N_{gt}^*/N_{gt}|t \geq 2012] \quad (D.1)$$

- (ii) Construct target headcounts net of D2 for the years 2001 to 2011 as $N_{gt}^{1*} = \frac{N_{gt}^*}{\delta_g}$

- (iii) Compute the weights for voluntary and mandatory workers as

$$w_{gt}^v = \begin{cases} \frac{N_{gt}^v + (N_{gt} - N_{gt}^{1*})}{N_{gt}^v} \delta_g & \text{if } t < 2012 \\ \frac{N_{gt}^*}{N_{gt}} & \text{if } t \geq 2012 \end{cases} \quad (D.2)$$

$$w_{gt}^m = \begin{cases} \delta_g & \text{if } t < 2012 \\ \frac{N_{gt}^*}{N_{gt}} & \text{if } t \geq 2012 \end{cases} \quad (D.3)$$

Procedure for Workers Above the IAB Top-Coding Cutoff. We partition the top earnings bin (above 60,000) into 20 fractiles by gender and age group. Let H be the combination of gender, age group and this partition. We use the same notation as for below-cutoff workers but replace G and g by H and h . The key assumption is that D2 is constant over time and that the share of non-filers is time-invariant within each combination of gender, age group and earnings fractile.

- (i) Compute the average share of non-filers in each group h for the years 2012 to 2016

$$\eta_h = E_t[N_{ht}^n/N_{ht}|h, t \geq 2012] \quad (D.4)$$

- (ii) Compute the number of missing non-filers in $t < 2012$ as

$$\hat{N}_{ht}^n = \frac{N_{ht}}{1 - \eta_h} - N_{ht} \quad (D.5)$$

- (iii) To correct for D2, compute the auxiliary weights for voluntary and mandatory workers as

$$\tilde{w}_{ht}^v = \begin{cases} \frac{N_{ht}^v + \hat{N}_{ht}^n}{N_{ht}^v} & \text{if } t < 2012 \\ 1 & \text{if } t \geq 2012 \end{cases} \quad (D.6)$$

$$\tilde{w}_{ht}^m = 1 \quad \text{for all } t \quad (D.7)$$

- (iv) Compute the total headcount implied by the auxiliary weights in the original top earnings bin (by gender and age group):

$$\tilde{N}_t = \sum_h \left(\tilde{w}_{ht}^v N_{ht}^v + \tilde{w}_{ht}^m N_{ht}^m \right) \quad (\text{D.8})$$

- (v) To correct for D1, we rescale the auxiliary weights to match the target headcount in the top earnings bin. This gives:

$$w_{ht}^x = \tilde{w}_{ht}^x \frac{N_t^*}{\tilde{N}_t} \quad \text{for } x \in \{v, m\} \quad (\text{D.9})$$

D.1.2 Reweighting Non-Social-Security Workers in the TPP

For non-social-security workers, we only use post-2012 TPP data for reweighting as these workers are not included in the IAB data. Hence, the reweighted data match the share of non-filers by gender, age group and earnings observed in the post-2012 TPP data. For brevity, we sometimes refer to social-security workers as regular workers and to non-social-security workers as other workers.

The reweighting procedure to account for non-filing non-social-security workers is very similar to the one used for social security workers above the cutoff. The main difference is that we have no data to correct for D2 as we cannot rely on IAB data for non-social-security workers. We first group civil servants based on gender, age group and (time-invariant) earnings fractiles.⁹ We use the same notation as above.

- (i) Compute the average share of non-filers for the years 2012 to 2016

$$\eta_g = E_t [N_{gt}^n / N_{gt} | t \geq 2012] \quad (\text{D.10})$$

- (ii) Compute the number of missing non-filers in $t < 2012$ as

$$\hat{N}_{gt}^n = \frac{N_{gt}}{1 - \eta_g} - N_{gt} \quad (\text{D.11})$$

- (iii) Correcting for D2, compute the weights for voluntary and mandatory workers as

$$w_{gt}^v = \begin{cases} \frac{N_{gt}^v + \hat{N}_{gt}^n}{N_{gt}^v} & \text{if } t < 2012 \\ 1 & \text{if } t \geq 2012 \end{cases} \quad (\text{D.12})$$

$$w_{gt}^m = 1 \quad \text{for all } t \quad (\text{D.13})$$

⁹Age groups are the four quartiles and earnings bins are defined by the gender and age group specific P5, P10, P20, ..., P90, P95 of the real earnings distribution pooled over the entire sample period. This gives $2 \times 4 \times 12 = 96$ groups in each year.

D.2 Combined IAB-TPP Data in Earnings Analysis (Section 3)

D.2.1 Combined Earnings Distribution

For the core analysis of labor earnings, we focus exclusively on social-security workers as we do not have IAB data for non-social-security workers. The main idea in constructing the combined distribution of earnings is the following: Below the top-coding threshold of 60,000 Euro, we use the (true) earnings distribution from the IAB data. Above the cutoff, we use the conditional earnings distribution from the (reweighted) TPP along with the (true) number of workers above the cutoff in the IAB.

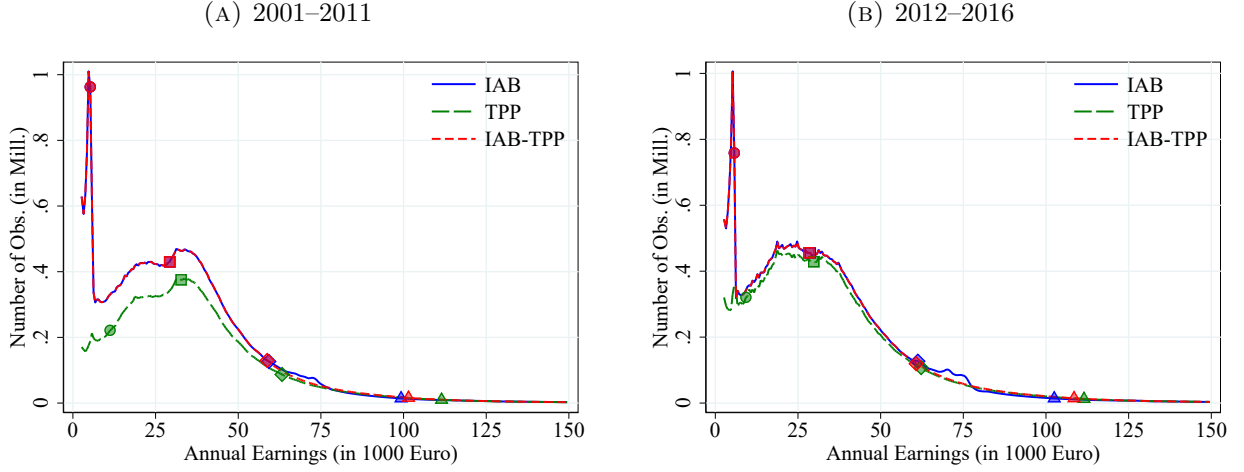
Technically, we (i) estimate the CDF of earnings in both data sources by monotonically interpolating a wide range of quantiles, and (ii) construct the combined CDF as:

$$F(y) = \begin{cases} F^{IAB}(y) & \text{if } y \leq \bar{y} \\ F^{IAB}(\bar{y}) + F^{TPP}(y|y > \bar{y}) \left(1 - F^{IAB}(\bar{y})\right) & \text{if } y > \bar{y} \end{cases} \quad (\text{D.14})$$

Figure 1 in the main text, Figure D.1 as well as Tables D.2, D.3 and D.4 show selected percentiles of the earnings distribution in the combined IAB-TPP data (CS sample) as well as in the IAB and TPP data for men, women and in the population respectively. Percentiles below 60,000 Euro (P75 and below) are practically identical in the IAB-TPP and IAB data, while higher percentiles are closer to the TPP data.¹⁰

¹⁰The small deviations below 60,000 Euro are the result of how we combine the IAB and TPP data. After interpolating the quantiles, we discretize the respective distributions on a very fine grid and then combine the discrete distributions. The deviations for higher percentiles are mostly driven by the fact that the number of observations in the TPP is smaller than in the combined IAB-TPP data. Adding individuals mostly at the bottom of the distribution moves the higher percentiles to different (lower) points in the income distribution (Krolage et al., forthcoming).

FIGURE D.1: ANNUAL EARNINGS DISTRIBUTION IN IAB, TPP AND COMBINED DATA – POPULATION



Notes: This figure shows the number of observations in real earnings bins for the IAB, the TPP and the combined data (IAB-TPP) in the full population (men and women). A complementing figure by gender can be found in Figure 1. Panel A shows averages across the years 2001 to 2011 where non-filing workers (Lohnsteuerfälle) are not included in the TPP and Panel B shows averages across the years 2012 to 2016 where the TPP data include these workers. We exclude earnings from the TPP that are not subject to social security contributions (e.g. salaries of civil servants) which are not covered in the IAB. The circular, square, diamond and triangle-shaped markers depict the 10th, 50th, 90th and 99th earnings percentile in the respective data sets. We use 500 Euro bins below 80,000 Euro and 1,000 Euro bins above 80,000 Euro but always plot the number of observations per 1,000 Euro bins. The IAB data are imputed above the social security contribution limit. Table D.4 shows selected earnings of these distributions percentiles across the different datasets.

TABLE D.2: EARNINGS PERCENTILES IN IAB, TPP AND COMBINED IAB-TPP DATA – MEN

| Year | N | Mean | P5 | P10 | P25 | P50 | P75 | P90 | P95 | P99 | P99.9 | P99.99 |
|---------------------|--------|--------|--------|--------|--------|--------|--------|--------|---------|---------|---------|-----------|
| IAB-TPP Data | | | | | | | | | | | | |
| 2001 | 13.113 | 41,587 | 6,331 | 11,893 | 25,563 | 37,599 | 50,524 | 69,835 | 86,881 | 141,835 | 330,996 | 939,346 |
| 2002 | 12.822 | 41,712 | 6,028 | 11,410 | 25,382 | 37,700 | 50,959 | 70,711 | 87,866 | 142,542 | 331,218 | 894,501 |
| 2003 | 12.593 | 41,808 | 5,677 | 10,939 | 25,104 | 37,809 | 51,455 | 71,576 | 88,916 | 143,772 | 319,169 | 824,191 |
| 2004 | 12.419 | 41,580 | 5,406 | 10,088 | 24,398 | 37,434 | 51,274 | 71,861 | 89,514 | 145,724 | 335,621 | 897,846 |
| 2005 | 12.165 | 41,643 | 5,300 | 10,029 | 23,991 | 37,066 | 51,192 | 72,262 | 90,409 | 149,328 | 358,986 | 994,054 |
| 2006 | 12.214 | 41,624 | 5,308 | 9,886 | 23,364 | 36,615 | 51,229 | 72,642 | 91,483 | 153,898 | 377,426 | 1,135,289 |
| 2007 | 12.373 | 41,650 | 5,434 | 10,126 | 23,142 | 36,169 | 50,978 | 72,824 | 92,428 | 157,847 | 401,395 | 1,146,261 |
| 2008 | 12.430 | 41,595 | 5,409 | 10,282 | 23,013 | 36,029 | 50,985 | 72,919 | 92,541 | 158,374 | 397,129 | 1,114,094 |
| 2009 | 12.223 | 41,261 | 5,274 | 9,657 | 22,696 | 35,843 | 50,622 | 72,996 | 92,605 | 157,500 | 376,763 | 1,005,888 |
| 2010 | 12.275 | 41,193 | 5,293 | 9,678 | 22,230 | 35,687 | 51,048 | 73,272 | 92,718 | 157,072 | 376,669 | 1,011,054 |
| 2011 | 12.464 | 41,442 | 5,226 | 9,934 | 22,373 | 35,518 | 51,155 | 73,681 | 93,802 | 160,563 | 392,593 | 1,060,607 |
| 2012 | 12.535 | 41,548 | 5,121 | 10,104 | 22,479 | 35,528 | 51,330 | 74,077 | 94,194 | 160,311 | 387,167 | 1,024,964 |
| 2013 | 12.638 | 41,527 | 5,153 | 9,871 | 22,385 | 35,536 | 51,306 | 74,094 | 94,204 | 160,158 | 391,983 | 1,068,718 |
| 2014 | 12.796 | 41,820 | 5,172 | 9,615 | 22,332 | 35,669 | 51,782 | 74,963 | 95,331 | 162,526 | 396,397 | 1,092,608 |
| 2015 | 12.958 | 42,392 | 5,355 | 9,858 | 22,519 | 35,922 | 52,422 | 76,017 | 96,743 | 165,881 | 415,286 | 1,183,578 |
| 2016 | 13.096 | 42,798 | 5,445 | 9,991 | 22,831 | 36,206 | 52,817 | 76,752 | 97,772 | 168,009 | 414,687 | 1,189,451 |
| IAB Data | | | | | | | | | | | | |
| 2001 | 13.113 | 40,781 | 6,336 | 11,898 | 25,568 | 37,604 | 50,529 | 67,943 | 83,324 | 136,482 | 241,005 | 385,336 |
| 2002 | 12.822 | 41,012 | 6,033 | 11,415 | 25,387 | 37,705 | 50,964 | 68,722 | 85,047 | 138,889 | 243,050 | 387,108 |
| 2003 | 12.593 | 40,952 | 5,682 | 10,944 | 25,109 | 37,814 | 51,460 | 71,056 | 84,389 | 133,374 | 229,630 | 343,458 |
| 2004 | 12.419 | 40,696 | 5,411 | 10,093 | 24,403 | 37,439 | 51,279 | 71,116 | 85,040 | 136,700 | 233,129 | 358,618 |
| 2005 | 12.165 | 40,658 | 5,305 | 10,034 | 23,996 | 37,071 | 51,197 | 71,404 | 86,047 | 139,593 | 243,628 | 371,340 |
| 2006 | 12.214 | 40,559 | 5,313 | 9,891 | 23,369 | 36,620 | 51,234 | 71,520 | 86,821 | 144,560 | 253,764 | 388,987 |
| 2007 | 12.373 | 40,516 | 5,439 | 10,131 | 23,147 | 36,174 | 50,983 | 71,249 | 87,457 | 148,149 | 268,884 | 418,559 |
| 2008 | 12.430 | 40,562 | 5,414 | 10,287 | 23,018 | 36,034 | 50,990 | 70,667 | 88,129 | 151,311 | 272,001 | 422,230 |
| 2009 | 12.223 | 40,260 | 5,279 | 9,662 | 22,701 | 35,848 | 50,627 | 71,190 | 87,712 | 149,965 | 267,141 | 424,189 |
| 2010 | 12.275 | 40,370 | 5,298 | 9,683 | 22,235 | 35,692 | 51,053 | 71,757 | 88,669 | 152,613 | 275,987 | 435,400 |
| 2011 | 12.464 | 40,438 | 5,231 | 9,939 | 22,378 | 35,523 | 51,160 | 71,030 | 89,298 | 152,928 | 274,180 | 428,103 |
| 2012 | 12.535 | 40,499 | 5,126 | 10,109 | 22,484 | 35,533 | 51,335 | 71,043 | 89,484 | 152,143 | 266,911 | 406,601 |
| 2013 | 12.638 | 40,348 | 5,158 | 9,876 | 22,390 | 35,541 | 51,311 | 71,638 | 88,450 | 151,273 | 262,699 | 391,128 |
| 2014 | 12.796 | 40,620 | 5,177 | 9,620 | 22,337 | 35,674 | 51,787 | 72,468 | 89,232 | 152,698 | 271,648 | 412,144 |
| 2015 | 12.958 | 41,009 | 5,360 | 9,863 | 22,524 | 35,927 | 52,427 | 73,256 | 89,845 | 153,676 | 271,108 | 407,527 |
| 2016 | 13.096 | 41,347 | 5,450 | 9,996 | 22,836 | 36,211 | 52,822 | 74,101 | 90,180 | 154,300 | 274,245 | 415,386 |
| TPP Data | | | | | | | | | | | | |
| 2001 | 10.570 | 45,203 | 10,396 | 17,269 | 28,853 | 39,482 | 53,752 | 75,043 | 93,188 | 154,641 | 370,435 | 1,065,239 |
| 2002 | 10.722 | 44,927 | 9,601 | 16,295 | 28,308 | 39,302 | 53,898 | 75,336 | 93,294 | 153,864 | 366,047 | 1,005,616 |
| 2003 | 10.396 | 45,073 | 9,396 | 16,037 | 28,262 | 39,495 | 54,283 | 76,071 | 94,201 | 154,617 | 349,718 | 914,305 |
| 2004 | 10.057 | 45,240 | 9,230 | 15,884 | 28,032 | 39,409 | 54,413 | 76,692 | 95,174 | 157,625 | 371,057 | 1,007,321 |
| 2005 | 9.662 | 46,041 | 9,486 | 16,089 | 27,826 | 39,232 | 54,519 | 77,338 | 96,326 | 162,219 | 397,859 | 1,135,101 |
| 2006 | 9.404 | 45,889 | 9,923 | 16,383 | 27,648 | 39,156 | 54,767 | 78,085 | 97,752 | 168,418 | 420,974 | 1,307,196 |
| 2007 | 9.308 | 46,664 | 10,493 | 16,930 | 27,686 | 39,013 | 54,829 | 78,846 | 99,425 | 173,991 | 454,312 | 1,330,240 |
| 2008 | 9.058 | 46,406 | 10,875 | 17,191 | 27,582 | 38,890 | 54,923 | 79,042 | 99,794 | 174,872 | 448,526 | 1,301,801 |
| 2009 | 8.851 | 46,016 | 10,128 | 16,533 | 26,994 | 38,473 | 54,675 | 79,374 | 100,190 | 173,638 | 423,921 | 1,166,691 |
| 2010 | 8.634 | 46,250 | 10,366 | 16,576 | 27,049 | 38,965 | 55,360 | 79,845 | 100,482 | 173,391 | 427,305 | 1,161,216 |
| 2011 | 8.688 | 47,027 | 11,069 | 17,385 | 27,464 | 39,235 | 56,177 | 80,965 | 102,446 | 179,416 | 448,392 | 1,232,365 |
| 2012 | 11.392 | 42,685 | 7,099 | 12,690 | 23,776 | 36,071 | 52,039 | 75,327 | 95,646 | 163,445 | 395,580 | 1,046,754 |
| 2013 | 11.782 | 42,345 | 6,684 | 11,869 | 23,239 | 35,774 | 51,846 | 75,096 | 95,361 | 162,571 | 399,861 | 1,089,920 |
| 2014 | 11.973 | 42,762 | 6,571 | 11,663 | 23,260 | 35,987 | 52,275 | 75,876 | 96,417 | 164,851 | 403,828 | 1,117,001 |
| 2015 | 12.150 | 43,114 | 6,680 | 11,789 | 23,339 | 36,139 | 52,830 | 76,832 | 97,749 | 168,104 | 421,757 | 1,193,978 |
| 2016 | 12.034 | 44,278 | 7,535 | 13,218 | 24,439 | 36,942 | 53,713 | 78,095 | 99,400 | 171,702 | 425,666 | 1,219,048 |

Notes: This table shows selected earnings percentiles for men in the combined IAB-TPP, the (imputed) IAB and TPP data. CS sample.

TABLE D.3: EARNINGS PERCENTILES IN IAB, TPP AND COMBINED IAB-TPP DATA – WOMEN

| Year | N | Mean | P5 | P10 | P25 | P50 | P75 | P90 | P95 | P99 | P99.9 | P99.99 |
|---------------------|--------|--------|-------|-------|--------|--------|--------|--------|--------|--------|---------|---------|
| IAB-TPP Data | | | | | | | | | | | | |
| 2001 | 11.476 | 24,636 | 3,671 | 4,673 | 10,881 | 22,162 | 34,723 | 46,032 | 54,306 | 78,212 | 137,387 | 284,602 |
| 2002 | 11.363 | 24,802 | 3,664 | 4,642 | 10,950 | 22,240 | 34,917 | 46,413 | 55,011 | 79,380 | 138,566 | 286,406 |
| 2003 | 11.166 | 24,905 | 3,630 | 4,692 | 10,759 | 22,288 | 35,129 | 46,840 | 55,509 | 79,964 | 137,893 | 273,567 |
| 2004 | 11.101 | 24,546 | 3,599 | 4,686 | 9,905 | 21,778 | 34,760 | 46,552 | 55,572 | 81,093 | 141,436 | 280,022 |
| 2005 | 10.965 | 24,434 | 3,596 | 4,646 | 9,736 | 21,580 | 34,540 | 46,490 | 55,603 | 81,674 | 145,817 | 299,067 |
| 2006 | 11.012 | 24,258 | 3,601 | 4,614 | 9,520 | 21,237 | 34,199 | 46,300 | 55,609 | 82,692 | 154,896 | 310,177 |
| 2007 | 11.146 | 24,088 | 3,653 | 4,686 | 9,602 | 20,907 | 33,722 | 45,957 | 55,521 | 84,019 | 158,477 | 327,358 |
| 2008 | 11.228 | 24,123 | 3,695 | 4,736 | 9,811 | 20,845 | 33,761 | 46,100 | 55,701 | 84,339 | 158,192 | 334,125 |
| 2009 | 11.223 | 24,370 | 3,706 | 4,745 | 9,897 | 21,013 | 34,147 | 46,845 | 56,361 | 85,317 | 158,036 | 324,891 |
| 2010 | 11.304 | 24,503 | 3,749 | 4,811 | 10,105 | 21,020 | 34,148 | 47,123 | 57,013 | 86,414 | 160,806 | 341,459 |
| 2011 | 11.439 | 24,554 | 3,791 | 4,812 | 10,376 | 20,987 | 34,034 | 47,034 | 57,110 | 87,208 | 165,112 | 360,119 |
| 2012 | 11.510 | 24,702 | 3,832 | 4,861 | 10,664 | 21,107 | 34,085 | 47,149 | 57,410 | 88,087 | 167,966 | 363,439 |
| 2013 | 11.585 | 24,953 | 3,868 | 4,994 | 10,913 | 21,353 | 34,345 | 47,430 | 57,738 | 88,991 | 169,977 | 366,792 |
| 2014 | 11.667 | 25,416 | 3,937 | 5,002 | 11,219 | 21,693 | 34,896 | 48,180 | 58,802 | 91,191 | 175,507 | 380,028 |
| 2015 | 11.756 | 26,038 | 4,090 | 5,218 | 11,846 | 22,188 | 35,487 | 48,998 | 59,914 | 93,356 | 181,516 | 404,868 |
| 2016 | 11.799 | 26,671 | 4,178 | 5,366 | 12,387 | 22,806 | 36,241 | 49,859 | 61,045 | 95,358 | 185,797 | 412,456 |
| IAB Data | | | | | | | | | | | | |
| 2001 | 11.476 | 24,558 | 3,676 | 4,678 | 10,886 | 22,167 | 34,728 | 46,037 | 54,311 | 77,237 | 126,590 | 185,188 |
| 2002 | 11.363 | 24,751 | 3,669 | 4,647 | 10,955 | 22,245 | 34,922 | 46,418 | 55,016 | 79,544 | 130,643 | 194,996 |
| 2003 | 11.166 | 24,823 | 3,635 | 4,697 | 10,764 | 22,293 | 35,134 | 46,845 | 55,514 | 77,428 | 126,331 | 187,049 |
| 2004 | 11.101 | 24,455 | 3,604 | 4,691 | 9,910 | 21,783 | 34,765 | 46,557 | 55,577 | 78,231 | 128,578 | 194,310 |
| 2005 | 10.965 | 24,334 | 3,601 | 4,651 | 9,741 | 21,585 | 34,545 | 46,495 | 55,608 | 78,982 | 132,035 | 197,873 |
| 2006 | 11.012 | 24,135 | 3,606 | 4,619 | 9,525 | 21,242 | 34,204 | 46,305 | 55,614 | 79,686 | 136,445 | 205,417 |
| 2007 | 11.146 | 23,954 | 3,658 | 4,691 | 9,607 | 20,912 | 33,727 | 45,962 | 55,526 | 80,683 | 140,641 | 222,722 |
| 2008 | 11.228 | 24,010 | 3,700 | 4,741 | 9,816 | 20,850 | 33,766 | 46,105 | 55,706 | 81,403 | 144,091 | 225,600 |
| 2009 | 11.223 | 24,248 | 3,711 | 4,750 | 9,902 | 21,018 | 34,152 | 46,850 | 56,366 | 82,021 | 144,067 | 226,262 |
| 2010 | 11.304 | 24,394 | 3,754 | 4,816 | 10,110 | 21,025 | 34,153 | 47,128 | 57,018 | 83,606 | 149,191 | 242,831 |
| 2011 | 11.439 | 24,380 | 3,796 | 4,817 | 10,381 | 20,992 | 34,039 | 47,039 | 57,115 | 83,009 | 145,068 | 224,170 |
| 2012 | 11.510 | 24,522 | 3,837 | 4,866 | 10,669 | 21,112 | 34,090 | 47,154 | 57,415 | 84,084 | 143,421 | 220,381 |
| 2013 | 11.585 | 24,744 | 3,873 | 4,999 | 10,918 | 21,358 | 34,350 | 47,435 | 57,743 | 83,587 | 143,554 | 220,791 |
| 2014 | 11.667 | 25,182 | 3,942 | 5,007 | 11,224 | 21,698 | 34,901 | 48,185 | 58,807 | 84,805 | 146,710 | 221,356 |
| 2015 | 11.756 | 25,758 | 4,095 | 5,223 | 11,851 | 22,193 | 35,492 | 49,003 | 59,919 | 85,858 | 149,172 | 219,736 |
| 2016 | 11.799 | 26,362 | 4,183 | 5,371 | 12,392 | 22,811 | 36,246 | 49,864 | 61,022 | 86,977 | 148,851 | 226,320 |
| TPP Data | | | | | | | | | | | | |
| 2001 | 7.979 | 28,866 | 5,536 | 8,757 | 15,945 | 25,605 | 37,069 | 48,836 | 57,770 | 83,738 | 152,169 | 325,486 |
| 2002 | 8.316 | 28,430 | 5,548 | 8,769 | 15,966 | 25,674 | 37,366 | 49,393 | 58,579 | 84,761 | 153,326 | 314,386 |
| 2003 | 8.040 | 28,607 | 5,744 | 8,885 | 16,071 | 25,829 | 37,775 | 49,892 | 59,066 | 85,574 | 153,277 | 307,138 |
| 2004 | 7.870 | 28,541 | 5,826 | 8,709 | 15,878 | 25,622 | 37,624 | 49,906 | 59,557 | 87,223 | 156,735 | 319,596 |
| 2005 | 7.624 | 29,001 | 5,848 | 8,708 | 15,852 | 25,511 | 37,520 | 49,852 | 59,595 | 88,426 | 163,486 | 351,511 |
| 2006 | 7.468 | 28,433 | 5,828 | 8,671 | 15,703 | 25,267 | 37,282 | 49,710 | 59,685 | 89,720 | 175,841 | 360,206 |
| 2007 | 7.480 | 28,201 | 5,817 | 8,594 | 15,493 | 24,901 | 36,816 | 49,290 | 59,671 | 91,234 | 177,866 | 378,985 |
| 2008 | 7.409 | 28,089 | 5,752 | 8,514 | 15,279 | 24,618 | 36,667 | 49,015 | 59,452 | 91,185 | 175,199 | 384,082 |
| 2009 | 7.363 | 28,338 | 5,736 | 8,492 | 15,297 | 24,753 | 37,174 | 50,068 | 60,632 | 92,857 | 177,928 | 369,099 |
| 2010 | 7.330 | 28,402 | 5,731 | 8,489 | 15,242 | 24,746 | 37,242 | 50,487 | 61,074 | 93,594 | 178,933 | 391,947 |
| 2011 | 7.409 | 28,402 | 5,737 | 8,501 | 15,224 | 24,658 | 37,103 | 50,513 | 61,299 | 94,831 | 184,320 | 414,964 |
| 2012 | 9.528 | 27,466 | 5,186 | 7,555 | 13,978 | 23,270 | 35,739 | 48,838 | 59,463 | 91,641 | 177,184 | 386,224 |
| 2013 | 9.883 | 27,137 | 5,155 | 7,524 | 13,957 | 23,319 | 35,879 | 49,124 | 59,839 | 92,625 | 178,170 | 390,722 |
| 2014 | 10.043 | 28,046 | 5,139 | 7,531 | 14,166 | 23,667 | 36,405 | 49,889 | 60,877 | 94,553 | 183,884 | 398,124 |
| 2015 | 10.225 | 28,419 | 5,327 | 7,788 | 14,531 | 23,923 | 36,819 | 50,539 | 61,826 | 96,454 | 188,698 | 428,613 |
| 2016 | 10.108 | 28,836 | 5,678 | 8,374 | 15,194 | 24,784 | 37,796 | 51,689 | 63,149 | 98,685 | 193,986 | 441,093 |

Notes: This table shows selected earnings percentiles for women in the combined IAB-TPP, the (imputed) IAB and TPP data. CS sample.

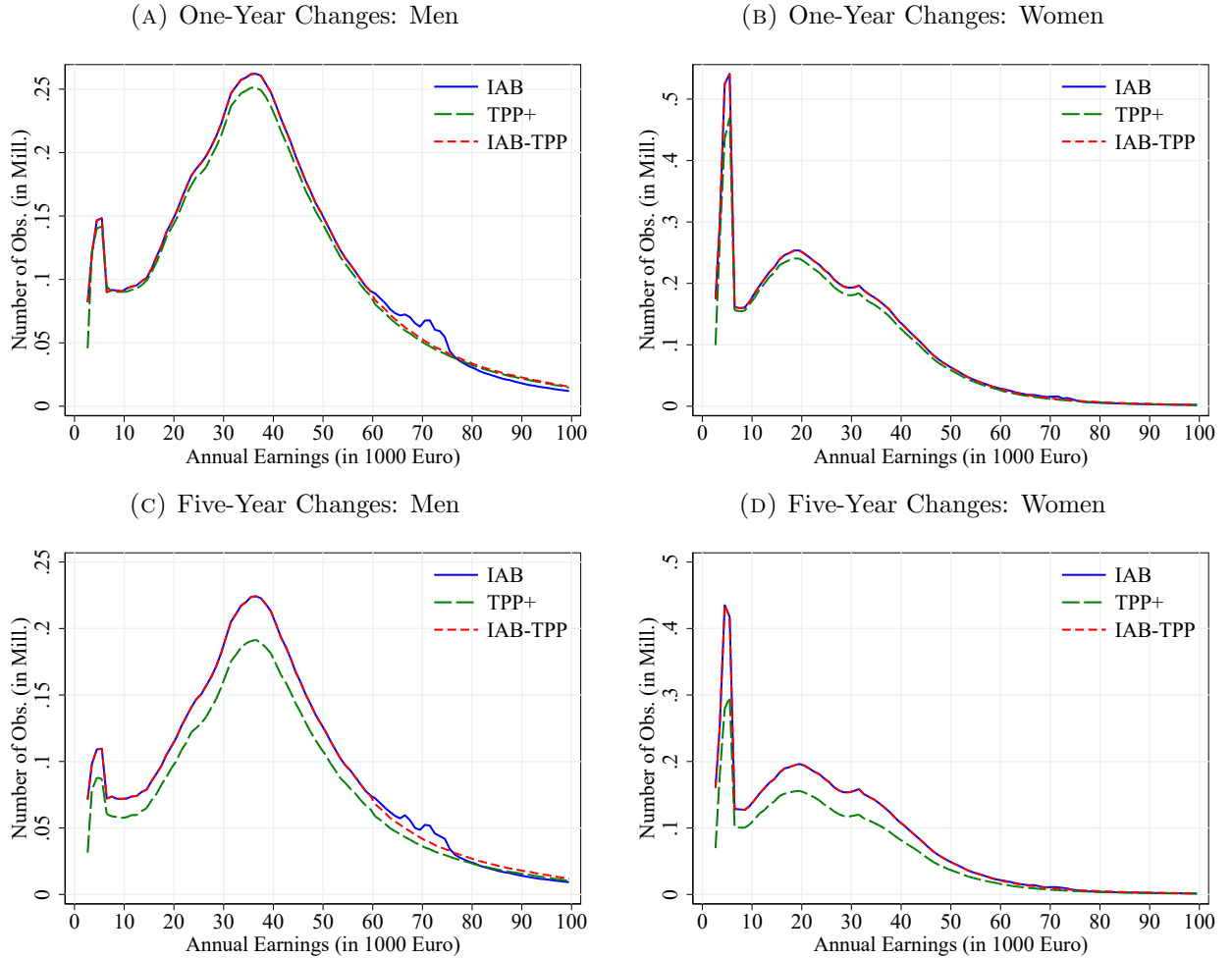
TABLE D.4: EARNINGS PERCENTILES IN IAB, TPP AND COMBINED IAB-TPP DATA – POPULATION

| Year | N | Mean | P5 | P10 | P25 | P50 | P75 | P90 | P95 | P99 | P99.9 | P99.99 |
|---------------------|--------|--------|-------|--------|--------|--------|--------|--------|--------|---------|---------|-----------|
| IAB-TPP Data | | | | | | | | | | | | |
| 2001 | 24,588 | 33,678 | 4,291 | 5,698 | 16,757 | 31,007 | 43,605 | 59,612 | 74,328 | 119,431 | 265,535 | 699,853 |
| 2002 | 24,185 | 33,765 | 4,253 | 5,613 | 16,587 | 30,994 | 43,835 | 60,284 | 75,186 | 119,907 | 263,680 | 680,196 |
| 2003 | 23,759 | 33,865 | 4,208 | 5,668 | 16,439 | 31,012 | 44,169 | 60,878 | 76,037 | 120,881 | 258,478 | 618,882 |
| 2004 | 23,520 | 33,542 | 4,127 | 5,652 | 15,746 | 30,440 | 43,860 | 60,900 | 76,427 | 122,138 | 267,480 | 668,308 |
| 2005 | 23,130 | 33,484 | 4,086 | 5,594 | 15,559 | 30,030 | 43,671 | 60,959 | 76,877 | 124,474 | 282,478 | 758,804 |
| 2006 | 23,226 | 33,389 | 4,105 | 5,585 | 15,249 | 29,517 | 43,438 | 61,108 | 77,446 | 126,999 | 297,407 | 838,076 |
| 2007 | 23,519 | 33,323 | 4,179 | 5,445 | 15,173 | 29,081 | 43,079 | 61,036 | 77,864 | 129,829 | 309,079 | 879,187 |
| 2008 | 23,657 | 33,298 | 4,226 | 5,409 | 15,179 | 28,901 | 43,053 | 61,096 | 77,945 | 130,047 | 310,816 | 848,239 |
| 2009 | 23,445 | 33,168 | 4,188 | 5,394 | 14,975 | 28,794 | 43,109 | 61,032 | 78,017 | 130,187 | 298,103 | 772,990 |
| 2010 | 23,579 | 33,189 | 4,210 | 5,335 | 14,946 | 28,566 | 43,243 | 61,487 | 78,355 | 130,007 | 297,627 | 776,410 |
| 2011 | 23,903 | 33,359 | 4,212 | 5,235 | 15,195 | 28,489 | 43,136 | 61,841 | 79,018 | 132,538 | 309,539 | 834,160 |
| 2012 | 24,044 | 33,487 | 4,258 | 5,419 | 15,378 | 28,525 | 43,165 | 62,230 | 79,513 | 133,125 | 308,711 | 807,367 |
| 2013 | 24,224 | 33,605 | 4,332 | 5,677 | 15,420 | 28,600 | 43,327 | 62,347 | 79,685 | 133,218 | 311,903 | 811,185 |
| 2014 | 24,463 | 33,994 | 4,391 | 5,638 | 15,588 | 28,808 | 43,806 | 63,182 | 80,822 | 135,321 | 313,688 | 829,623 |
| 2015 | 24,715 | 34,613 | 4,533 | 6,059 | 16,169 | 29,148 | 44,362 | 64,190 | 82,159 | 138,285 | 325,526 | 898,376 |
| 2016 | 24,895 | 35,155 | 4,629 | 6,371 | 16,599 | 29,688 | 44,908 | 64,980 | 83,158 | 140,084 | 329,513 | 906,087 |
| IAB Data | | | | | | | | | | | | |
| 2001 | 24,588 | 33,210 | 4,296 | 5,703 | 16,762 | 31,012 | 43,610 | 59,617 | 71,786 | 116,085 | 211,082 | 344,387 |
| 2002 | 24,185 | 33,372 | 4,258 | 5,618 | 16,592 | 30,999 | 43,840 | 60,285 | 73,015 | 118,395 | 212,391 | 341,753 |
| 2003 | 23,759 | 33,372 | 4,213 | 5,673 | 16,444 | 31,017 | 44,174 | 60,873 | 74,597 | 114,596 | 201,486 | 311,873 |
| 2004 | 23,520 | 33,030 | 4,132 | 5,657 | 15,751 | 30,445 | 43,865 | 60,885 | 74,394 | 116,687 | 207,185 | 323,429 |
| 2005 | 23,130 | 32,920 | 4,091 | 5,599 | 15,564 | 30,035 | 43,676 | 60,944 | 74,251 | 119,296 | 212,988 | 337,394 |
| 2006 | 23,226 | 32,772 | 4,110 | 5,590 | 15,254 | 29,522 | 43,443 | 61,079 | 74,469 | 121,980 | 221,301 | 352,315 |
| 2007 | 23,519 | 32,667 | 4,184 | 5,558 | 15,178 | 29,086 | 43,084 | 61,013 | 74,345 | 124,768 | 232,408 | 369,988 |
| 2008 | 23,657 | 32,706 | 4,231 | 5,414 | 15,184 | 28,906 | 43,058 | 61,037 | 74,385 | 126,856 | 238,196 | 379,414 |
| 2009 | 23,445 | 32,595 | 4,193 | 5,398 | 14,980 | 28,799 | 43,114 | 60,976 | 74,260 | 125,714 | 233,515 | 376,841 |
| 2010 | 23,579 | 32,711 | 4,215 | 5,340 | 14,951 | 28,571 | 43,248 | 61,399 | 74,985 | 128,004 | 240,624 | 388,516 |
| 2011 | 23,903 | 32,753 | 4,217 | 5,232 | 15,200 | 28,494 | 43,141 | 61,746 | 75,041 | 128,224 | 239,371 | 382,640 |
| 2012 | 24,044 | 32,851 | 4,263 | 5,424 | 15,383 | 28,530 | 43,170 | 62,025 | 75,396 | 128,269 | 234,306 | 365,760 |
| 2013 | 24,224 | 32,885 | 4,337 | 5,682 | 15,425 | 28,605 | 43,332 | 62,090 | 74,753 | 127,146 | 230,019 | 353,749 |
| 2014 | 24,463 | 33,257 | 4,396 | 5,642 | 15,593 | 28,813 | 43,811 | 62,851 | 75,566 | 128,792 | 236,544 | 369,246 |
| 2015 | 24,715 | 33,754 | 4,538 | 6,064 | 16,174 | 29,153 | 44,367 | 63,866 | 76,383 | 129,541 | 236,768 | 371,329 |
| 2016 | 24,895 | 34,245 | 4,634 | 6,376 | 16,604 | 29,693 | 44,913 | 64,684 | 77,146 | 130,054 | 238,211 | 374,076 |
| TPP Data | | | | | | | | | | | | |
| 2001 | 18,549 | 37,871 | 7,288 | 11,741 | 21,331 | 33,960 | 46,672 | 64,861 | 80,946 | 131,421 | 302,132 | 821,165 |
| 2002 | 19,039 | 37,687 | 7,029 | 11,465 | 20,961 | 33,768 | 46,693 | 65,048 | 81,171 | 130,897 | 297,634 | 788,232 |
| 2003 | 18,436 | 37,897 | 7,094 | 11,442 | 21,000 | 33,939 | 47,116 | 65,627 | 81,964 | 131,644 | 288,229 | 706,906 |
| 2004 | 17,927 | 37,932 | 6,964 | 11,249 | 20,738 | 33,722 | 47,070 | 65,987 | 82,740 | 133,523 | 302,414 | 772,816 |
| 2005 | 17,286 | 38,013 | 7,039 | 11,344 | 20,663 | 33,440 | 46,989 | 66,238 | 83,462 | 136,557 | 318,894 | 875,066 |
| 2006 | 16,872 | 38,800 | 7,107 | 11,371 | 20,529 | 33,227 | 47,001 | 66,601 | 84,488 | 140,408 | 336,207 | 972,140 |
| 2007 | 16,788 | 38,227 | 7,158 | 11,445 | 20,451 | 32,925 | 46,776 | 66,902 | 85,430 | 144,138 | 356,608 | 1,015,749 |
| 2008 | 16,467 | 38,078 | 7,177 | 11,406 | 20,309 | 32,724 | 46,565 | 66,887 | 85,427 | 144,384 | 354,247 | 997,556 |
| 2009 | 16,214 | 37,918 | 7,013 | 11,148 | 20,108 | 32,524 | 46,607 | 67,244 | 85,755 | 144,617 | 339,340 | 867,527 |
| 2010 | 15,964 | 38,053 | 7,027 | 11,138 | 19,993 | 32,586 | 47,056 | 67,588 | 86,058 | 144,271 | 340,943 | 910,341 |
| 2011 | 16,097 | 38,448 | 7,167 | 11,353 | 20,171 | 32,691 | 47,327 | 68,238 | 87,254 | 147,770 | 358,185 | 944,551 |
| 2012 | 20,921 | 36,136 | 5,769 | 9,301 | 17,949 | 30,315 | 44,495 | 64,207 | 81,958 | 137,468 | 321,052 | 831,895 |
| 2013 | 21,665 | 35,403 | 5,693 | 9,082 | 17,716 | 30,077 | 44,438 | 64,153 | 81,904 | 136,935 | 323,332 | 844,378 |
| 2014 | 22,016 | 36,293 | 5,713 | 9,076 | 17,846 | 30,300 | 44,870 | 64,871 | 82,915 | 138,998 | 325,341 | 869,572 |
| 2015 | 22,375 | 36,187 | 5,822 | 9,286 | 18,255 | 30,458 | 45,305 | 65,710 | 84,056 | 141,618 | 335,844 | 921,056 |
| 2016 | 22,142 | 37,560 | 6,263 | 10,069 | 19,098 | 31,344 | 46,237 | 66,867 | 85,536 | 144,316 | 342,562 | 946,374 |

Notes: This table shows selected earnings percentiles for men and women in the combined IAB-TPP, the (imputed) IAB and TPP data. CS sample.

For the LS samples, we follow the same procedure. While the cross-sectional earnings distribution in the reweighted TPP data matches the IAB data (by construction of the weights), this is no longer the case for the LS and H samples due to attrition in the TPP.¹¹ The LS samples differ from the CS sample in that workers have to be in the data in year t and $t + 1$ or $t + 5$. Figure D.2 shows the earnings distribution in these samples in the IAB, the reweighted TPP and the combined IAB-TPP data. The attrition in the reweighted TPP data becomes particularly visible in Panels C and D which plots the earnings distribution in the LS sample for 5-year earnings changes.

FIGURE D.2: IAB vs. TPP: EARNINGS DISTRIBUTION IN LONGITUDINAL SAMPLES



Notes: LS sample. Annual earnings. Averaged over years 2001-2015 for one-year changes and 2001-2011 for five-year changes. Source: IAB and TPP.

D.3 Combined Earnings Growth Distribution

For the analysis of earnings dynamics, we are interested in the distribution of earnings growth, i.e. the distribution of earnings *changes* in addition to the earnings distribution shown in Figure D.2

¹¹Recall, that the reweighting does not target moments of earnings changes over time. For example, many workers who switch from a regular job to a mini-job will drop out of the TPP.

for the different samples. To construct this distribution of changes, we proceed as follows. For simplicity, we drop time subscripts for all variables and use the following notation:

- earnings y (continuous)
- earnings bins Y (discrete and finite support)
- earnings growth $g = \log(y_{t+k}) - \log(y)$ (continuous)
- earnings growth bins G (discrete and finite support)

Available Data. For each data source (IAB and reweighted TPP) and by year and gender, we have

- the share of workers in each earnings bin: $\Pr(Y)$
- summary statistics (e.g. mean, standard deviation, skewness, kurtosis) and selected quantiles¹² of earnings growth by earnings bin:

$$q^p(g|Y) \equiv F_{g|Y}^{-1}(p/100|Y) \quad \text{for selected values of } q \in (0, 1) \quad (\text{D.15})$$

Conditional Growth Rate Distributions by Earnings Bins. In a first step, we approximate the conditional CDF of earnings growth, $F_{g|Y}$ in both the IAB and reweighted TPP data using a continuous interpolation of its quantiles.¹³

Figures D.4 shows the P90-P10 differential, Kelley Skewness and Excess Crow-Siddiqui kurtosis of 1-year earnings growth by current earnings in the IAB and reweighted TPP data. In the middle of the earnings distribution, the conditional earnings growth distributions are very similar in the IAB and reweighted TPP data. However, there are stark differences at the bottom (where the TPP has a lot of attrition because of missing mini-jobs) and even more so above the top-coding threshold where imputed earnings in the IAB are essentially iid. Figure D.5 shows the corresponding statistics for 5-year earnings growth. While the IAB and (reweighted) TPP are again remarkably similar in the middle of the male earnings distribution, the fit becomes slightly worse for women.

In order to construct the combined IAB-TPP data set for earnings changes, we proceed as follows. First, for high (above-cutoff) initial earnings bins, we use the conditional earnings growth distribution from the TPP as the entire conditional distribution of earnings growth in the IAB is affected by top-coding. Figure D.6 plots the share of 1- and 5-year log earnings growth rates affected by top-coding in the IAB. For initial-earnings groups below the top-coding cutoff, only (the right) part of the conditional growth distribution in the IAB is affected by top-coding. For earnings bins where more than 50% of future earnings are top-coded, we use the earnings-growth distribution from the TPP. If less than 50% are top-coded in the IAB, we replace the top-coded earnings-growth

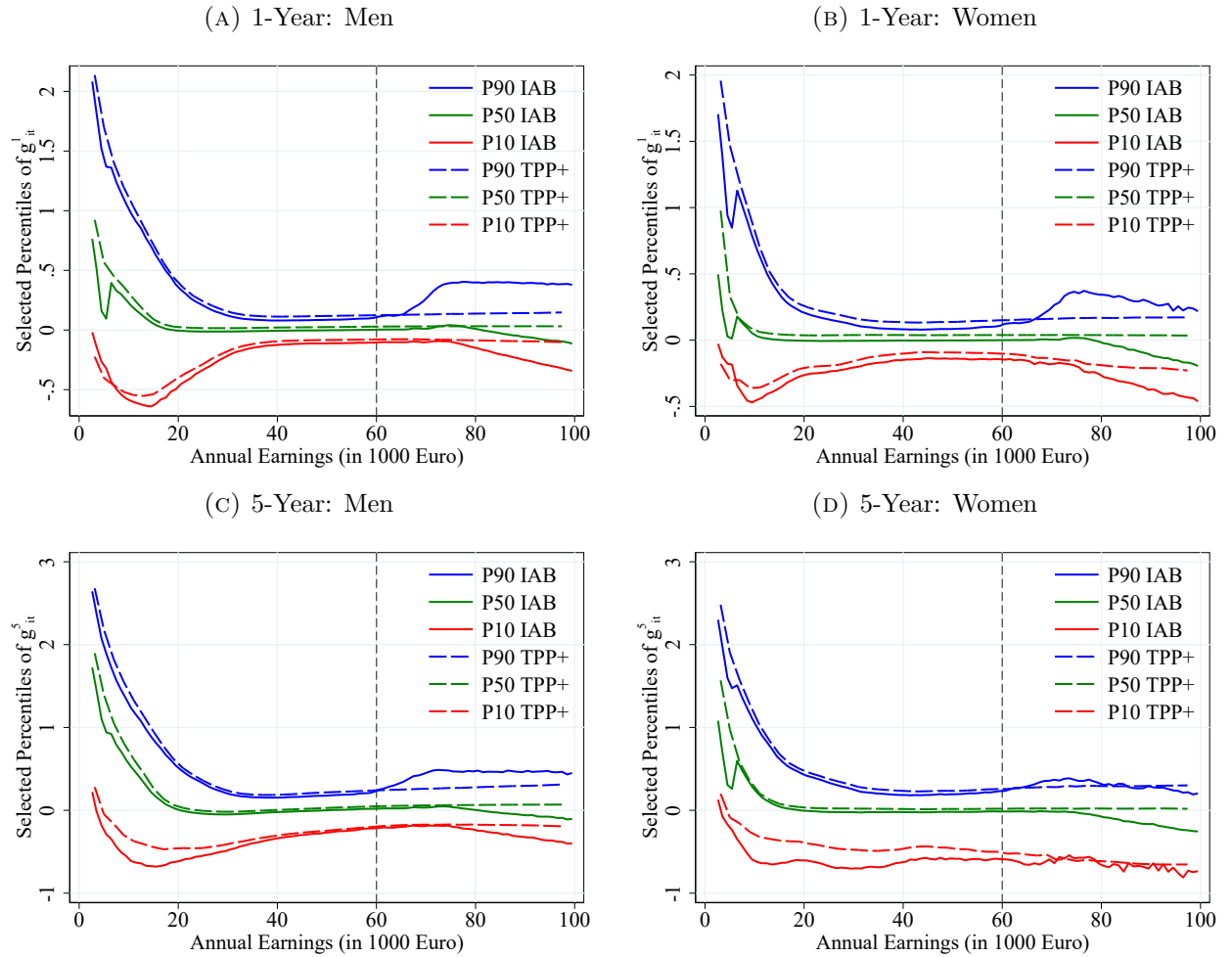
¹²We have the following percentiles: 0.1, 0.5, 1, 2, ..., 10, 15, ..., 90, 91, ..., 99, 99.5, 99.9.

¹³In order to approximate the CDF using monotonic spline interpolation of quantiles in a given dataset, we have to impose a minimum and maximum for g , i.e. q^0 and q^{100} . Let $\hat{F}_{g|Y}$ denote the resulting approximation of the CDF of earnings growth. We set the minimum and the maximum such that the standard deviation and skewness of $\hat{F}_{g|Y}$ equal the values that we observe in the data.

quantiles in the IAB with a re-centered counterpart from the TPP data. For example, if the top 10% of the earnings growth distribution is affected by top-coding, we use the 85th percentile from the IAB and the 95th percentile from the TPP and subtract from it the difference in the 50th percentile between TPP and IAB to account for the fact that the TPP distribution is slightly upward biased (see Figure D.3). That is, the p -quantile of the earnings growth distribution in bin G where a share s of growth rates is top-coded, is given by:

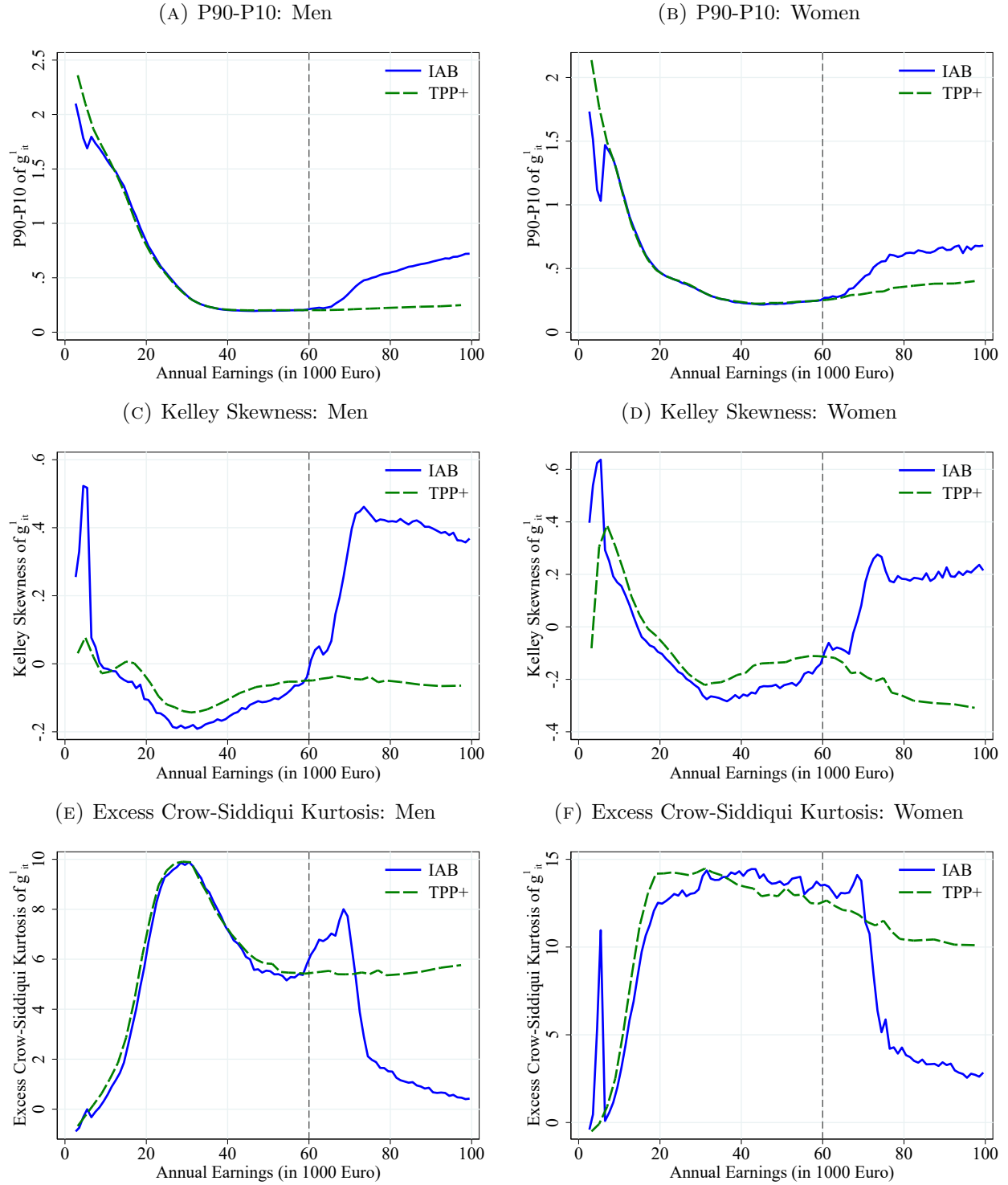
$$q^p(g|G) = \begin{cases} q_{IAB}^p(g|G) & \text{if } s < 0.5 \text{ and } p \leq s \\ q_{TPP}^p(g|G) - (q_{TPP}^{0.5}(g|G) - q_{IAB}^{0.5}(g|G)) & \text{if } s < 0.5 \text{ and } p > s \\ q_{TPP}^p(g|G) & \text{if } s > 0.5 \end{cases} \quad (D.16)$$

FIGURE D.3: IAB vs. TPP: PERCENTILES OF LOG EARNINGS CHANGES BY CURRENT EARNINGS



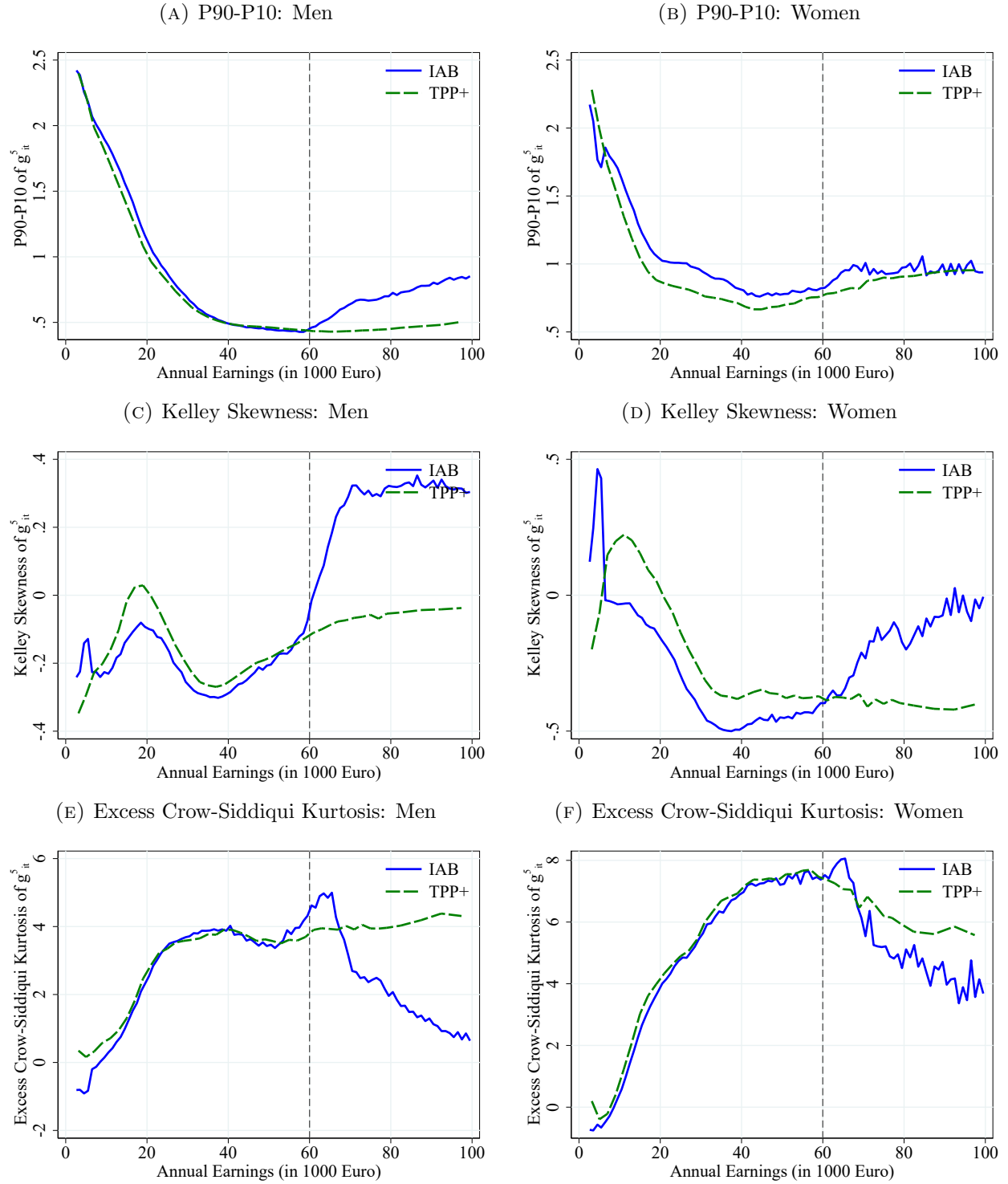
Notes: LS sample. Selected percentiles of 1- and 5-year residualized log earnings growth distribution in IAB and (re-weighted) TPP data. Averaged over years 2001-2015 for 1-year growth rates and over 2001-2011 for 5-year growth rates. Source: IAB and TPP.

FIGURE D.4: IAB vs. TPP: 1-YEAR LOG EARNINGS CHANGES BY CURRENT EARNINGS



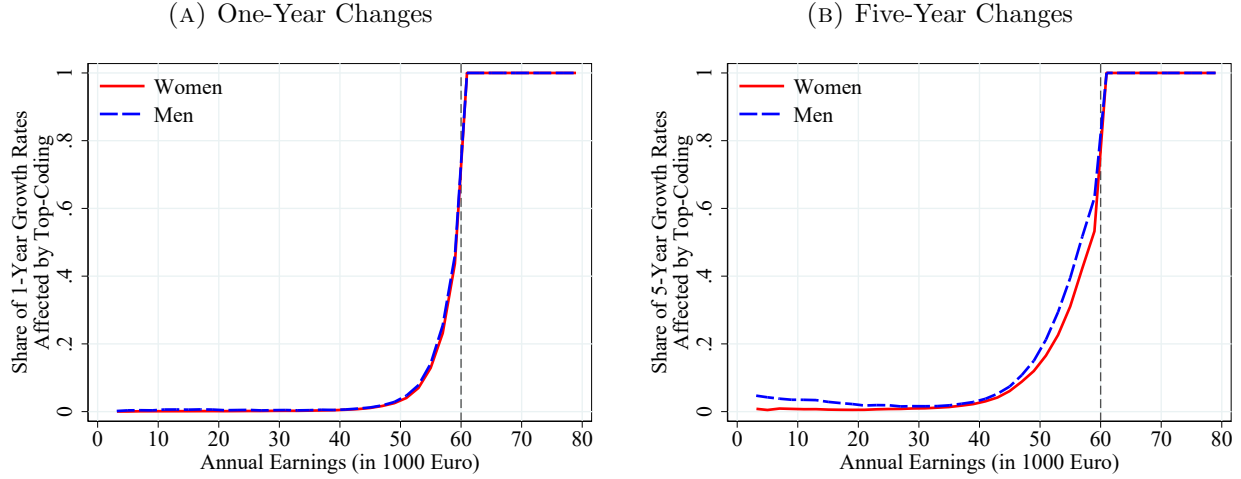
Notes: LS sample. One-year residualized log earnings growth. Averaged over years 2001-2015. Source: IAB and TPP.

FIGURE D.5: IAB vs. TPP: 5-YEAR LOG EARNINGS CHANGES BY CURRENT EARNINGS



Notes: LS sample. Five-year residualized log earnings growth. Averaged over years 2001-2011. Source: IAB and TPP.

FIGURE D.6: SHARE OF 1-YEAR LOG EARNINGS CHANGES AFFECTED BY TOP-CODING IN THE IAB DATA



Notes: LS sample. Averages over years. Earnings changes are affected by top-coding if current or future earnings are above 60,000 Euro. The dashed vertical depicts the point where 2% of earnings changes are affected by top-coding. Source: IAB.

In the next step, we discretize the continuous conditional earnings growth distributions. To do so, we set up a fine grid for g ranging from the global minimum to the global maximum of the support of $\hat{F}_{g|Y}$. The grid defines earnings growth bins G with upper and lower bounds denoted by G^+ and G^- respectively. Using those, we discretize the continuous conditional distributions to obtain $\Pr(G|Y)$ for all G and Y :

$$\Pr(G|Y) = \Pr(G^- \leq g \leq G^+|Y) = \hat{F}_{g|Y}(G^+|Y) - \hat{F}_{g|Y}(G^-|Y) \quad (\text{D.17})$$

Unconditional Growth Rate Distribution. Finally, this discretized conditional growth distribution allows us to recover the (unconditional) marginal probability mass function of earnings growth (discretized) defined by the probabilities

$$\Pr(G) = \sum_Y \Pr(G, Y) = \sum_Y \Pr(G|Y) \Pr(Y) \quad (\text{D.18})$$

where $\Pr(Y)$ is the discretized combined IAB-TPP earnings distribution in the corresponding LS sample (see Figure D.2). As the bins are very fine, we simply use their midpoints along with the above probabilities to compute summary statistics and selected percentiles of the unconditional distribution of earnings growth. Tables D.6 and D.7 show selected percentiles of the 1-year earnings growth distribution using the combined IAB-TPP data as well as the IAB and (reweighted) TPP data. Tables D.9 and D.10 show the corresponding statistics for 5-year earnings growth.

TABLE D.5: PERCENTILES OF REAL ANNUAL EARNINGS – LS SAMPLE WITH 1-YEAR CHANGES

| Year | N | Mean | P5 | P10 | P25 | P50 | P75 | P90 | P95 | P99 | P99.9 | P99.99 |
|--------------|--------|--------|-------|--------|--------|--------|--------|--------|--------|---------|---------|---------|
| Men | | | | | | | | | | | | |
| 2001 | 11.912 | 43,190 | 8,791 | 15,719 | 27,982 | 38,659 | 51,532 | 70,995 | 88,108 | 143,467 | 332,694 | 919,859 |
| 2005 | 11.179 | 43,016 | 6,707 | 13,234 | 26,104 | 38,036 | 52,112 | 73,334 | 91,449 | 149,354 | 355,640 | 980,704 |
| 2010 | 11.379 | 42,171 | 6,183 | 11,844 | 23,742 | 36,483 | 51,748 | 74,004 | 93,290 | 155,921 | 368,195 | 994,093 |
| Women | | | | | | | | | | | | |
| 2001 | 10.214 | 25,858 | 4,056 | 4,891 | 12,765 | 23,666 | 35,769 | 46,942 | 55,245 | 79,670 | 139,035 | 290,365 |
| 2005 | 9.874 | 25,434 | 3,951 | 4,992 | 11,545 | 22,757 | 35,423 | 47,268 | 56,446 | 82,755 | 145,573 | 296,609 |
| 2010 | 10.253 | 25,279 | 4,080 | 5,095 | 11,241 | 21,930 | 34,863 | 47,791 | 57,719 | 87,235 | 159,738 | 335,921 |

Notes: This table shows the number of observations (in millions) and selected percentiles of real annual earnings (in 2018 Euro) in the combined IAB-TPP data. LS sample with non-missing 1-year log earnings changes (from t to $t + 1$). Sources: IAB and TPP.

TABLE D.6: PERCENTILES OF 1-YEAR EARNINGS GROWTH IN COMBINED IAB-TPP DATA – MEN

| Year | N | P1 | P2.5 | P10 | P25 | P50 | P75 | P90 | P95 | P97.5 | P99 |
|---------------------|--------|--------|--------|--------|--------|--------|-------|-------|-------|-------|-------|
| IAB-TPP Data | | | | | | | | | | | |
| 2001 | 11.912 | -1.579 | -1.044 | -0.235 | -0.062 | -0.008 | 0.045 | 0.201 | 0.497 | 0.869 | 1.358 |
| 2005 | 11.179 | -1.376 | -0.816 | -0.165 | -0.053 | -0.005 | 0.053 | 0.228 | 0.563 | 0.954 | 1.395 |
| 2010 | 11.379 | -1.261 | -0.730 | -0.168 | -0.058 | -0.006 | 0.068 | 0.282 | 0.619 | 0.985 | 1.424 |
| IAB Data | | | | | | | | | | | |
| 2001 | 11.912 | -1.571 | -1.052 | -0.294 | -0.067 | -0.006 | 0.053 | 0.270 | 0.563 | 0.901 | 1.370 |
| 2005 | 11.179 | -1.365 | -0.813 | -0.204 | -0.057 | -0.004 | 0.059 | 0.280 | 0.593 | 0.946 | 1.396 |
| 2010 | 11.379 | -1.259 | -0.764 | -0.212 | -0.064 | -0.006 | 0.076 | 0.342 | 0.660 | 1.001 | 1.433 |
| TPP+ Data | | | | | | | | | | | |
| 2001 | 9.694 | -1.678 | -1.032 | -0.215 | -0.041 | 0.013 | 0.064 | 0.244 | 0.583 | 1.006 | 1.484 |
| 2005 | 8.413 | -1.455 | -0.801 | -0.135 | -0.031 | 0.012 | 0.069 | 0.270 | 0.657 | 1.088 | 1.556 |
| 2010 | 7.791 | -1.257 | -0.661 | -0.128 | -0.033 | 0.014 | 0.095 | 0.349 | 0.772 | 1.205 | 1.653 |

Notes: This table shows the number of observations (in millions) and selected percentiles of the combined IAB-TPP distribution of 1-year changes in residualized log earnings (from t to $t + 1$) for men and selected years. LS sample. Sources: IAB and TPP.

TABLE D.7: Percentiles of 1-Year Earnings Growth in Combined IAB-TPP Data – Women

| Year | N | P1 | P2.5 | P10 | P25 | P50 | P75 | P90 | P95 | P97.5 | P99 |
|---------------------|--------|--------|--------|--------|--------|--------|-------|-------|-------|-------|-------|
| IAB-TPP Data | | | | | | | | | | | |
| 2001 | 11.912 | -1.579 | -1.044 | -0.235 | -0.062 | -0.008 | 0.045 | 0.201 | 0.497 | 0.869 | 1.358 |
| 2005 | 11.179 | -1.376 | -0.816 | -0.165 | -0.053 | -0.005 | 0.053 | 0.228 | 0.563 | 0.954 | 1.395 |
| 2010 | 11.379 | -1.261 | -0.730 | -0.168 | -0.058 | -0.006 | 0.068 | 0.282 | 0.619 | 0.985 | 1.424 |
| IAB Data | | | | | | | | | | | |
| 2001 | 11.912 | -1.571 | -1.052 | -0.294 | -0.067 | -0.006 | 0.053 | 0.270 | 0.563 | 0.901 | 1.370 |
| 2005 | 11.179 | -1.365 | -0.813 | -0.204 | -0.057 | -0.004 | 0.059 | 0.280 | 0.593 | 0.946 | 1.396 |
| 2010 | 11.379 | -1.259 | -0.764 | -0.212 | -0.064 | -0.006 | 0.076 | 0.342 | 0.660 | 1.001 | 1.433 |
| TPP+ Data | | | | | | | | | | | |
| 2001 | 9.694 | -1.678 | -1.032 | -0.215 | -0.041 | 0.013 | 0.064 | 0.244 | 0.583 | 1.006 | 1.484 |
| 2005 | 8.413 | -1.455 | -0.801 | -0.135 | -0.031 | 0.012 | 0.069 | 0.270 | 0.657 | 1.088 | 1.556 |
| 2010 | 7.791 | -1.257 | -0.661 | -0.128 | -0.033 | 0.014 | 0.095 | 0.349 | 0.772 | 1.205 | 1.653 |

Notes: This table shows the number of observations (in millions) and selected percentiles of the combined IAB-TPP distribution of 1-year changes in residualized log earnings (from t to $t+1$) for women and selected years. LS sample. Sources: IAB and TPP.

TABLE D.8: PERCENTILES OF REAL ANNUAL EARNINGS – LS SAMPLE WITH 5-YEAR CHANGES

| Year | N | Mean | P5 | P10 | P25 | P50 | P75 | P90 | P95 | P99 | P99.9 | P99.99 |
|--------------|-------|--------|-------|--------|--------|--------|--------|--------|--------|---------|---------|---------|
| Men | | | | | | | | | | | | |
| 2001 | 9.599 | 43,175 | 9,561 | 16,990 | 28,791 | 38,951 | 51,505 | 70,347 | 86,758 | 135,055 | 295,389 | 831,361 |
| 2005 | 9.189 | 42,397 | 6,782 | 13,277 | 26,222 | 37,984 | 51,743 | 72,300 | 89,449 | 139,269 | 305,821 | 836,836 |
| 2010 | 9.253 | 41,383 | 6,169 | 11,788 | 23,680 | 36,211 | 51,149 | 72,478 | 90,820 | 147,462 | 330,275 | 873,678 |
| Women | | | | | | | | | | | | |
| 2001 | 7.982 | 25,977 | 4,039 | 4,903 | 12,973 | 23,888 | 35,891 | 46,956 | 55,198 | 79,399 | 136,241 | 277,011 |
| 2005 | 7.899 | 25,203 | 3,845 | 4,918 | 11,282 | 22,560 | 35,182 | 46,965 | 56,099 | 82,304 | 141,291 | 284,028 |
| 2010 | 8.119 | 25,056 | 3,999 | 5,043 | 11,015 | 21,751 | 34,595 | 47,442 | 57,398 | 86,895 | 156,575 | 321,349 |

Notes: This table shows the number of observations (in millions) and selected percentiles of real annual earnings (in 2018 Euro) in the combined IAB-TPP data. LS sample with non-missing 5-year log earnings changes (from t to $t+5$). Sources: IAB and TPP.

TABLE D.9: PERCENTILES OF 5-YEAR EARNINGS GROWTH IN COMBINED IAB-TPP DATA – MEN

| Year | N | P1 | P2.5 | P10 | P25 | P50 | P75 | P90 | P95 | P97.5 | P99 |
|---------------------|-------|--------|--------|--------|--------|--------|-------|-------|-------|-------|-------|
| IAB-TPP Data | | | | | | | | | | | |
| 2001 | 9.599 | -1.990 | -1.402 | -0.440 | -0.152 | -0.011 | 0.109 | 0.344 | 0.725 | 1.213 | 1.790 |
| 2005 | 9.189 | -1.808 | -1.180 | -0.381 | -0.130 | 0.004 | 0.133 | 0.465 | 0.931 | 1.410 | 1.932 |
| 2010 | 9.253 | -1.646 | -1.013 | -0.340 | -0.109 | 0.025 | 0.178 | 0.554 | 1.023 | 1.472 | 1.967 |
| IAB Data | | | | | | | | | | | |
| 2001 | 9.599 | -1.974 | -1.378 | -0.452 | -0.160 | -0.012 | 0.115 | 0.386 | 0.752 | 1.203 | 1.794 |
| 2005 | 9.189 | -1.777 | -1.155 | -0.391 | -0.139 | 0.001 | 0.139 | 0.512 | 0.950 | 1.420 | 1.941 |
| 2010 | 9.253 | -1.631 | -1.018 | -0.369 | -0.128 | 0.012 | 0.172 | 0.582 | 1.024 | 1.479 | 1.977 |
| TPP+ Data | | | | | | | | | | | |
| 2001 | 6.531 | -2.026 | -1.276 | -0.372 | -0.123 | 0.007 | 0.123 | 0.379 | 0.811 | 1.331 | 1.898 |
| 2005 | 5.708 | -1.835 | -1.090 | -0.340 | -0.106 | 0.023 | 0.148 | 0.478 | 0.976 | 1.488 | 1.992 |
| 2010 | 6.383 | -1.713 | -0.975 | -0.299 | -0.085 | 0.040 | 0.195 | 0.616 | 1.161 | 1.659 | 2.130 |

Notes: This table shows the number of observations (in millions) and selected percentiles of the combined IAB-TPP distribution of 5-year changes in residualized log earnings (from t to $t + 5$) for men and selected years. LS sample. Sources: IAB and TPP.

TABLE D.10: Percentiles of 5-Year Earnings Growth in Combined IAB-TPP Data – Women

| Year | N | P1 | P2.5 | P10 | P25 | P50 | P75 | P90 | P95 | P97.5 | P99 |
|---------------------|--------|--------|--------|--------|--------|--------|-------|-------|-------|-------|-------|
| IAB-TPP Data | | | | | | | | | | | |
| 2001 | 11.912 | -1.579 | -1.044 | -0.235 | -0.062 | -0.008 | 0.045 | 0.201 | 0.497 | 0.869 | 1.358 |
| 2005 | 11.179 | -1.376 | -0.816 | -0.165 | -0.053 | -0.005 | 0.053 | 0.228 | 0.563 | 0.954 | 1.395 |
| 2010 | 11.379 | -1.261 | -0.730 | -0.168 | -0.058 | -0.006 | 0.068 | 0.282 | 0.619 | 0.985 | 1.424 |
| IAB Data | | | | | | | | | | | |
| 2001 | 11.912 | -1.571 | -1.052 | -0.294 | -0.067 | -0.006 | 0.053 | 0.270 | 0.563 | 0.901 | 1.370 |
| 2005 | 11.179 | -1.365 | -0.813 | -0.204 | -0.057 | -0.004 | 0.059 | 0.280 | 0.593 | 0.946 | 1.396 |
| 2010 | 11.379 | -1.259 | -0.764 | -0.212 | -0.064 | -0.006 | 0.076 | 0.342 | 0.660 | 1.001 | 1.433 |
| TPP+ Data | | | | | | | | | | | |
| 2001 | 9.694 | -1.678 | -1.032 | -0.215 | -0.041 | 0.013 | 0.064 | 0.244 | 0.583 | 1.006 | 1.484 |
| 2005 | 8.413 | -1.455 | -0.801 | -0.135 | -0.031 | 0.012 | 0.069 | 0.270 | 0.657 | 1.088 | 1.556 |
| 2010 | 7.791 | -1.257 | -0.661 | -0.128 | -0.033 | 0.014 | 0.095 | 0.349 | 0.772 | 1.205 | 1.653 |

Notes: This table shows the number of observations (in millions) and selected percentiles of the combined IAB-TPP distribution of 5-year changes in residualized log earnings (from t to $t + 5$) for women and selected years. LS sample. Sources: IAB and TPP.

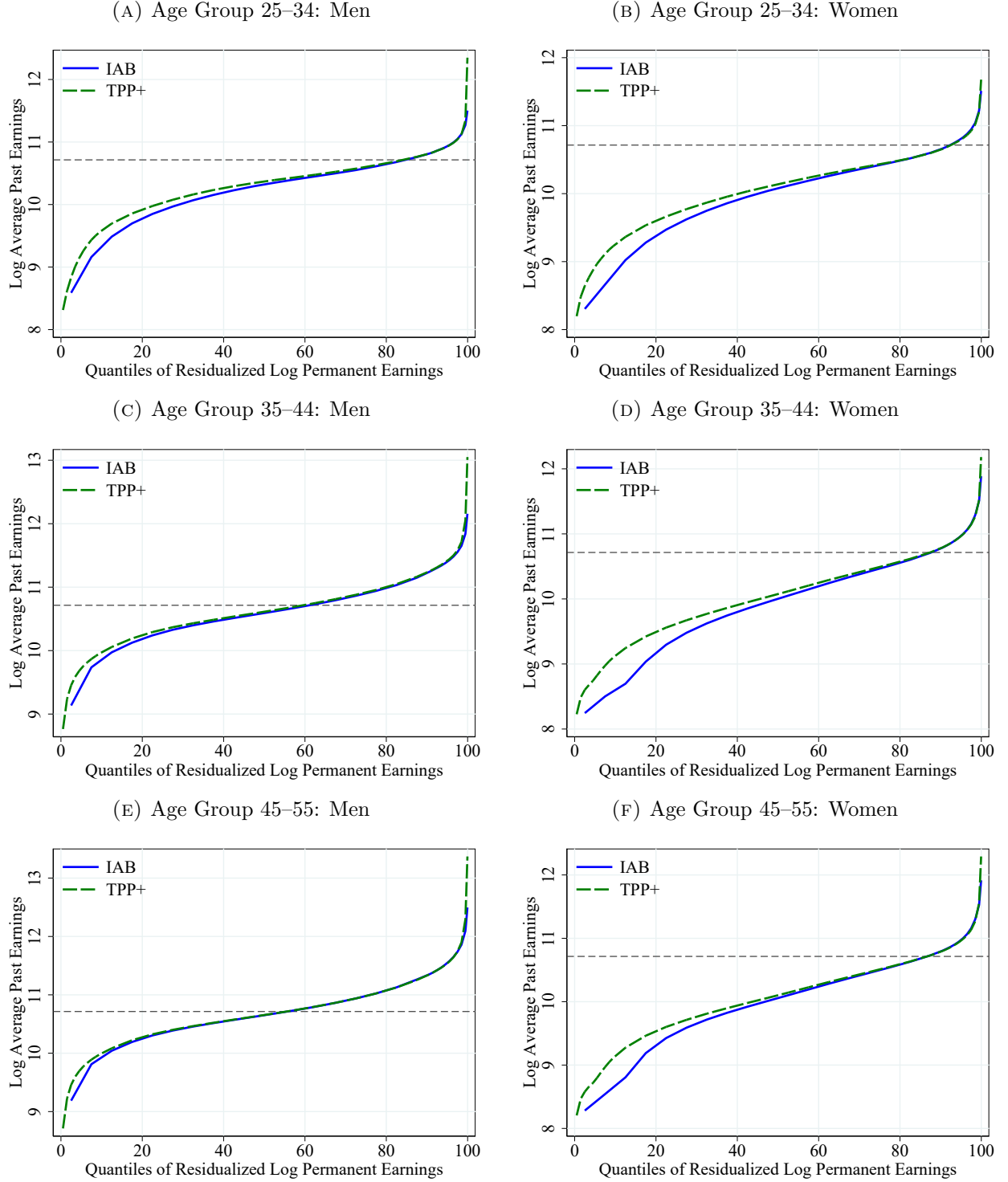
D.4 Earnings Growth by Permanent Earnings

For the heterogeneity analysis by permanent earnings, we use a simple cut-off rule to combine IAB and TPP data. For 1-year growth rates, this cutoff is equal to 45,000 Euro. Hence, for all quantiles of the residualized permanent earnings distribution above this cutoff, we use the conditional growth rate distribution computed from the IAB data. Above this cutoff, we use the corresponding condi-

tional statistics from the reweighted TPP data. There are two reasons for the choice of 45,000 Euro as the cutoff. First, Figure D.7 shows that both residualized permanent earnings and raw average past earnings converge in the middle of the distribution and are almost identical at the cutoff of 45,000 Euro. Second, we argue it is reasonable to assume that average past earnings below the cutoff are mostly unaffected by the top-coding threshold of 60,000 Euro such that the IAB data is reliable. Figures D.8, D.9 and D.10 show the P90-P10 differential, Kelley Skewness and Excess Crow-Siddiqui kurtosis by permanent earnings quantiles in the IAB and reweighted TPP data.

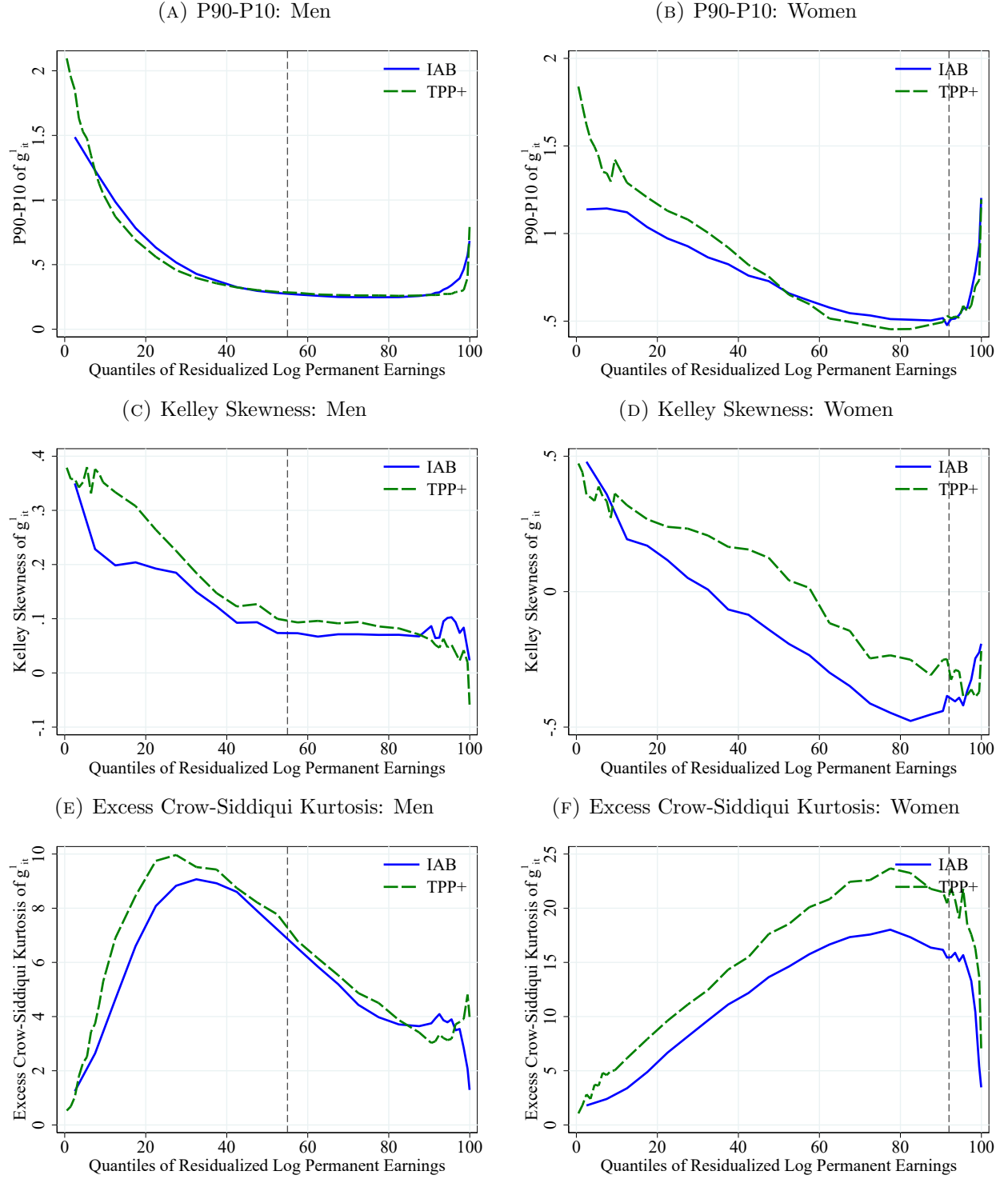
For 5-year earnings changes we proceed analogously, but use a cutoff of 40,000 Euro as jumping into the top-coded range is more likely over a period of five years. Figures D.11, D.12 and D.13 show the P90-P10 differential, Kelley Skewness and Excess Crow-Siddiqui kurtosis by permanent earnings quantiles in the IAB and reweighted TPP data.

FIGURE D.7: IAB vs. TPP: PERMANENT EARNINGS (H SAMPLE)



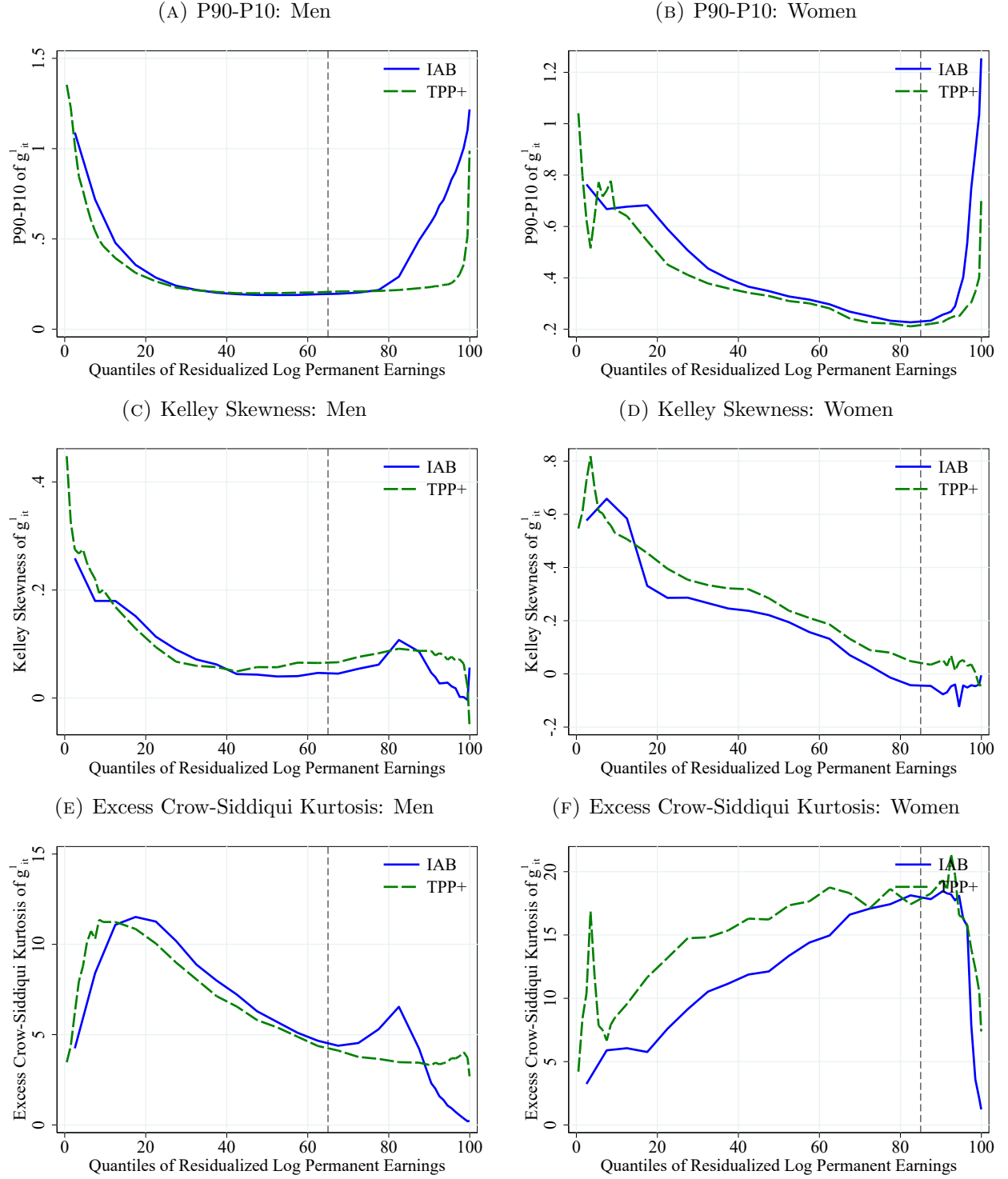
Notes: H sample, averages from 2004 to 2011. The dashed vertical depicts the point where permanent earnings is equal to 45,000 Euro, i.e. the point where the lines in Figure 9 in the main text switch from IAB to TPP data. Source: IAB and TPP.

FIGURE D.8: IAB vs. TPP: 1-YEAR LOG EARNINGS CHANGES BY PERMANENT EARNINGS, AGE GROUP 25–34



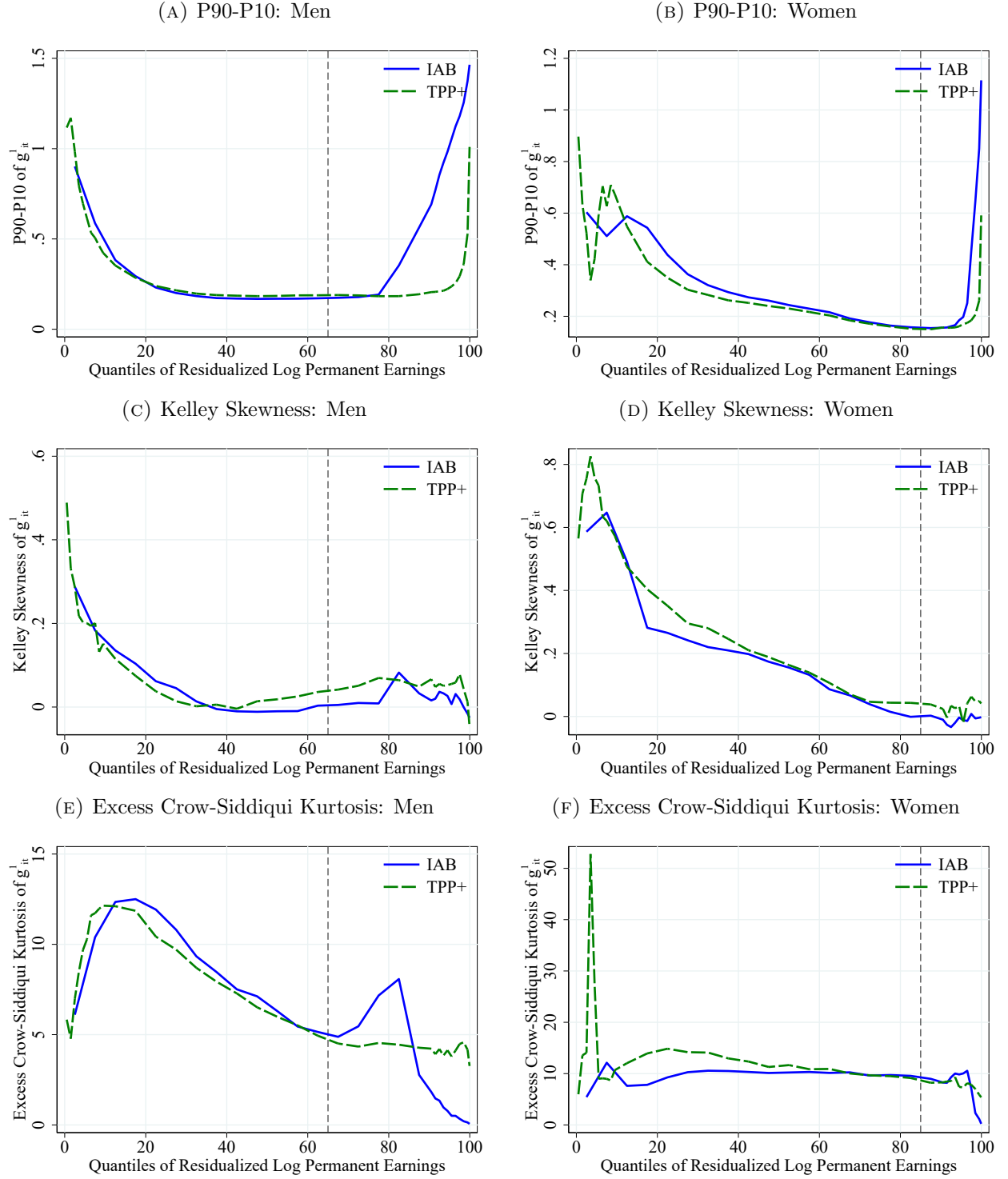
Notes: H sample, averages from 2004 to 2011. The dashed vertical depicts the point where permanent earnings is equal to 45,000 Euro, i.e. the point where the lines in Figure 9 in the main text switch from IAB to TPP data. Source: IAB and TPP.

FIGURE D.9: IAB vs. TPP: 1-YEAR LOG EARNINGS CHANGES BY PERMANENT EARNINGS, AGE GROUP 35–44



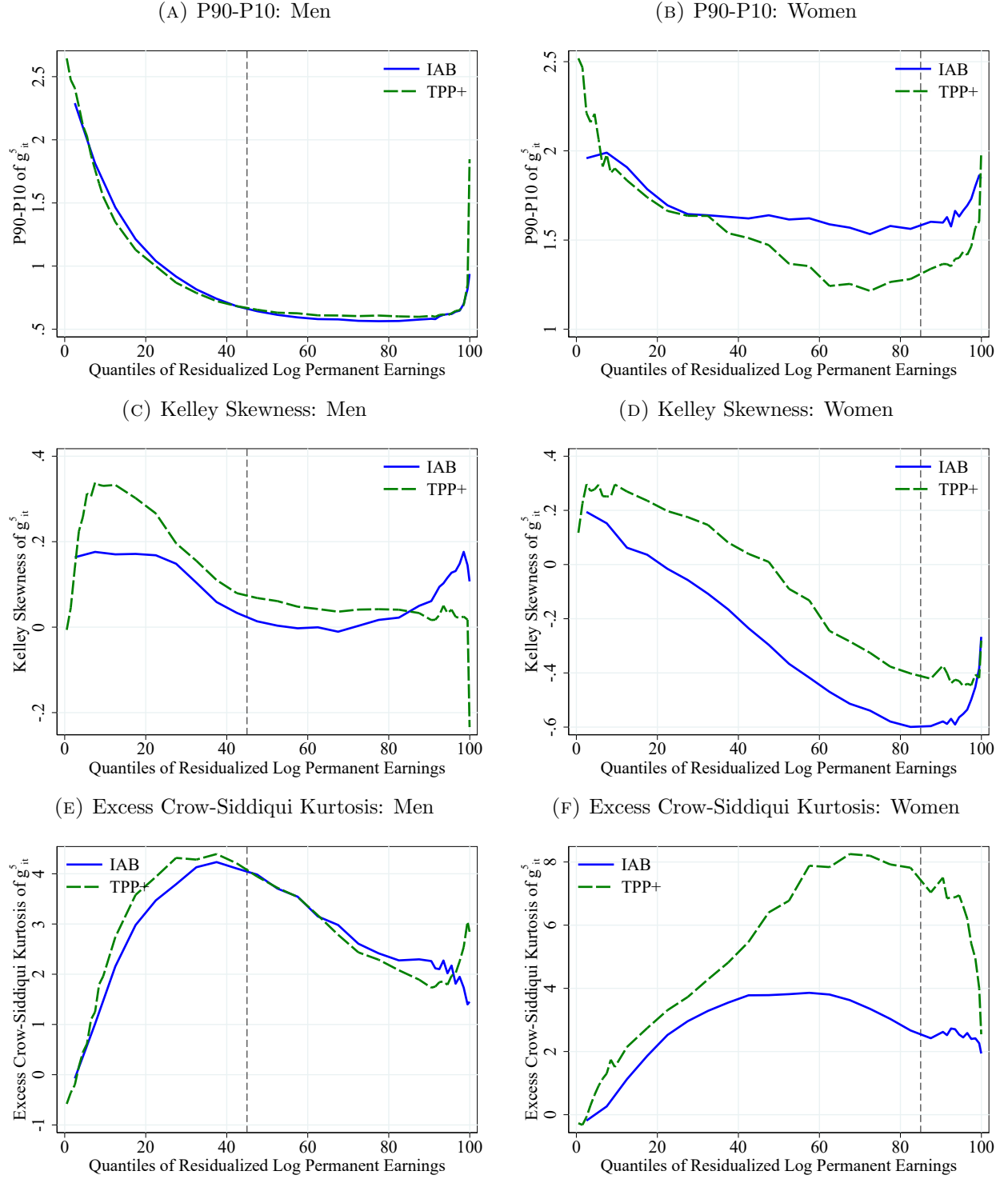
Notes: H sample, averages from 2004 to 2011. The dashed vertical depicts the point where permanent earnings is equal to 45,000 Euro, i.e. the point where the lines in Figure 9 in the main text switch from IAB to TPP data. Source: IAB and TPP.

FIGURE D.10: IAB vs. TPP: 1-YEAR LOG EARNINGS CHANGES BY PERMANENT EARNINGS, AGE GROUP 45–55



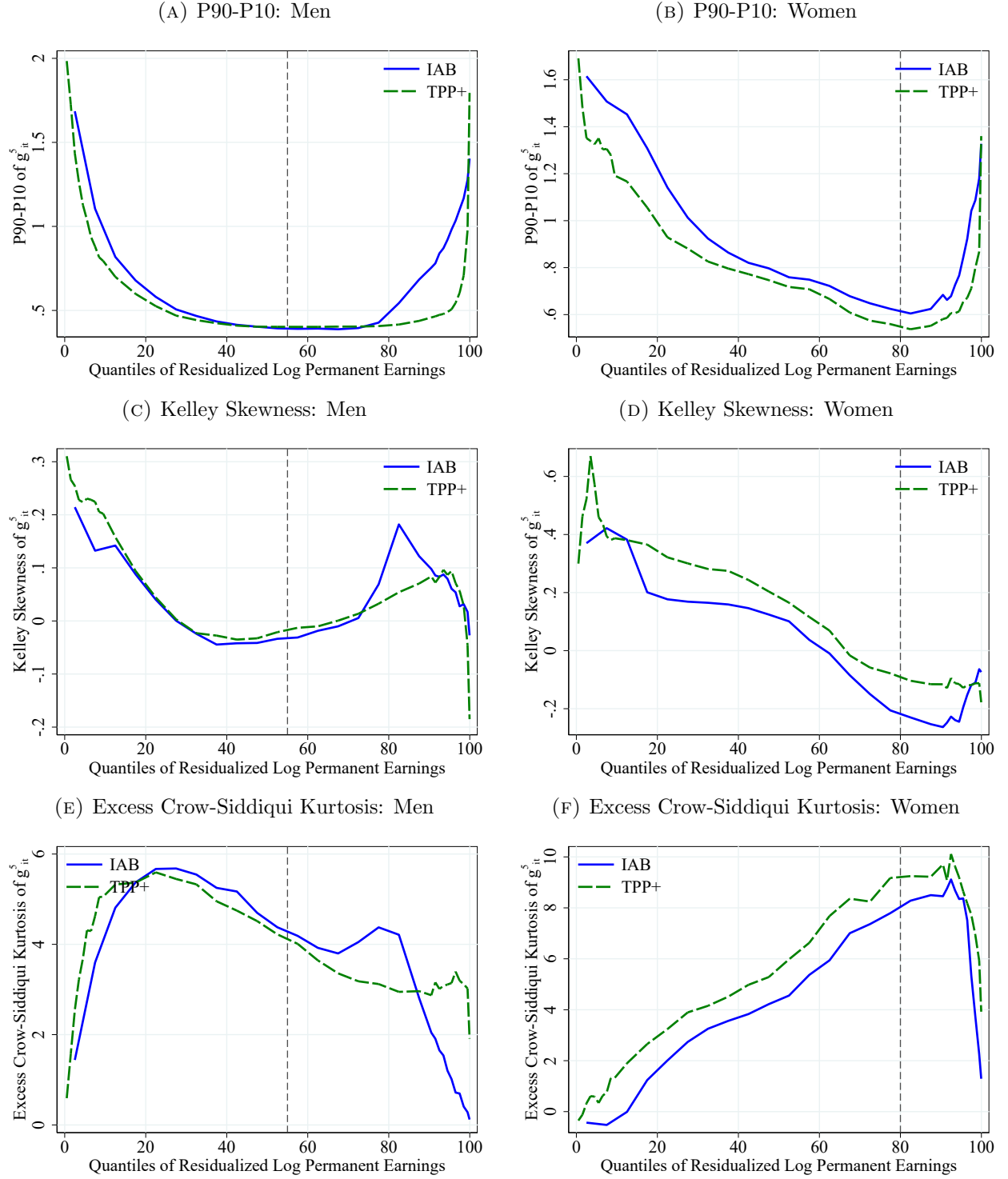
Notes: H sample, averages from 2004 to 2011. The dashed vertical depicts the point where permanent earnings is equal to 45,000 Euro, i.e. the point where the lines in Figure 9 in the main text switch from IAB to TPP data. Source: IAB and TPP.

FIGURE D.11: IAB vs. TPP: 5-YEAR LOG EARNINGS CHANGES BY PERMANENT EARNINGS, AGE GROUP 25–34



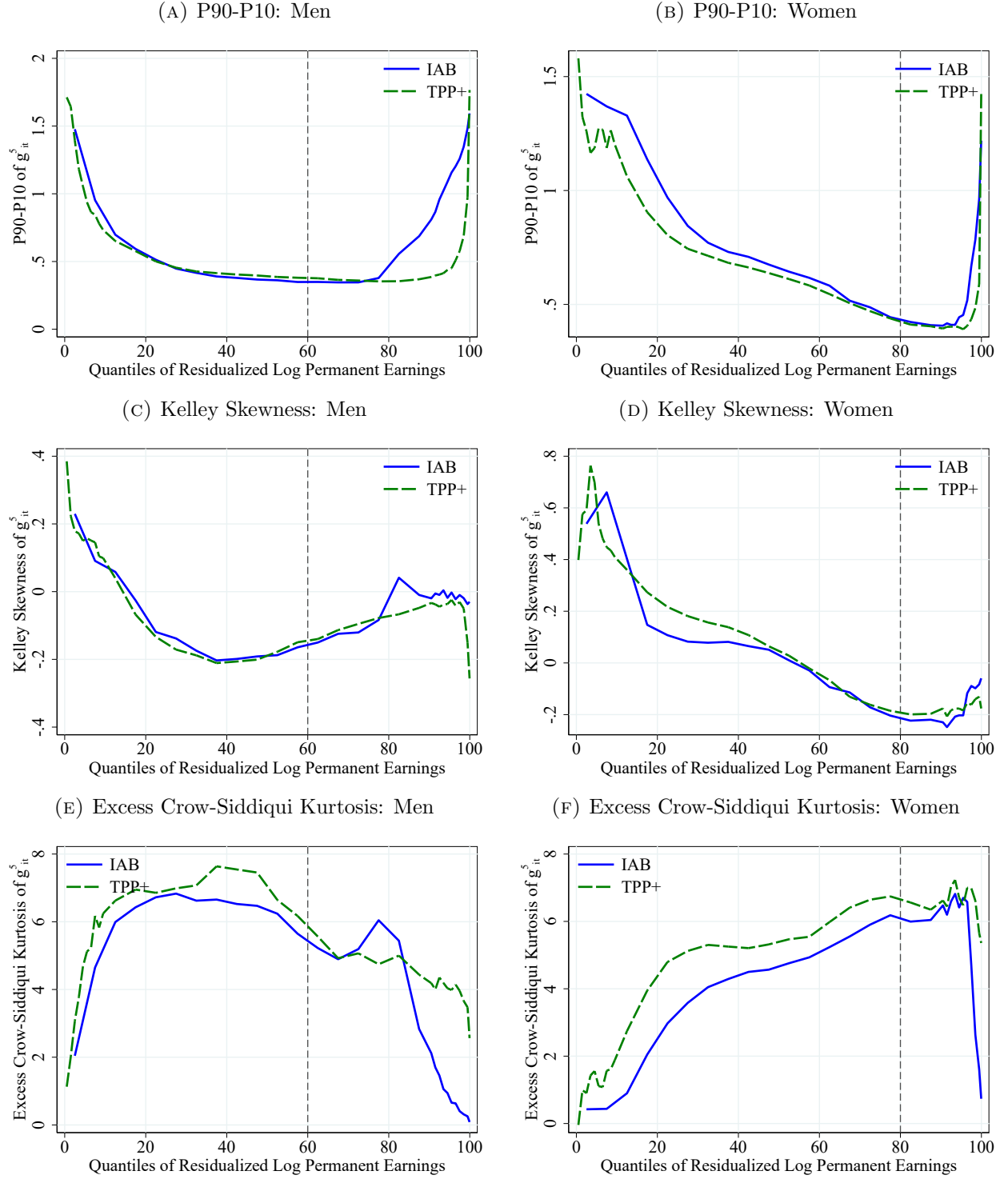
Notes: H sample, averages from 2004 to 2011. The dashed vertical depicts the point where permanent earnings is equal to 40,000 Euro, i.e. the point where the lines in Figure E.28 in the main text switch from IAB to TPP data. Source: IAB and TPP.

FIGURE D.12: IAB vs. TPP: 5-YEAR LOG EARNINGS CHANGES BY PERMANENT EARNINGS, AGE GROUP 35–44



Notes: H sample, averages from 2004 to 2011. The dashed vertical depicts the point where permanent earnings is equal to 40,000 Euro, i.e. the point where the lines in Figure E.28 in the main text switch from IAB to TPP data. Source: IAB and TPP.

FIGURE D.13: IAB vs. TPP: 5-YEAR LOG EARNINGS CHANGES BY PERMANENT EARNINGS, AGE GROUP 45–55



Notes: H sample, averages from 2004 to 2011. The dashed vertical depicts the point where permanent earnings is equal to 40,000 Euro, i.e. the point where the lines in Figure E.28 in the main text switch from IAB to TPP data. Source: IAB and TPP.

D.5 Combined IAB-TPP Data in Total Income Analysis (Section 4)

For the analysis of total income, we use the reweighted TPP data. Recall that the distribution of earnings in the subsample of social-security workers in the reweighted data matches the earnings distribution of the combined IAB-TPP data in the earnings analysis (see Appendix D.1). Note that the reweighting procedure does not distort the distribution of non-labor income as only workers who were not obliged to file a tax return are assigned a weight larger than one. The key point is that voluntary filers must not have annual non-labor income above 410 Euro.

The total income analysis sample additionally includes non-social-security workers (e.g. civil servants) and taxpayers who do not receive labor income (self-employed, business owners, landlords). Table D.11 shows how we arrive at the analysis sample starting from the unweighted TPP data (columns 1 and 4). Columns 2 and 5 show the reweighted TPP data before imposing the minimum income threshold of 2,300 Euro and columns 3 and 6 refer to the analysis sample used in Section 4 (see Table 2). In particular, Panel E shows that 1.1% of men and 1.8% of women have negative total income in 2008. While those observations are excluded from the analysis sample, there are still observations with above-threshold total income but negative non-labor income.

In Tables D.13 and D.14 we show pairwise correlations between the different income components. As expected, labor income is negatively correlated with business and self-employment income, and all income components are positively correlated with total income. The surprisingly low correlation of labor and total income is due to the presence of outliers, i.e. entrepreneurs (mostly business owners) who have no labor income but business and hence total income of more than 1 million Euro (up to 25 million Euro).

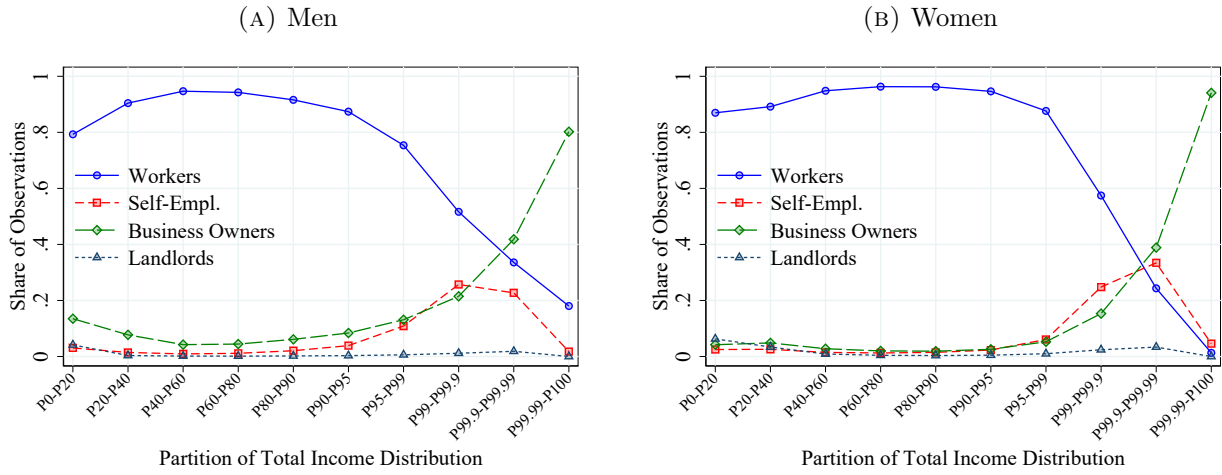
Table D.12 shows selected percentiles of the earnings distribution in the combined IAB-TPP data (CS sample). As mentioned above, percentiles below 60,000 Euro (P75 and below) are practically identical in the IAB-TPP and IAB data, while higher percentiles are closer to the TPP data.

TABLE D.11: SUMMARY STATISTICS FOR TOTAL INCOME DATA

| | Men | | | Women | | |
|--|-----------|-----------|-----------------------|-----------|---------|-----------------------|
| | TPP | IAB-TPP | IAB-TPP (analysis) | TPP | IAB-TPP | IAB-TPP (analysis) |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Observations (in mill.) | 11.584 | 15.006 | 14.756 | 8.986 | 12.836 | 12.479 |
| <i>A. Income Distribution</i> | | | | | | |
| Mean | 49,323 | 44,583 | 45,576 | 28,406 | 24,892 | 25,952 |
| P50 | 38,664 | 36,064 | 36,487 | 24,227 | 20,869 | 21,509 |
| P90 | 83,998 | 77,415 | 77,947 | 50,888 | 47,844 | 48,254 |
| P99.9 | 786,435 | 688,636 | 694,412 | 312,805 | 266,979 | 269,998 |
| P99.99 | 3,313,004 | 2,898,255 | 2,914,545 | 1,154,808 | 893,585 | 916,726 |
| <i>B. Share of Total Income</i> | | | | | | |
| Labor | 0.813 | 0.841 | 0.835 | 0.886 | 0.907 | 0.893 |
| Non-Labor | 0.187 | 0.159 | 0.165 | 0.114 | 0.093 | 0.107 |
| Self-Empl. | 0.063 | 0.054 | 0.054 | 0.052 | 0.043 | 0.043 |
| Business | 0.123 | 0.105 | 0.109 | 0.053 | 0.042 | 0.053 |
| Rental | 0.001 | 0.000 | 0.001 | 0.009 | 0.008 | 0.011 |
| Capital* | 0.016 | 0.014 | 0.013 | 0.006 | 0.006 | 0.005 |
| <i>C. Main Income Source</i> | | | | | | |
| Workers | 0.835 | 0.868 | 0.882 | 0.866 | 0.896 | 0.918 |
| Entrepreneurs | 0.165 | 0.132 | 0.118 | 0.134 | 0.104 | 0.082 |
| Self-Employed | 0.034 | 0.027 | 0.026 | 0.037 | 0.028 | 0.024 |
| Business Owners | 0.115 | 0.092 | 0.082 | 0.060 | 0.047 | 0.036 |
| Landlords | 0.015 | 0.013 | 0.010 | 0.037 | 0.030 | 0.022 |
| <i>D. Non-Zero Income</i> | | | | | | |
| Labor | 0.851 | 0.885 | 0.895 | 0.884 | 0.918 | 0.935 |
| Non-Labor | 0.364 | 0.312 | 0.300 | 0.272 | 0.230 | 0.208 |
| Self-Empl. | 0.064 | 0.053 | 0.052 | 0.065 | 0.053 | 0.049 |
| Business | 0.206 | 0.175 | 0.165 | 0.107 | 0.093 | 0.080 |
| Rental | 0.170 | 0.147 | 0.144 | 0.134 | 0.112 | 0.102 |
| Capital* | 0.123 | 0.103 | 0.103 | 0.055 | 0.044 | 0.042 |
| <i>E. Negative Income (if $\neq 0$)</i> | | | | | | |
| Total | 0.012 | 0.011 | 0.000 | 0.020 | 0.018 | 0.000 |
| Non-Labor | 0.101 | 0.098 | 0.087 | 0.074 | 0.071 | 0.051 |
| Self-Empl. | 0.007 | 0.006 | 0.005 | 0.009 | 0.009 | 0.006 |
| Business | 0.043 | 0.043 | 0.034 | 0.027 | 0.029 | 0.019 |
| Rental | 0.084 | 0.075 | 0.073 | 0.051 | 0.043 | 0.035 |
| Capital* | 0.007 | 0.006 | 0.006 | 0.003 | 0.003 | 0.002 |

Notes: This table shows descriptive statistics for the full TPP and IAB-TPP data by gender for the year 2008. The data includes all workers independent of their social-security status and individuals with non-labor income. Columns 1 and 4 refer to the raw TPP data (earnings not reweighted using IAB data). Columns 2 and 5 refer to the combined IAB-TPP data where observations with earnings are reweighted using IAB data (see Appendix D). Columns 3 and 6 refer to the analysis sample of the combined IAB-TPP where we require total income to be above the minimum income threshold of 2,300 Euro (2018 prices). Panel A shows the mean and selected percentiles of the total income distribution in 2018 Euro (excluding capital income). Panel B shows the share of each income source in total income (excluding capital income). Hence, the capital share is not part of the non-labor income share. Panel C reports the share of observations whose most important source of income is labor, non-labor (and sub-categories of non-labor income). Panel D shows the share of observations with non-zero income from different sources. Panel E shows the share of observations with negative income from different sources provided that the person has non-zero income from this source.

FIGURE D.14: MAIN INCOME SOURCES ACROSS THE INCOME DISTRIBUTION



Notes: This figure shows the share of observations classified as workers, self-employed, business owners and landlords in different parts of the total income distribution in the combined IAB-TPP data for the year 2008. A person is classified as a worker if her labor income is positive and (pairwise) larger than incomes from other sources. Source: IAB and TPP.

TABLE D.12: TOTAL INCOME PERCENTILES IN THE COMBINED IAB-TPP DATA

| Year | N | Mean | P5 | P10 | P25 | P50 | P75 | P90 | P95 | P99 | P99.9 | P99.99 |
|-------------------|--------|--------|-------|--------|--------|--------|--------|--------|---------|---------|---------|-----------|
| Men | | | | | | | | | | | | |
| 2001 | 15.373 | 43,989 | 6,654 | 11,937 | 25,329 | 37,808 | 51,579 | 72,826 | 93,353 | 176,187 | 517,722 | 1,984,431 |
| 2002 | 15.127 | 43,912 | 6,275 | 11,295 | 24,964 | 37,853 | 51,937 | 73,469 | 93,746 | 176,093 | 502,011 | 1,773,271 |
| 2003 | 14.866 | 43,903 | 5,889 | 10,754 | 24,574 | 37,924 | 52,403 | 74,273 | 94,649 | 176,508 | 491,658 | 1,725,813 |
| 2004 | 14.741 | 44,157 | 5,598 | 10,081 | 23,848 | 37,564 | 52,348 | 74,945 | 96,180 | 183,802 | 541,297 | 2,033,983 |
| 2005 | 14.565 | 44,453 | 5,523 | 9,854 | 23,259 | 37,169 | 52,323 | 75,668 | 97,818 | 192,677 | 598,642 | 2,369,003 |
| 2006 | 14.621 | 44,826 | 5,558 | 9,875 | 22,806 | 36,824 | 52,567 | 76,633 | 99,720 | 200,946 | 635,102 | 2,485,810 |
| 2007 | 14.758 | 45,253 | 5,674 | 10,132 | 22,757 | 36,540 | 52,468 | 77,266 | 101,524 | 208,477 | 681,134 | 2,819,799 |
| 2008 | 14.768 | 45,542 | 5,650 | 10,272 | 22,678 | 36,465 | 52,652 | 77,918 | 102,878 | 214,685 | 694,181 | 2,914,545 |
| 2009 | 14.498 | 44,690 | 5,527 | 9,817 | 22,324 | 36,184 | 52,133 | 77,544 | 102,111 | 208,188 | 637,002 | 2,280,813 |
| 2010 | 14.630 | 45,107 | 5,565 | 9,988 | 22,084 | 36,203 | 52,765 | 78,271 | 103,285 | 211,273 | 655,038 | 2,521,691 |
| 2011 | 14.796 | 45,650 | 5,697 | 10,358 | 22,360 | 36,182 | 53,091 | 79,270 | 105,317 | 217,863 | 683,204 | 2,540,327 |
| 2012 | 14.854 | 45,730 | 5,574 | 10,257 | 22,388 | 36,202 | 53,345 | 79,771 | 105,600 | 217,914 | 674,841 | 2,568,046 |
| 2013 | 14.892 | 45,729 | 5,550 | 10,092 | 22,332 | 36,265 | 53,393 | 79,874 | 105,720 | 218,549 | 682,847 | 2,719,594 |
| 2014 | 14.974 | 46,199 | 5,523 | 9,932 | 22,371 | 36,501 | 53,977 | 80,933 | 107,136 | 223,002 | 704,432 | 2,746,560 |
| 2015 | 15.054 | 47,085 | 5,663 | 10,237 | 22,632 | 36,829 | 54,736 | 82,248 | 109,164 | 228,277 | 736,971 | 2,885,246 |
| 2016 | 15.079 | 47,768 | 5,783 | 10,540 | 23,111 | 37,216 | 55,323 | 83,213 | 110,686 | 233,049 | 758,905 | 2,944,148 |
| Women | | | | | | | | | | | | |
| 2001 | 12.558 | 26,126 | 3,704 | 4,662 | 11,602 | 22,908 | 35,732 | 47,757 | 56,790 | 89,124 | 218,577 | 690,697 |
| 2002 | 12.531 | 26,274 | 3,685 | 4,624 | 11,573 | 22,980 | 35,949 | 48,267 | 57,596 | 89,996 | 216,063 | 641,789 |
| 2003 | 12.363 | 26,280 | 3,652 | 4,760 | 11,347 | 22,910 | 36,120 | 48,631 | 57,970 | 90,759 | 216,306 | 630,920 |
| 2004 | 12.345 | 26,001 | 3,628 | 4,761 | 10,571 | 22,396 | 35,786 | 48,433 | 58,234 | 92,789 | 227,721 | 741,022 |
| 2005 | 12.294 | 25,957 | 3,600 | 4,729 | 10,304 | 22,081 | 35,533 | 48,386 | 58,325 | 94,709 | 239,373 | 823,981 |
| 2006 | 12.330 | 25,855 | 3,597 | 4,749 | 10,125 | 21,756 | 35,200 | 48,276 | 58,432 | 96,490 | 250,291 | 881,431 |
| 2007 | 12.486 | 25,799 | 3,691 | 4,813 | 10,163 | 21,492 | 34,812 | 48,094 | 58,703 | 99,261 | 262,099 | 924,239 |
| 2008 | 12.522 | 25,876 | 3,723 | 4,841 | 10,322 | 21,429 | 34,832 | 48,205 | 58,929 | 100,686 | 269,619 | 914,497 |
| 2009 | 12.544 | 26,000 | 3,734 | 4,870 | 10,386 | 21,573 | 35,217 | 48,808 | 59,419 | 100,528 | 261,006 | 841,224 |
| 2010 | 12.644 | 26,232 | 3,775 | 4,918 | 10,580 | 21,601 | 35,279 | 49,274 | 60,077 | 101,969 | 269,650 | 957,126 |
| 2011 | 12.774 | 26,355 | 3,814 | 4,951 | 10,824 | 21,629 | 35,224 | 49,202 | 60,336 | 103,478 | 275,482 | 976,390 |
| 2012 | 12.905 | 26,440 | 3,810 | 4,979 | 11,020 | 21,688 | 35,242 | 49,378 | 60,753 | 104,167 | 278,631 | 914,946 |
| 2013 | 12.962 | 26,723 | 3,891 | 5,115 | 11,253 | 21,933 | 35,518 | 49,673 | 61,141 | 105,084 | 281,024 | 963,589 |
| 2014 | 13.028 | 27,315 | 3,964 | 5,176 | 11,525 | 22,313 | 36,135 | 50,550 | 62,439 | 108,432 | 292,271 | 1,032,300 |
| 2015 | 13.092 | 27,996 | 4,093 | 5,467 | 12,158 | 22,811 | 36,777 | 51,477 | 63,736 | 110,838 | 303,084 | 1,050,728 |
| 2016 | 13.079 | 28,707 | 4,179 | 5,604 | 12,693 | 23,477 | 37,553 | 52,357 | 64,982 | 113,904 | 314,598 | 1,167,646 |
| Population | | | | | | | | | | | | |
| 2001 | 27.930 | 35,958 | 4,353 | 6,679 | 17,265 | 31,617 | 44,775 | 62,189 | 79,289 | 144,373 | 409,211 | 1,504,645 |
| 2002 | 27.658 | 35,921 | 4,295 | 6,483 | 17,004 | 31,554 | 45,005 | 62,730 | 79,860 | 144,218 | 397,504 | 1,374,766 |
| 2003 | 27.230 | 35,902 | 4,279 | 6,114 | 16,735 | 31,500 | 45,295 | 63,238 | 80,655 | 144,970 | 391,890 | 1,327,219 |
| 2004 | 27.086 | 35,882 | 4,214 | 5,838 | 16,057 | 30,976 | 45,033 | 63,502 | 81,621 | 149,589 | 424,043 | 1,530,896 |
| 2005 | 26.859 | 35,987 | 4,187 | 5,775 | 15,713 | 30,490 | 44,868 | 63,751 | 82,548 | 155,180 | 464,490 | 1,771,740 |
| 2006 | 26.951 | 36,147 | 4,206 | 5,763 | 15,491 | 30,060 | 44,785 | 64,255 | 83,873 | 160,990 | 492,269 | 1,935,221 |
| 2007 | 27.244 | 36,338 | 4,280 | 5,799 | 15,450 | 29,727 | 44,556 | 64,505 | 84,916 | 166,350 | 520,798 | 2,096,380 |
| 2008 | 27.291 | 36,518 | 4,326 | 5,849 | 15,483 | 29,599 | 44,579 | 64,881 | 85,714 | 170,816 | 534,067 | 2,142,705 |
| 2009 | 27.042 | 36,020 | 4,299 | 5,806 | 15,274 | 29,446 | 44,497 | 64,623 | 85,083 | 166,595 | 496,910 | 1,745,501 |
| 2010 | 27.274 | 36,357 | 4,335 | 5,901 | 15,325 | 29,369 | 44,840 | 65,310 | 86,026 | 169,209 | 510,570 | 1,892,326 |
| 2011 | 27.570 | 36,710 | 4,360 | 6,085 | 15,636 | 29,353 | 44,896 | 65,835 | 87,229 | 173,786 | 530,741 | 1,966,156 |
| 2012 | 27.760 | 36,762 | 4,346 | 6,154 | 15,641 | 29,342 | 44,938 | 66,273 | 87,759 | 173,904 | 522,565 | 1,887,114 |
| 2013 | 27.854 | 36,884 | 4,421 | 6,218 | 15,735 | 29,419 | 45,109 | 66,442 | 87,958 | 174,208 | 527,232 | 1,967,935 |
| 2014 | 28.002 | 37,413 | 4,450 | 6,287 | 15,897 | 29,693 | 45,663 | 67,384 | 89,260 | 177,867 | 542,424 | 2,047,636 |
| 2015 | 28.146 | 38,205 | 4,602 | 6,718 | 16,476 | 30,070 | 46,318 | 68,547 | 90,879 | 182,272 | 564,562 | 2,194,591 |
| 2016 | 28.158 | 38,914 | 4,719 | 7,120 | 17,013 | 30,659 | 46,947 | 69,538 | 92,185 | 185,773 | 582,065 | 2,291,329 |

Notes: This table shows selected total income percentiles for men, women and in the population in the combined IAB-TPP. Capital income is not included in total or non-labor income. Note that total incomes in the analysis sample must exceed the minimum income threshold of 2,300 Euro (in 2018 prices). CS sample. Source: IAB and TPP.

TABLE D.13: CORRELATIONS BETWEEN INCOME COMPONENTS – MEN

| | Total | Labor | Non-Labor | Business | Self-Empl. | Capital | Rental |
|------------|--------|---------|-----------|----------|------------|---------|---------|
| Total | 1.0000 | 0.3158 | 0.9432 | 0.9237 | 0.1871 | 0.0297 | 0.0525 |
| Labor | 0.3158 | 1.0000 | -0.0172 | -0.0060 | -0.0445 | 0.0237 | -0.0300 |
| Non-Labor | 0.9432 | -0.0172 | 1.0000 | 0.9755 | 0.2128 | 0.0230 | 0.0658 |
| Business | 0.9237 | -0.0060 | 0.9755 | 1.0000 | 0.0017 | 0.0215 | 0.0105 |
| Self-Empl. | 0.1871 | -0.0445 | 0.2128 | 0.0017 | 1.0000 | 0.0070 | -0.0318 |
| Capital | 0.0297 | 0.0237 | 0.0230 | 0.0215 | 0.0070 | 1.0000 | 0.0103 |
| Rental | 0.0525 | -0.0300 | 0.0658 | 0.0105 | -0.0318 | 0.0103 | 1.0000 |

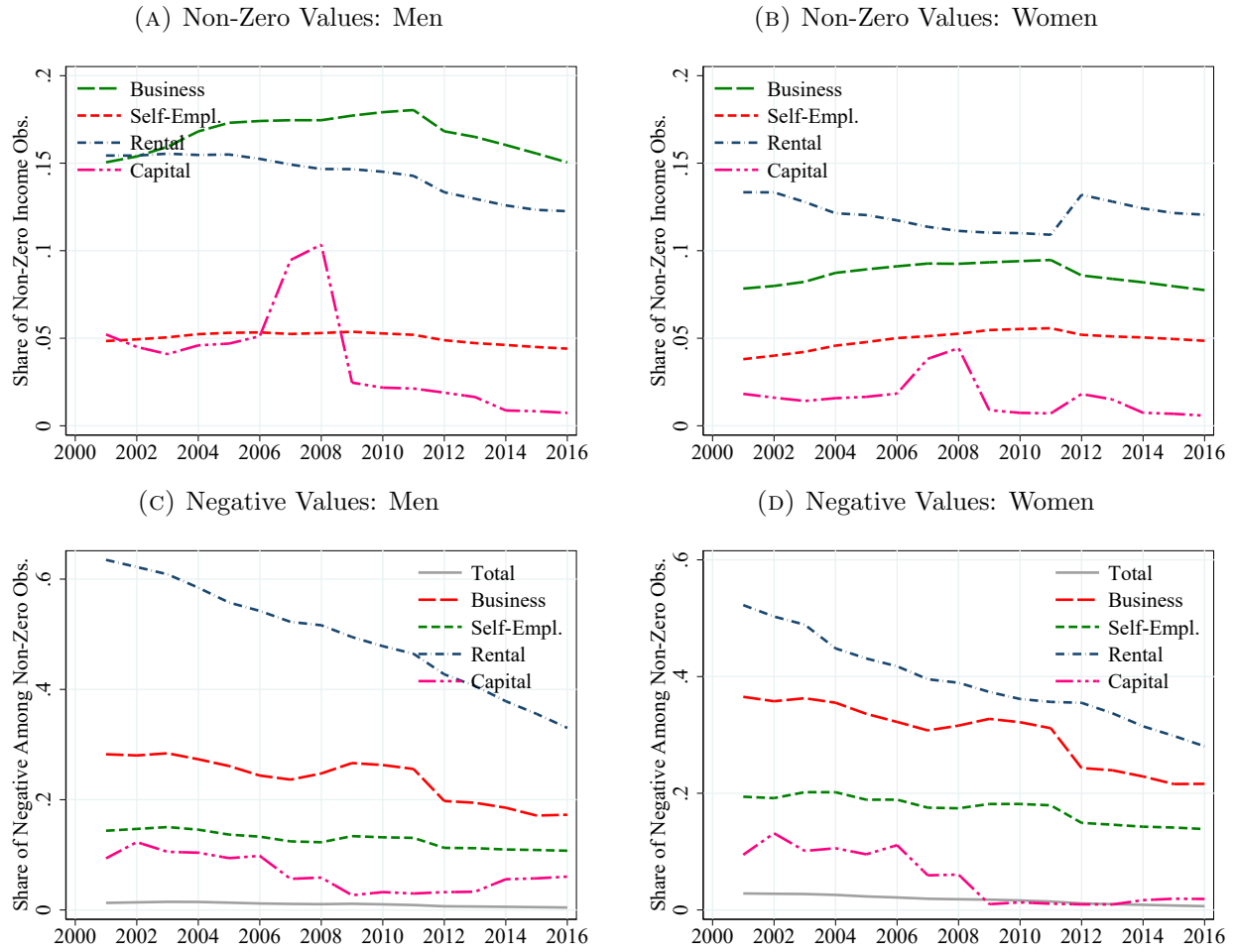
Notes: This table shows correlations between different income components in the combined IAB-TPP analysis sample for men in 2008. Capital income is not included in total or non-labor income. Note that total incomes in the analysis sample must exceed the minimum income threshold of 2,300 Euro (in 2018 prices). CS sample. Source: IAB and TPP.

TABLE D.14: CORRELATIONS BETWEEN INCOME COMPONENTS – WOMEN

| | Total | Labor | Non-Labor | Business | Self-Empl. | Capital | Rental |
|------------|--------|---------|-----------|----------|------------|---------|---------|
| Total | 1.0000 | 0.3054 | 0.9394 | 0.9088 | 0.1747 | 0.0801 | 0.1554 |
| Labor | 0.3054 | 1.0000 | -0.0395 | -0.0174 | -0.0875 | 0.0403 | -0.0220 |
| Non-Labor | 0.9394 | -0.0395 | 1.0000 | 0.9599 | 0.2148 | 0.0695 | 0.1710 |
| Business | 0.9088 | -0.0174 | 0.9599 | 1.0000 | 0.0004 | 0.0592 | -0.0084 |
| Self-Empl. | 0.1747 | -0.0875 | 0.2148 | 0.0004 | 1.0000 | 0.0195 | -0.0080 |
| Capital | 0.0801 | 0.0403 | 0.0695 | 0.0592 | 0.0195 | 1.0000 | 0.0467 |
| Rental | 0.1554 | -0.0220 | 0.1710 | -0.0084 | -0.0080 | 0.0467 | 1.0000 |

Notes: This table shows correlations between different income components in the combined IAB-TPP analysis sample for women in 2008. CS sample. Source: IAB and TPP.

FIGURE D.15: NON-ZERO AND NEGATIVE VALUES FOR NON-LABOR INCOME



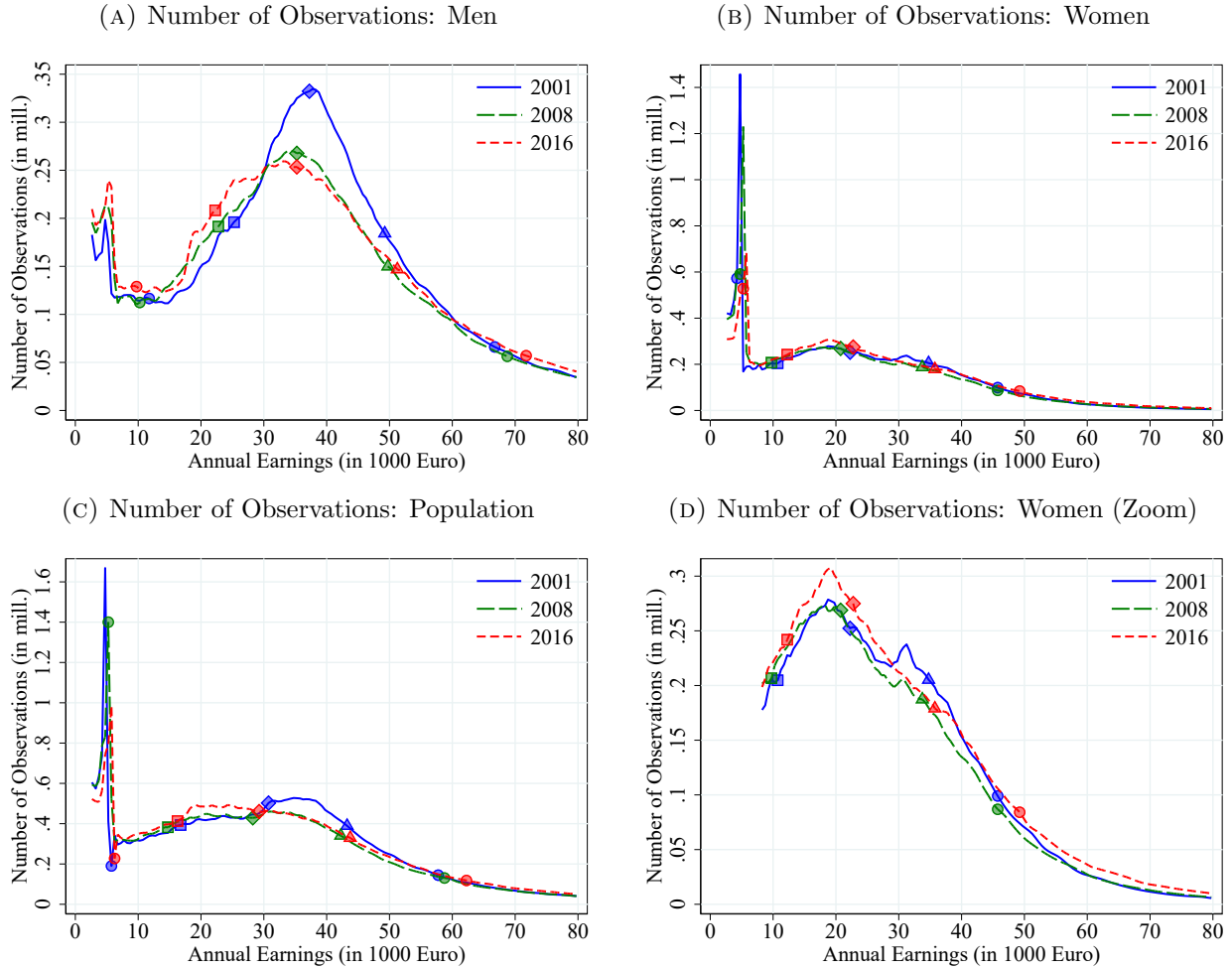
Notes: Panels A and B show the share of total income for different non-labor income components. Panels C and D show the share of observations with non-zero income from these components. Panels E and F show the share of observations out of all non-zero observations with negative income. Total income includes capital income. Source: TPP re-weighted using IAB data.

E Core Analysis: Additional Results for Combined IAB-TPP Data 2001–2016

In this Appendix we present additional results for the core analysis in Section 3.

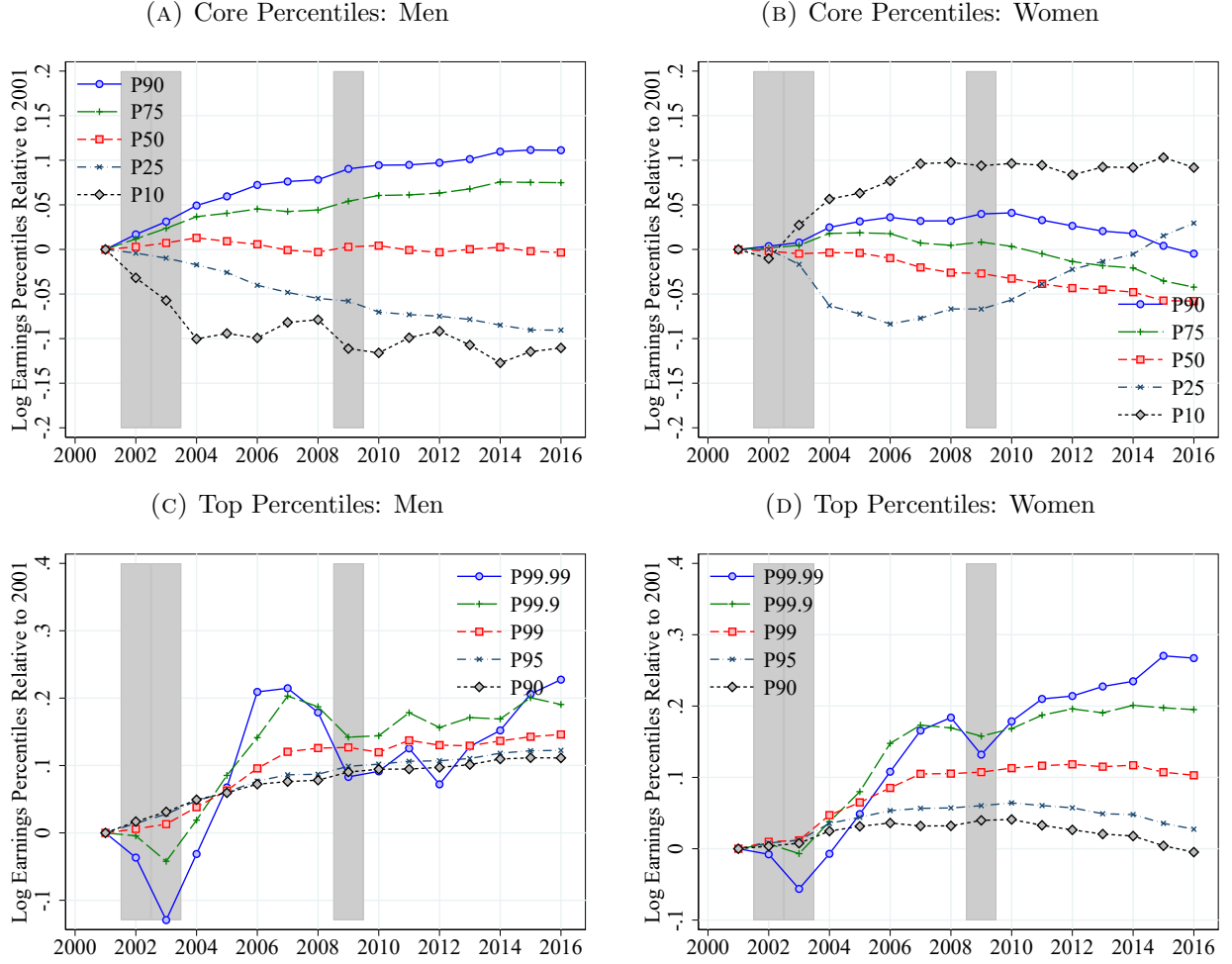
E.1 Additional Results for Earnings Inequality (Section 3.1)

FIGURE E.1: EARNINGS DISTRIBUTION



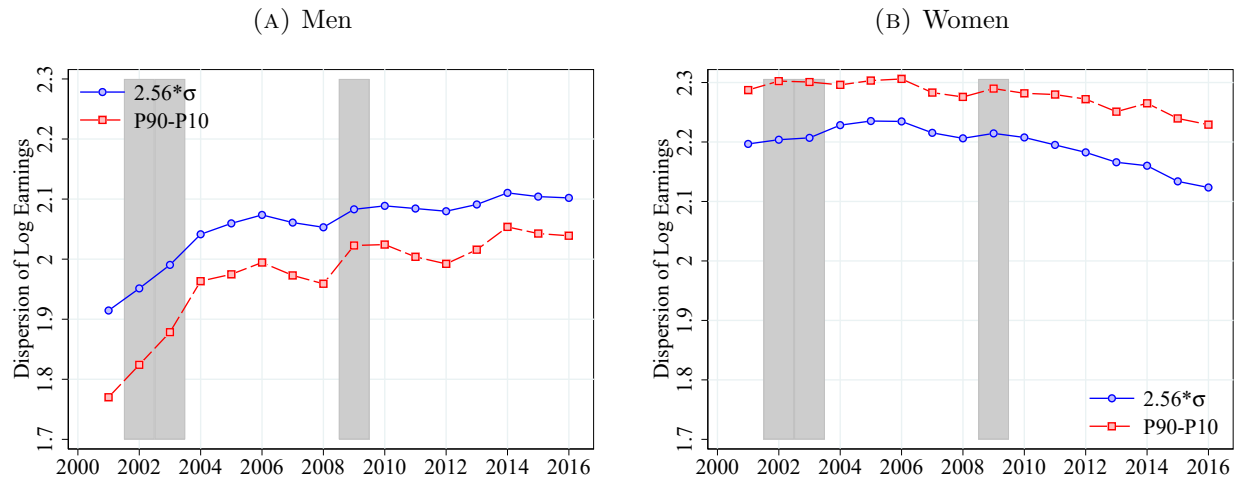
Notes: This figure shows the number of observations per 1,000 Euro earnings bins of real annual earnings for selected years in the combined IAB-TPP data (CS sample) separately for men and women. Panel A and B are depicted as shares in Figure 2. The data is smoothed (by year and gender) using a three-bin moving average for bins above 10,000 Euro. The markers indicate the 10th (circle), 25th (square), 50th (i.e. median; diamond), 75th (triangle) and 90th (circle again) percentiles of the respective distributions.

FIGURE E.2: EVOLUTION OF RESIDUAL LOG EARNINGS PERCENTILES (CONTROLLING FOR AGE)



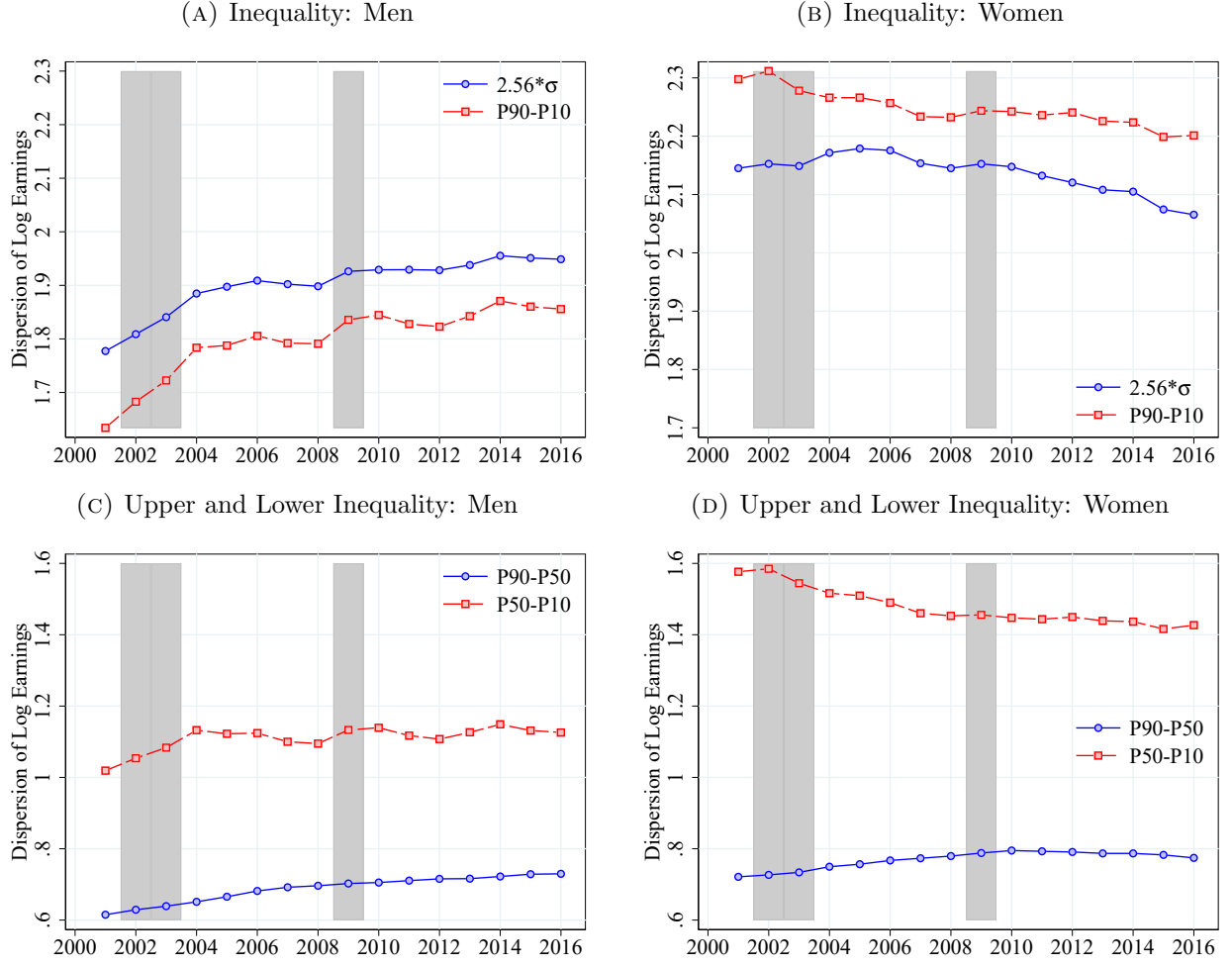
Notes: This figure shows the evolution of residualized log real annual earnings (controlling for age, for unconditioned percentiles, see Figure 3) in the combined IAB-TPP data (CS sample). Shaded areas indicate recessions.

FIGURE E.3: EVOLUTION OF EARNINGS INEQUALITY: STANDARD DEVIATION AND LOG PERCENTILE DIFFERENTIALS



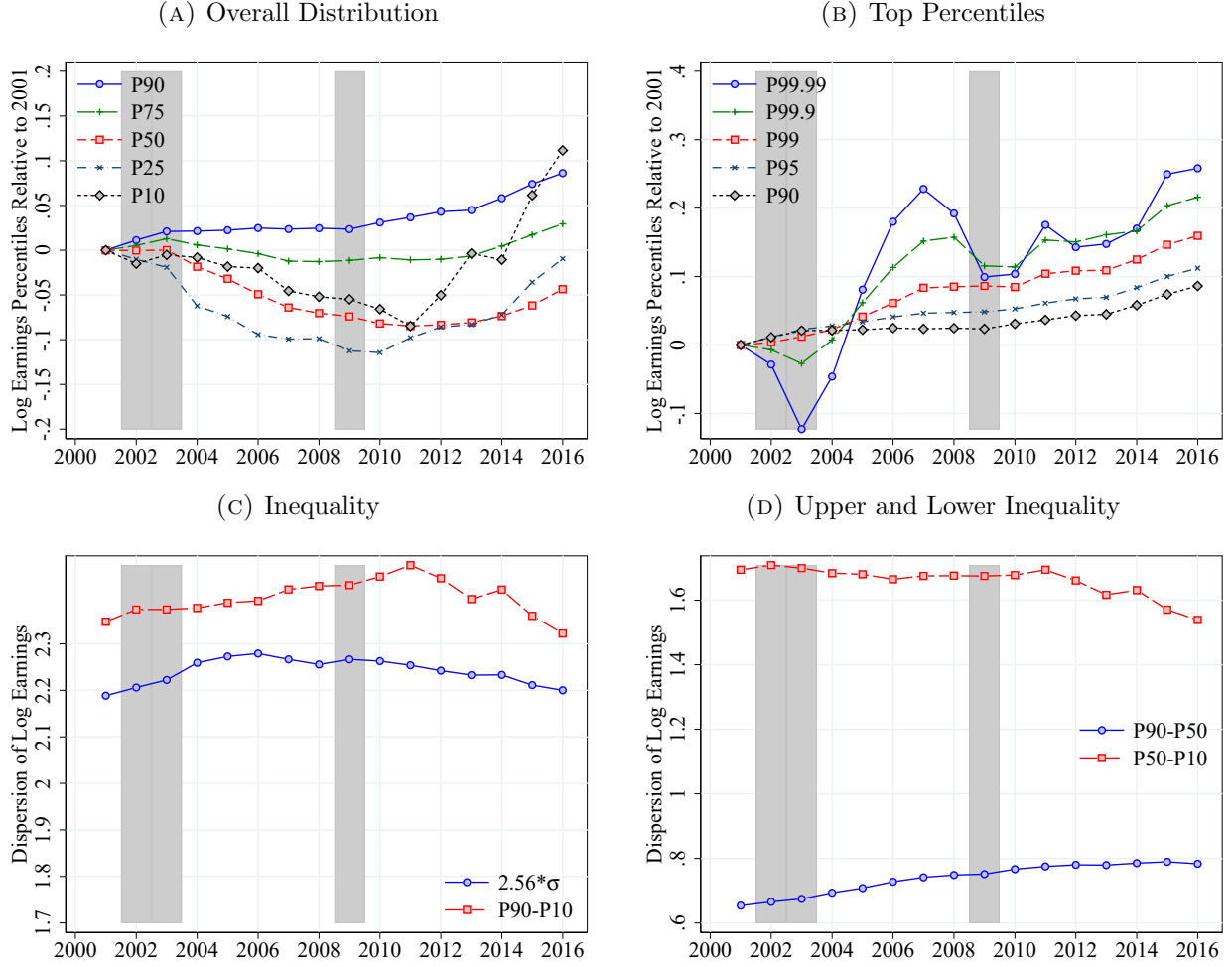
Notes: This figure shows the evolution of different log percentile differentials as well as the (rescaled) standard deviation of the log real annual earnings distribution over time in the combined IAB-TPP data (CS sample) separately for men and women. The standard deviation σ is rescaled as $2.56 * \sigma$ corresponds to P90-P10 differential for a Gaussian distribution. Shaded areas indicate recessions.

FIGURE E.4: RESIDUAL EARNINGS INEQUALITY (CONTROLLING FOR AGE)



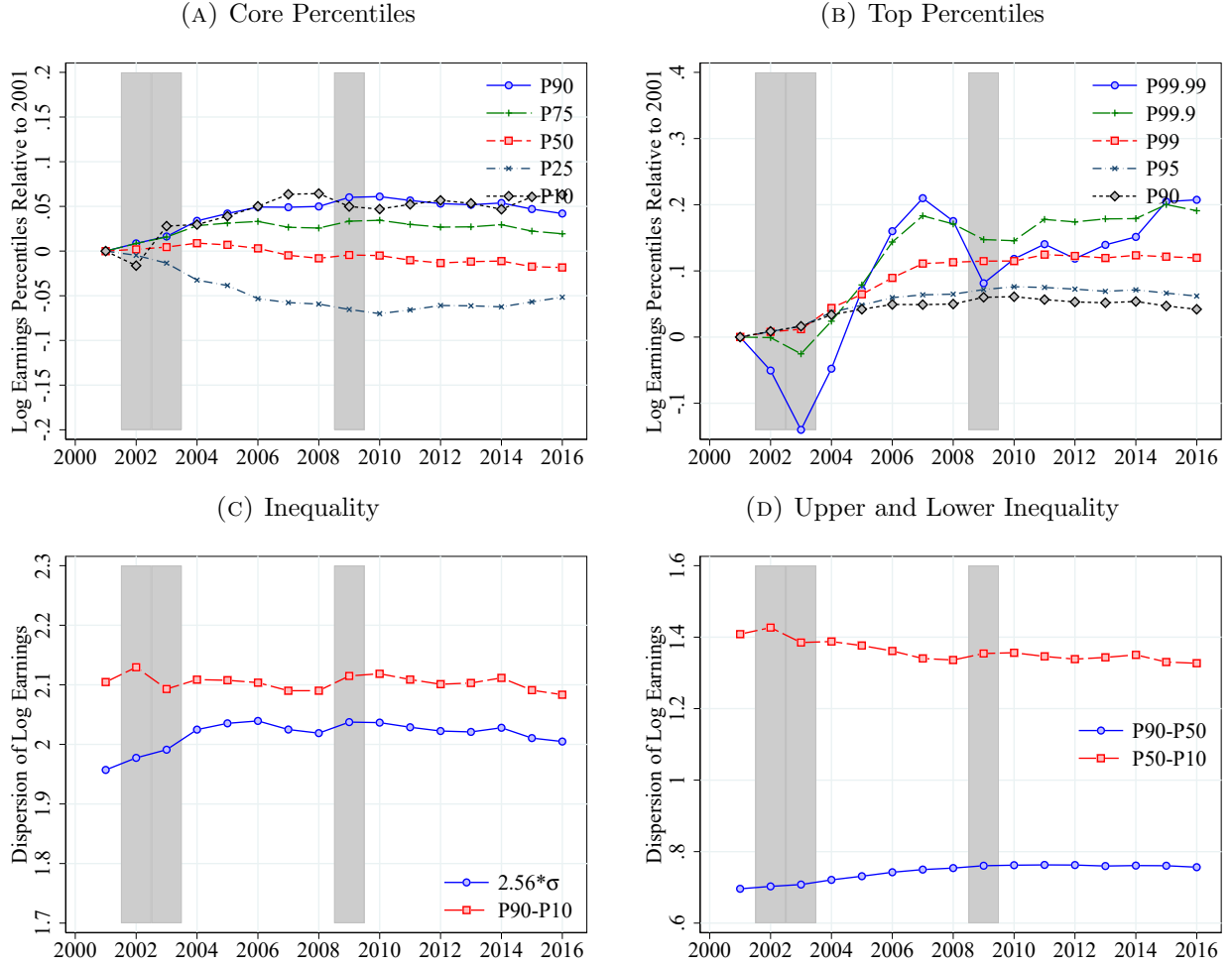
Notes: This figure shows the evolution of residualized log real annual earnings (controlling for age, unconditioned results can be found in Figure 4) in the combined IAB-TPP data (CS sample). Shaded areas indicate recessions.

FIGURE E.5: EVOLUTION OF LOG EARNINGS PERCENTILES IN THE POPULATION



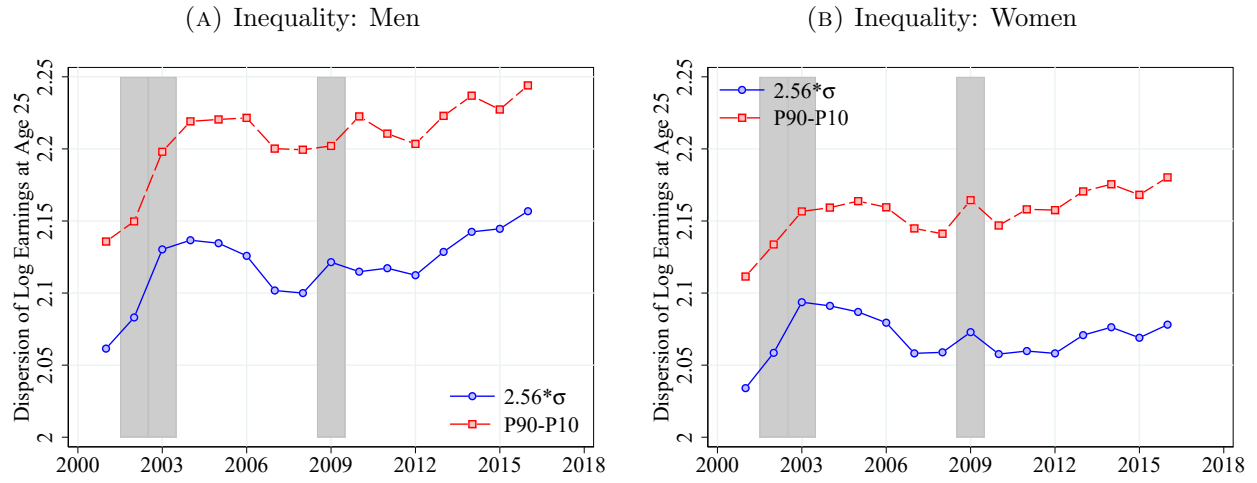
Notes: This figure shows the evolution of selected percentiles of log real annual earnings (relative to 2001) in the combined IAB-TPP data (CS sample) in the join data of men and women. Shaded areas indicate recessions.

FIGURE E.6: RESIDUAL LOG EARNINGS INEQUALITY IN THE POPULATION (CONTROLLING FOR GENDER AND AGE)



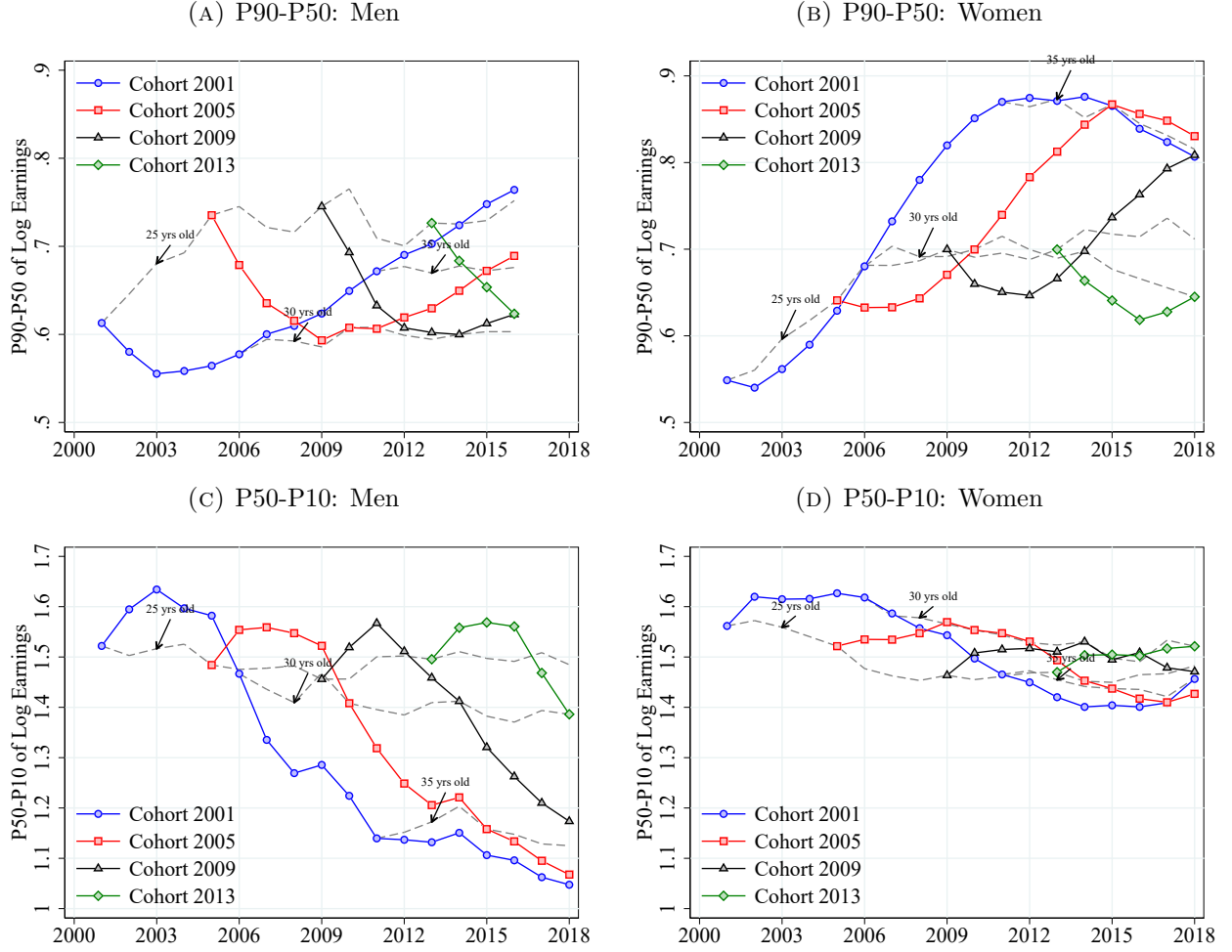
Notes: This figure shows the evolution of residualized log real annual earnings (controlling for gender and age, unconditioned results can be found in Figures 3 and 4.) in the combined IAB-TPP data (CS sample). Shaded areas indicate recessions.

FIGURE E.7: INITIAL INCOME INEQUALITY (AT AGE 25)



Notes: This figure shows the evolution of the P90-P10 log percentile differential as well as the (rescaled) standard deviation of the log real annual earnings distribution over time in the IAB data (CS sample) separately for men and women at the age of 25 in each year. The standard deviation σ is rescaled as $2.56 * \sigma$ corresponds to P90-P10 differential for a Gaussian distribution. Shaded areas indicate recessions.

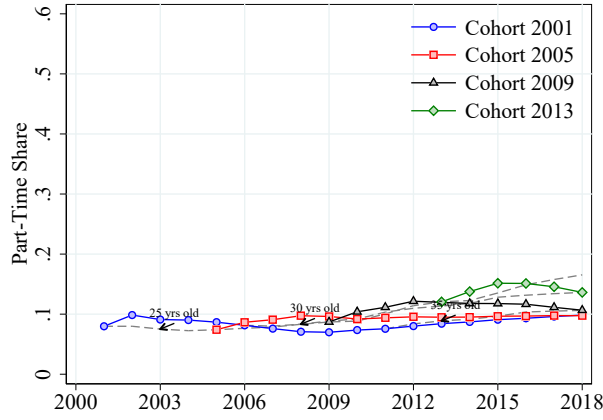
FIGURE E.8: UPPER AND LOWER EARNINGS INEQUALITY BY COHORT



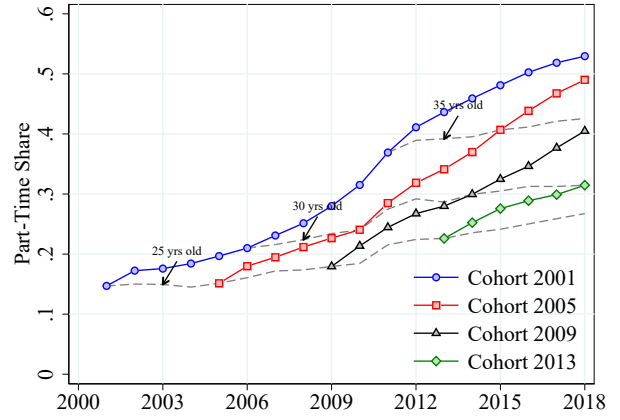
Notes: This figure shows the evolution of the P90-P50 and the P50-P10 differentials of the log real annual earnings distribution over time in the combined IAB-TPP data (CS sample) separately for men and women. As the P90 of men is imputed and the TPP data end in 2016, Panel A also ends in 2016. Grey dashed lines correspond to earnings inequality of 25, 30 and 35 year olds in each year as indicated by arrows. Each colored line corresponds to an individual cohort, where “cohort t ” represents the cohort aged 25 in year t .

FIGURE E.9: EMPLOYMENT LEVELS AND EDUCATION OVER THE LIFECYCLE

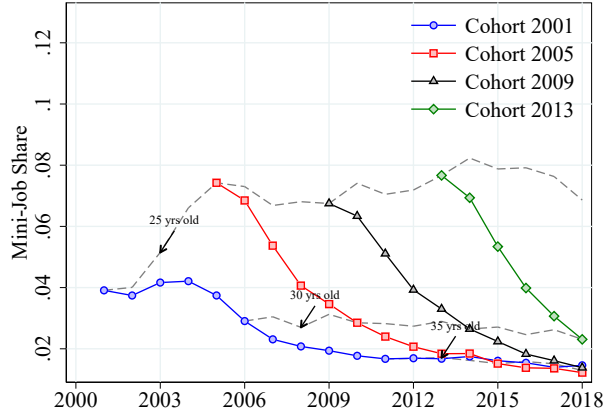
(A) Part-Time Share: Men



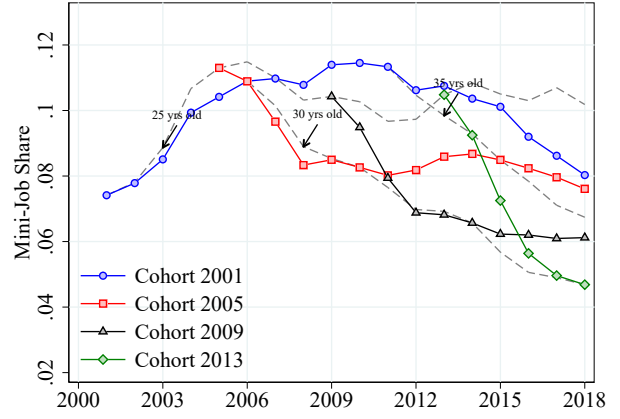
(B) Part-Time Share: Women



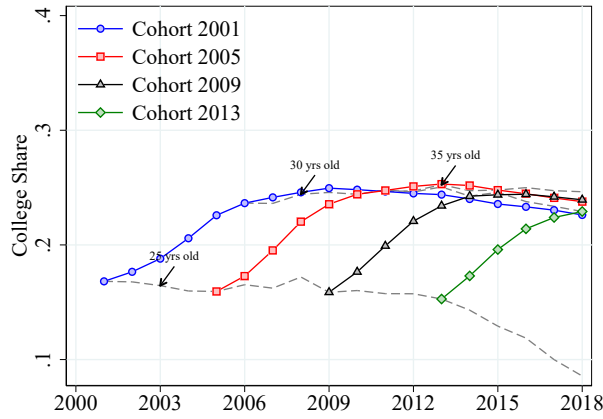
(C) Mini-Job Share: Men



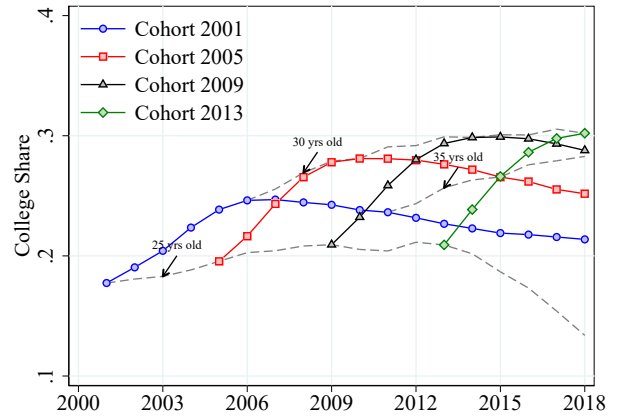
(D) Mini-Job Share: Women



(E) College Share: Men

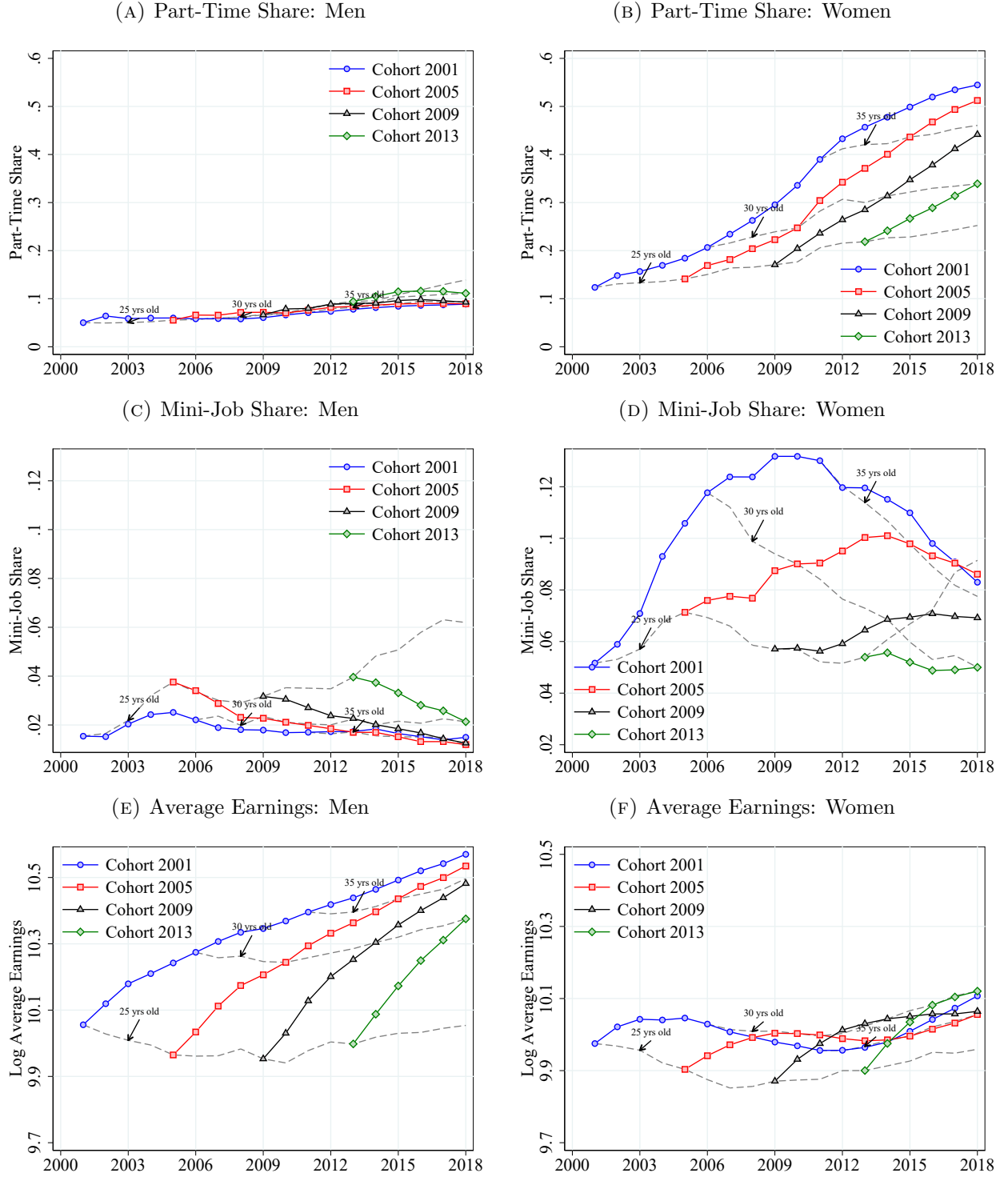


(F) College Share: Women



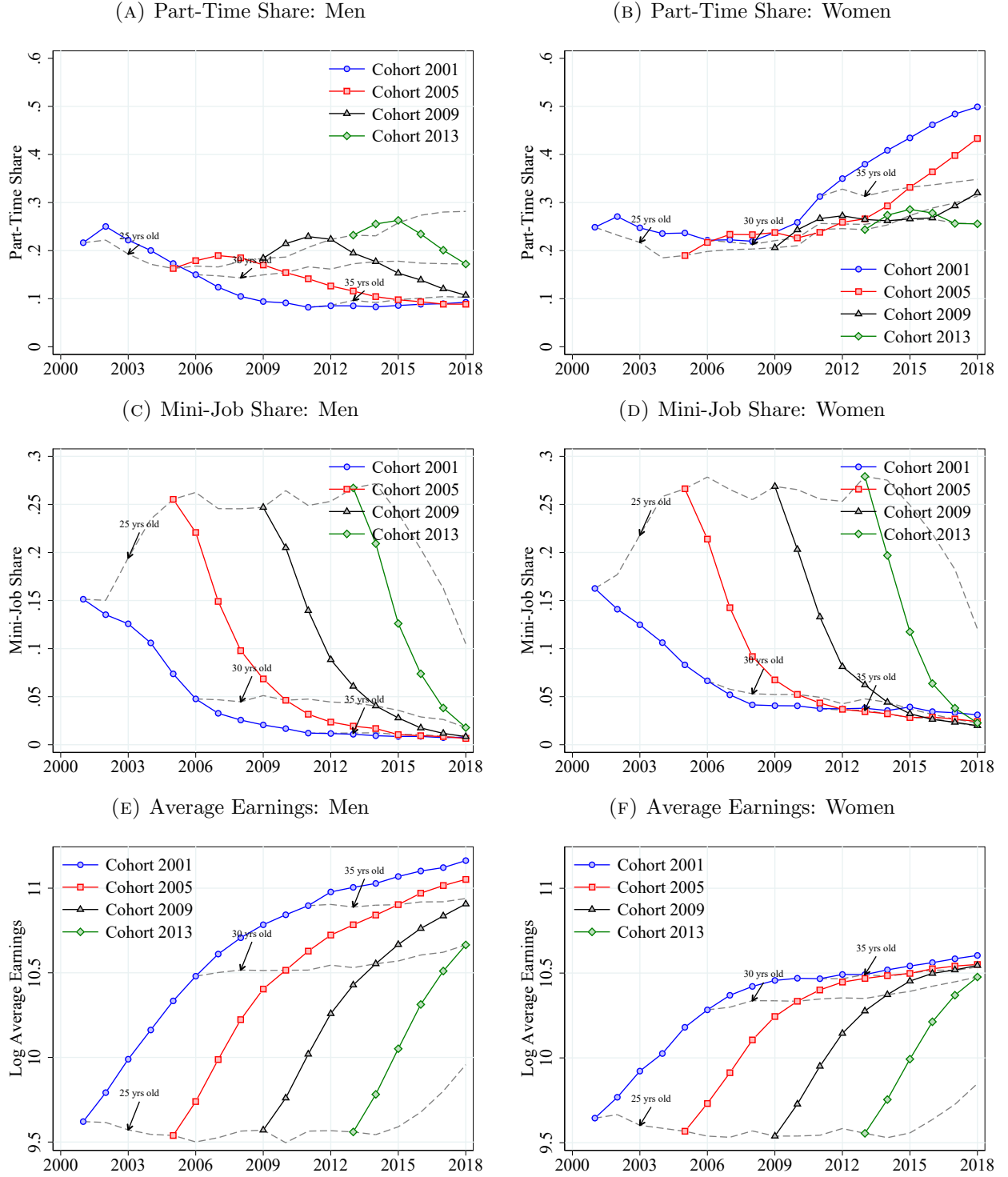
Notes: This figure shows selected employment and education shares in the IAB data (CS sample). Panels A and B show the part-time share over the lifecycle of selected cohorts. Panels C and D show the mini-job share. Panels E and F show the share of college graduates.

FIGURE E.10: EMPLOYMENT LEVELS AND AVERAGE EARNINGS OVER THE LIFECYCLE – NON-COLLEGE WORKERS



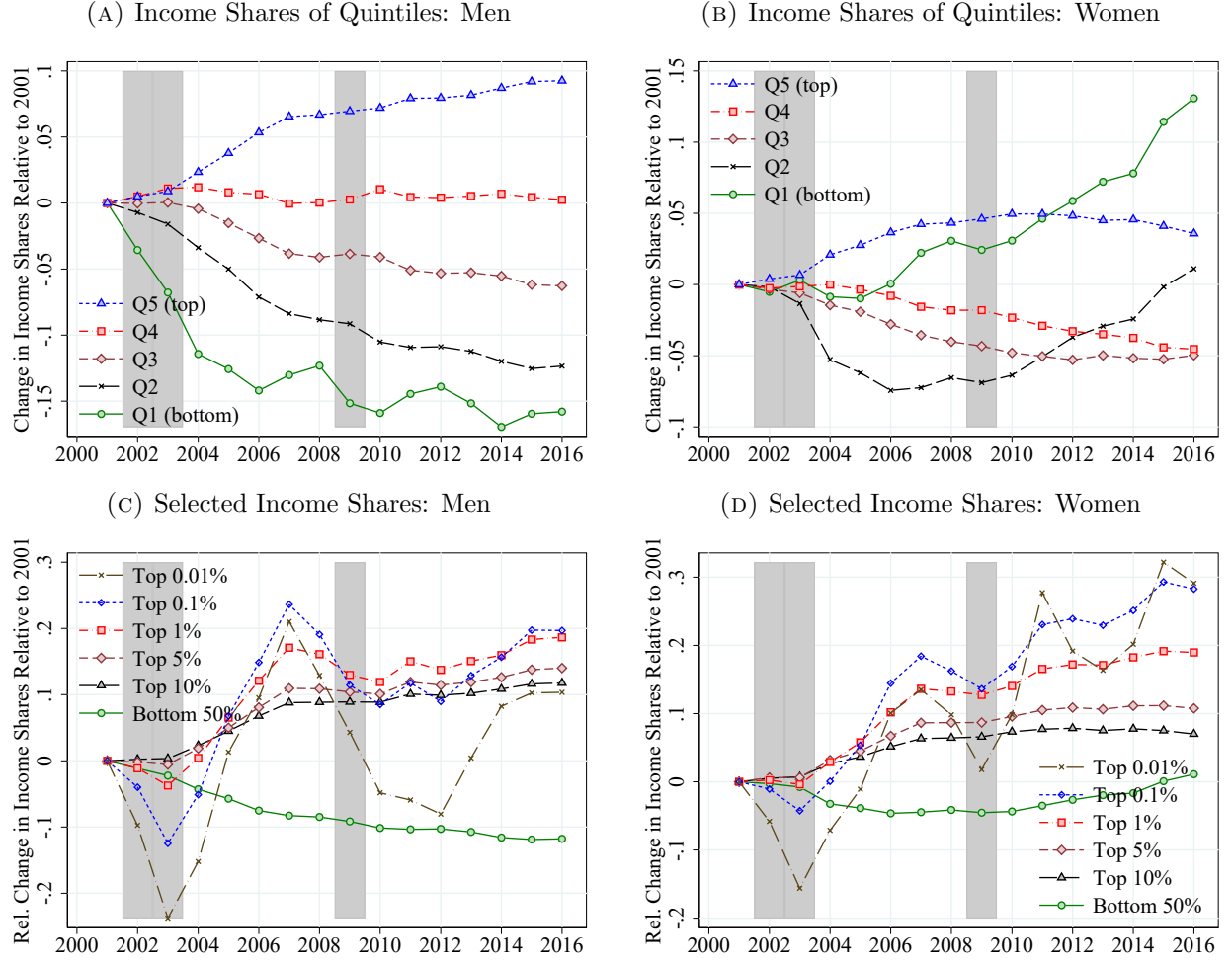
Notes: This figure shows employment levels and average earnings for workers without college degree by cohort in the IAB data (CS sample). Panels A and B show the part-time share over the lifecycle of selected cohorts for non-college workers. Panels C and D show the mini-job share. Panels E and F show average earnings.

FIGURE E.11: EMPLOYMENT LEVELS AND AVERAGE EARNINGS OVER THE LIFECYCLE – COLLEGE WORKERS



Notes: This figure shows employment levels and average earnings for workers with college degree by cohort in the IAB data (CS sample). Panels A and B show the part-time share over the lifecycle of selected cohorts for college workers. Panels C and D show the mini-job share. Panels E and F show average earnings.

FIGURE E.12: CHANGES IN LABOR INCOME SHARES RELATIVE TO 2001



Notes: This figure shows the evolution of selected income shares of real annual earnings (relative to 2001) in the combined IAB-TPP data (CS sample) separately for men and women. The relative change in income shares of each group relative to 2001 is the differences of the income share in year t minus the income share in 2001 divided by the income share in 2001. Shaded areas indicate recessions. See Tables E.1 and E.2 for more details.

TABLE E.1: LABOR INCOME SHARES – MEN

| Year | Q1 | Q2 | Q3 | Q4 | Q5 | Bot 50 | Bot 90 | Mid 40 | Top 10 | Top 5 | Top 1 | Top 0.1 | Top 0.01 |
|------|------|-------|-------|-------|-------|--------|--------|--------|--------|-------|-------|---------|----------|
| 2001 | 5.77 | 13.64 | 18.09 | 22.83 | 39.67 | 27.98 | 75.07 | 47.09 | 24.93 | 15.65 | 5.57 | 1.48 | 0.46 |
| 2002 | 5.57 | 13.55 | 18.08 | 22.94 | 39.87 | 27.66 | 75.01 | 47.35 | 24.99 | 15.62 | 5.50 | 1.42 | 0.41 |
| 2003 | 5.38 | 13.43 | 18.09 | 23.08 | 40.02 | 27.35 | 74.98 | 47.63 | 25.02 | 15.56 | 5.36 | 1.29 | 0.35 |
| 2004 | 5.11 | 13.18 | 18.01 | 23.10 | 40.60 | 26.77 | 74.49 | 47.72 | 25.51 | 15.94 | 5.59 | 1.40 | 0.39 |
| 2005 | 5.05 | 12.96 | 17.81 | 23.01 | 41.17 | 26.38 | 73.95 | 47.57 | 26.05 | 16.43 | 5.93 | 1.58 | 0.46 |
| 2006 | 4.95 | 12.67 | 17.60 | 22.98 | 41.79 | 25.87 | 73.38 | 47.50 | 26.62 | 16.91 | 6.24 | 1.70 | 0.50 |
| 2007 | 5.02 | 12.50 | 17.39 | 22.82 | 42.27 | 25.66 | 72.88 | 47.22 | 27.12 | 17.36 | 6.52 | 1.82 | 0.55 |
| 2008 | 5.06 | 12.44 | 17.34 | 22.83 | 42.33 | 25.60 | 72.86 | 47.26 | 27.14 | 17.35 | 6.46 | 1.76 | 0.52 |
| 2009 | 4.90 | 12.40 | 17.39 | 22.89 | 42.43 | 25.42 | 72.85 | 47.43 | 27.15 | 17.27 | 6.29 | 1.65 | 0.48 |
| 2010 | 4.85 | 12.21 | 17.34 | 23.06 | 42.53 | 25.14 | 72.86 | 47.72 | 27.14 | 17.22 | 6.23 | 1.60 | 0.43 |
| 2011 | 4.94 | 12.15 | 17.16 | 22.93 | 42.82 | 25.08 | 72.55 | 47.47 | 27.45 | 17.51 | 6.40 | 1.65 | 0.43 |
| 2012 | 4.97 | 12.16 | 17.12 | 22.92 | 42.83 | 25.10 | 72.60 | 47.50 | 27.40 | 17.44 | 6.33 | 1.61 | 0.42 |
| 2013 | 4.90 | 12.11 | 17.13 | 22.95 | 42.91 | 24.97 | 72.52 | 47.55 | 27.48 | 17.51 | 6.40 | 1.67 | 0.46 |
| 2014 | 4.79 | 12.01 | 17.09 | 22.99 | 43.13 | 24.74 | 72.36 | 47.63 | 27.64 | 17.62 | 6.45 | 1.71 | 0.49 |
| 2015 | 4.85 | 11.93 | 16.97 | 22.93 | 43.32 | 24.65 | 72.17 | 47.52 | 27.83 | 17.80 | 6.59 | 1.77 | 0.50 |
| 2016 | 4.86 | 11.96 | 16.95 | 22.88 | 43.35 | 24.68 | 72.14 | 47.46 | 27.86 | 17.84 | 6.60 | 1.77 | 0.50 |

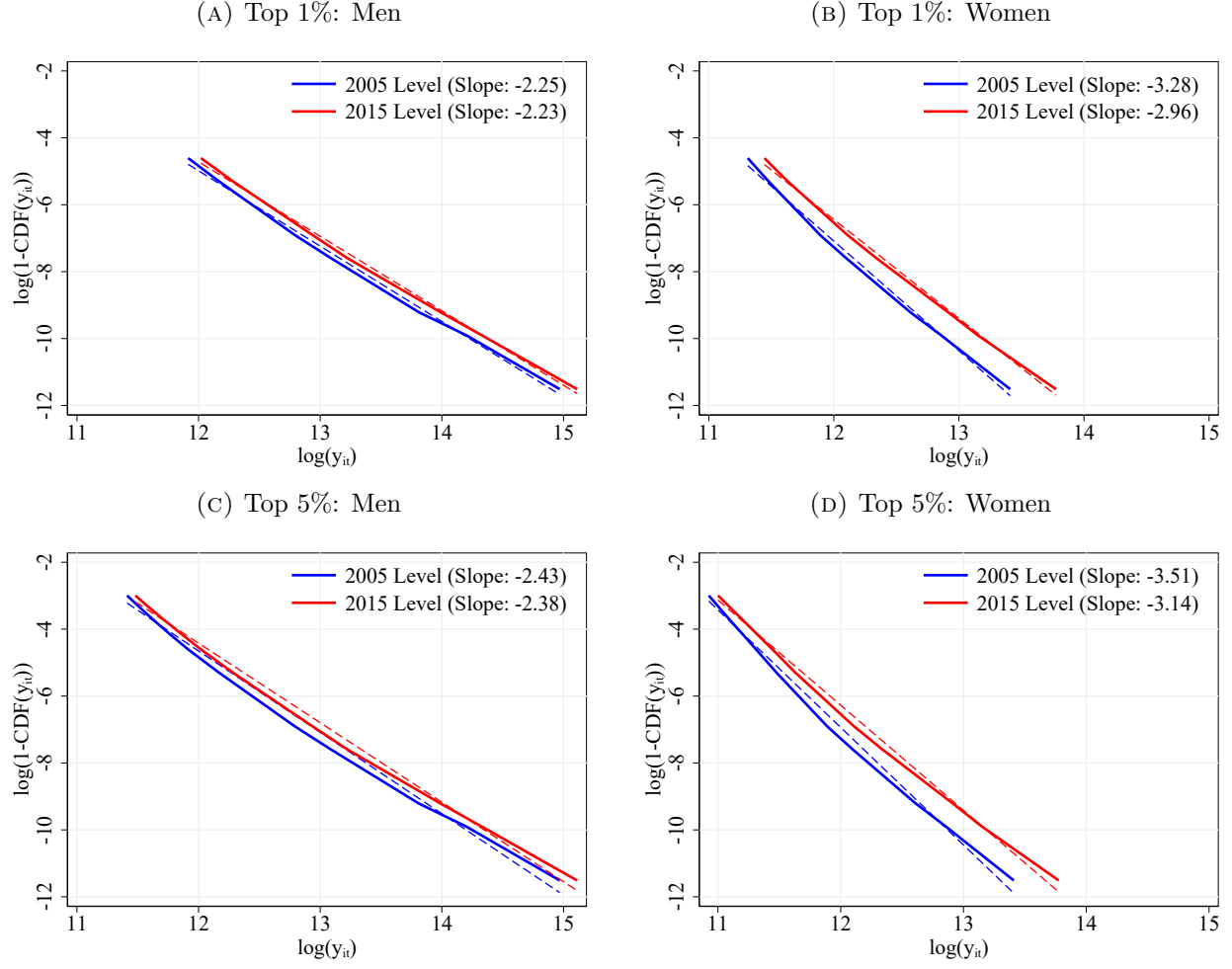
Notes: This table shows the share of earnings that goes to selected parts of the earnings distribution of men in the combined IAB-TPP data (CS sample). Q1 to Q5 refer to the five quintiles where Q1 (Q5) stands for the bottom (top) 20% of the earnings distribution. The quintile shares sum to one. Bot 50, Bot 90 and Mid 40 refer to observations in the bottom 50%, the bottom 90% and between the median and the 90th percentile of the earnings distribution. Top x refers to the top x % of the earnings distribution.

TABLE E.2: LABOR INCOME SHARES – WOMEN

| Year | Q1 | Q2 | Q3 | Q4 | Q5 | Bot 50 | Bot 90 | Mid 40 | Top 10 | Top 5 | Top 1 | Top 0.1 | Top 0.01 |
|------|------|-------|-------|-------|-------|--------|--------|--------|--------|-------|-------|---------|----------|
| 2001 | 3.71 | 10.78 | 18.08 | 26.13 | 41.30 | 22.62 | 75.50 | 52.88 | 24.50 | 14.41 | 4.27 | 0.83 | 0.19 |
| 2002 | 3.70 | 10.77 | 18.02 | 26.06 | 41.46 | 22.56 | 75.37 | 52.81 | 24.63 | 14.50 | 4.28 | 0.82 | 0.18 |
| 2003 | 3.73 | 10.64 | 17.97 | 26.09 | 41.58 | 22.45 | 75.32 | 52.87 | 24.68 | 14.50 | 4.25 | 0.79 | 0.16 |
| 2004 | 3.68 | 10.21 | 17.81 | 26.12 | 42.16 | 21.89 | 74.85 | 52.95 | 25.15 | 14.87 | 4.39 | 0.83 | 0.17 |
| 2005 | 3.68 | 10.11 | 17.73 | 26.03 | 42.44 | 21.75 | 74.61 | 52.86 | 25.39 | 15.06 | 4.52 | 0.87 | 0.19 |
| 2006 | 3.72 | 9.98 | 17.57 | 25.92 | 42.81 | 21.58 | 74.24 | 52.66 | 25.76 | 15.38 | 4.70 | 0.95 | 0.21 |
| 2007 | 3.80 | 10.00 | 17.43 | 25.72 | 43.05 | 21.61 | 73.95 | 52.33 | 26.05 | 15.66 | 4.85 | 0.98 | 0.21 |
| 2008 | 3.83 | 10.08 | 17.35 | 25.65 | 43.09 | 21.69 | 73.93 | 52.24 | 26.07 | 15.66 | 4.83 | 0.96 | 0.21 |
| 2009 | 3.80 | 10.04 | 17.29 | 25.65 | 43.21 | 21.60 | 73.89 | 52.29 | 26.11 | 15.66 | 4.81 | 0.94 | 0.19 |
| 2010 | 3.83 | 10.10 | 17.21 | 25.52 | 43.35 | 21.64 | 73.70 | 52.07 | 26.30 | 15.79 | 4.87 | 0.97 | 0.21 |
| 2011 | 3.89 | 10.24 | 17.16 | 25.37 | 43.35 | 21.83 | 73.61 | 51.78 | 26.39 | 15.93 | 4.98 | 1.02 | 0.24 |
| 2012 | 3.93 | 10.38 | 17.12 | 25.27 | 43.30 | 22.03 | 73.58 | 51.55 | 26.42 | 15.98 | 5.00 | 1.02 | 0.22 |
| 2013 | 3.98 | 10.47 | 17.18 | 25.21 | 43.16 | 22.18 | 73.66 | 51.48 | 26.34 | 15.94 | 5.00 | 1.02 | 0.22 |
| 2014 | 4.00 | 10.52 | 17.14 | 25.15 | 43.19 | 22.25 | 73.60 | 51.35 | 26.40 | 16.02 | 5.05 | 1.03 | 0.23 |
| 2015 | 4.14 | 10.76 | 17.13 | 24.97 | 43.00 | 22.64 | 73.66 | 51.02 | 26.34 | 16.02 | 5.09 | 1.07 | 0.25 |
| 2016 | 4.20 | 10.90 | 17.18 | 24.94 | 42.78 | 22.87 | 73.78 | 50.91 | 26.22 | 15.96 | 5.08 | 1.06 | 0.24 |

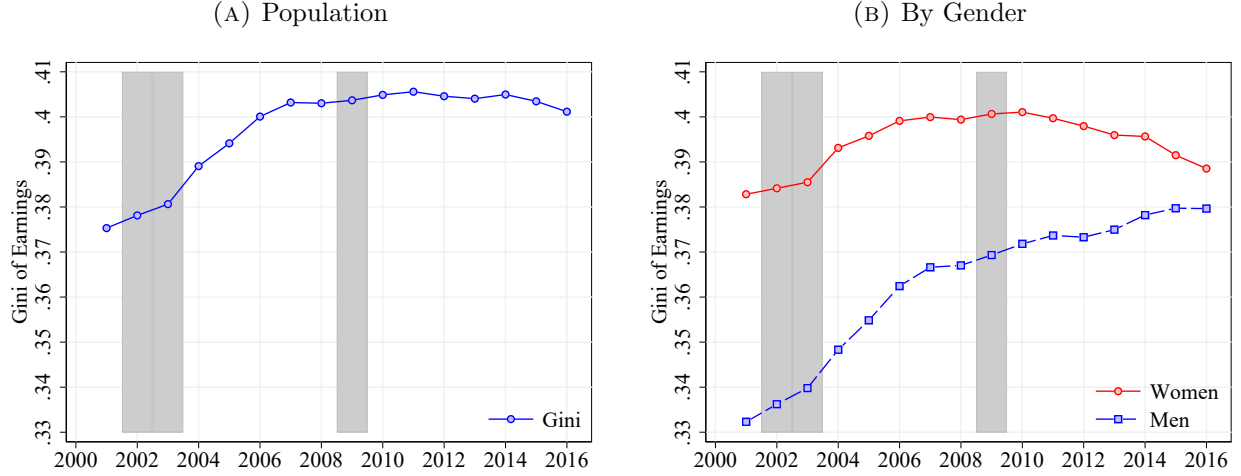
Notes: This table shows the share of earnings that goes to selected parts of the earnings distribution of women in the combined IAB-TPP data (CS sample). Q1 to Q5 refer to the five quintiles where Q1 (Q5) stands for the bottom (top) 20% of the earnings distribution. The quintile shares sum to one. Bot 50, Bot 90 and Mid 40 refer to observations in the bottom 50%, the bottom 90% and between the median and the 90th percentile of the earnings distribution. Top x refers to the top x % of the earnings distribution.

FIGURE E.13: TOP EARNINGS INEQUALITY: PARETO TAIL AT TOP 1% AND TOP 5%



Notes: This figure shows the log of the inverse empirical CDF of log earnings and a fitted linear regression line for observations with earnings in the top 1% and top 5% in the combined IAB-TPP data (CS sample). The absolute value of the slope of the regression line is the Pareto parameter above the respective cutoff.

FIGURE E.14: GINI COEFFICIENT OF LABOR INCOME



Notes: This figure shows the Gini coefficient of labor income in the population and by gender in the combined IAB-TTP data (CS sample). Shaded areas indicate recessions. .

E.2 Details on Reweighting Analysis (Section 3.1)

To shed light on the different development of the percentiles in more detail and reveal underlying drivers we use a reweighting similarly to the procedure proposed by DiNardo et al. (1996), henceforth DFL, to analyze the income distribution. We employ the reweighting function keeping different observable characteristics fixed at their 2001 value. For e.g. the year 2015, we can now observe the wage density that would have prevailed if employees were still equipped with their 2001 characteristics and received wages of 2015. The reweighting function is given by:

$$\psi_z(z) = \frac{dF(z|t_z = 2001)}{dF(z|t_z = 2015)}, \quad (\text{E.1})$$

where z denotes the respective attribute to be held constant and $F(z|t_z)$ the respective individual distribution of z in year t .

Figure E.15 displays the evolution of the demographic observables age, non-German nationality and educational attainment (2 groups) before reweighting separately for men and women. Mean age increases in the sample until about 2010 before slightly decreasing until 2018 as displayed in Panels A and B. It starts at 39.6 for men and 40.2 for women in 2010, peaks at 40.9 (men) and 41.3 (women) and ends at 40.3 (men) and 41 (women) in 2018.¹⁴ Panels C and D show that the share of non-German citizens is almost constant until 2010 and then almost doubles from 2010 to 2018 for both men and women. It is constantly higher for men (9 to 17.5 percent) than for women (6.5 to 12.5 percent). The share of workers with college degree plotted in Panels E and F, slightly

¹⁴This only holds for our sample with the restriction to prime age workers. The average age of the total population and the age of the workforce constantly increases during this time. The decrease in our sample tends to reflect larger birth cohorts leaving the sample when passing age 55.

increases from 2001 to 2018. For men it increases from 19% to 21% and for women from 15% to 21%.

FIGURE E.15: WEIGHTING VARIABLES: DEMOGRAPHICS

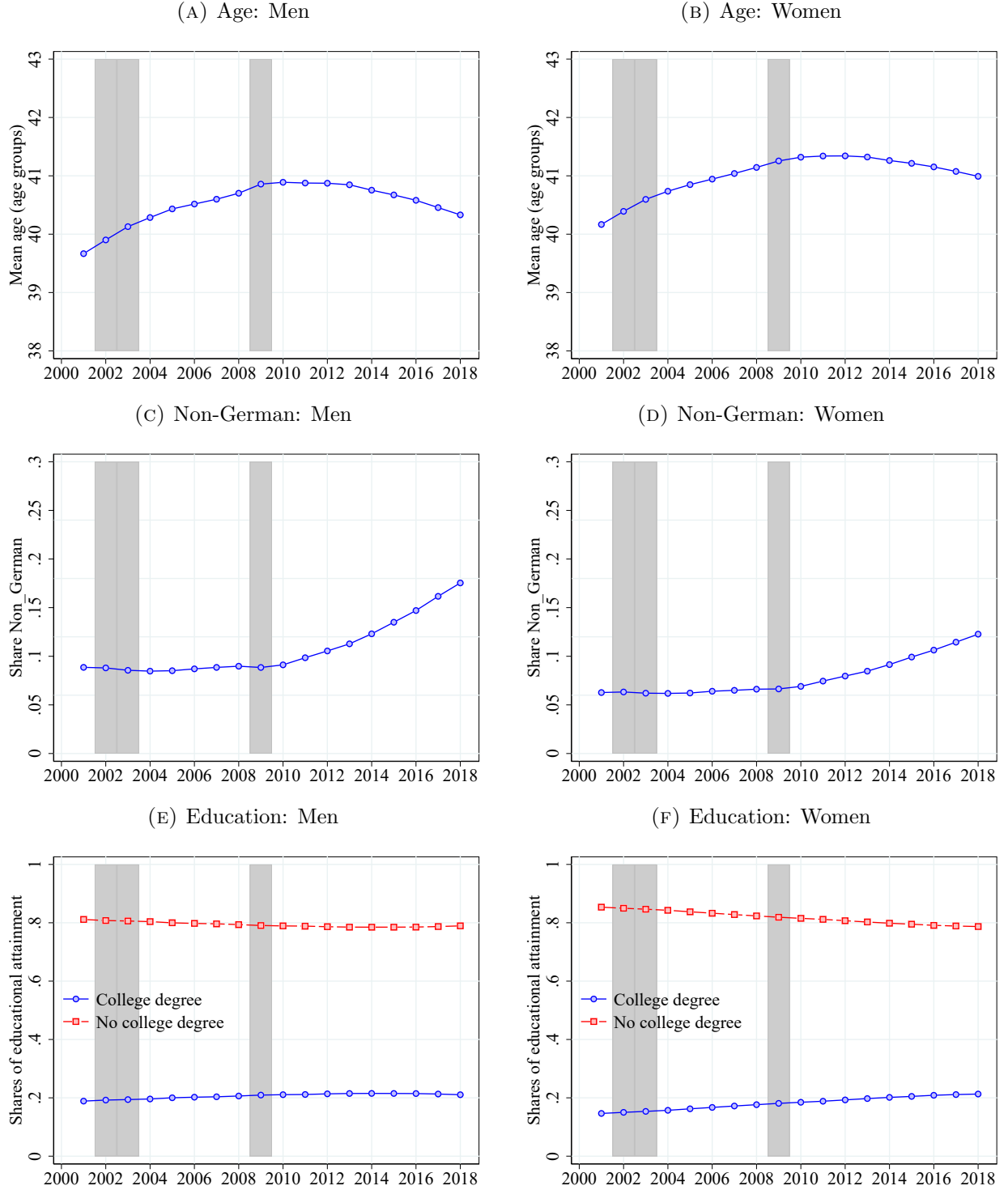
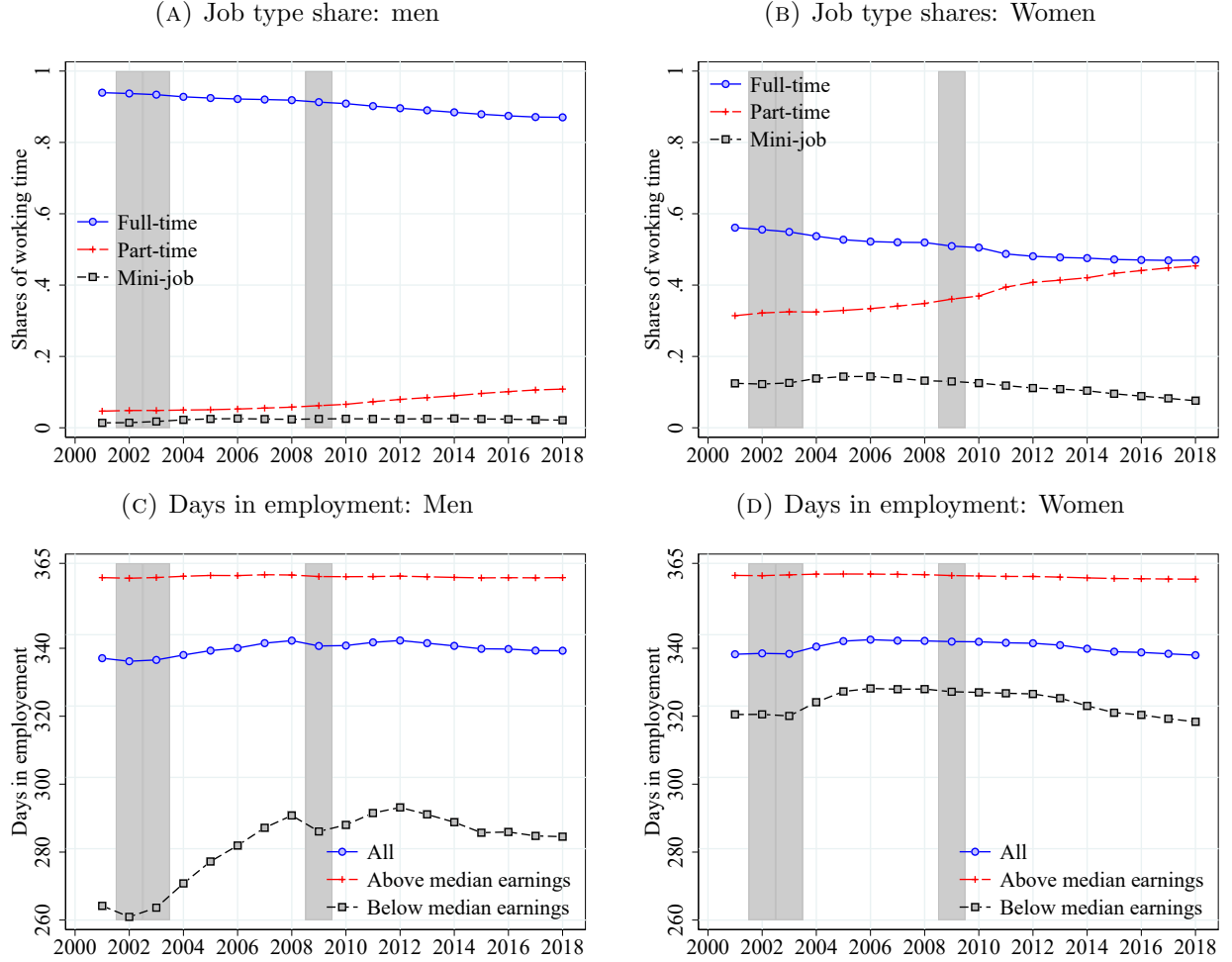


Figure E.16 plots the evolution of work characteristics before applying the weights separately for men and women. In Panels A and B, we show the evolution of full-time, part-time and mini-job shares in our sample before reweighting. The share of full-time workers decreased for men and women. While decreasing, it is consistently higher for men (94% to 87%) than for women (56% to 47%). The share of part-time workers increases over time, from 4.5% (men) and 31.5% (women) in 2001 to 11% (men) and 45.5% (women) in 2018. The share of mini-jobbers is comparatively small (men: 1.5-2.5%, women: 7.5-14.5%). In Panels C and D we depict mean days in employment for men and women for all workers as well as split by median earnings. For men, mean days in employment increase from 337 in 2001 to 342 in 2012 before decreasing again to 339 in 2018. Similarly, days in employment for women increase from 338 in 2001 to 342.5 in 2006 before decreasing again to 338 in 2018. For both genders this changes are almost purely driven by below median earnings workers. For above median earning men, days in employment even decrease slightly while below median earning men experience a notable overall increase from 264 in 2002 to 290 in 2008, 293 in 2012 and then slightly decreasing to 284 in 2018.

FIGURE E.16: WEIGHTING VARIABLES - WORK CHARACTERISTICS

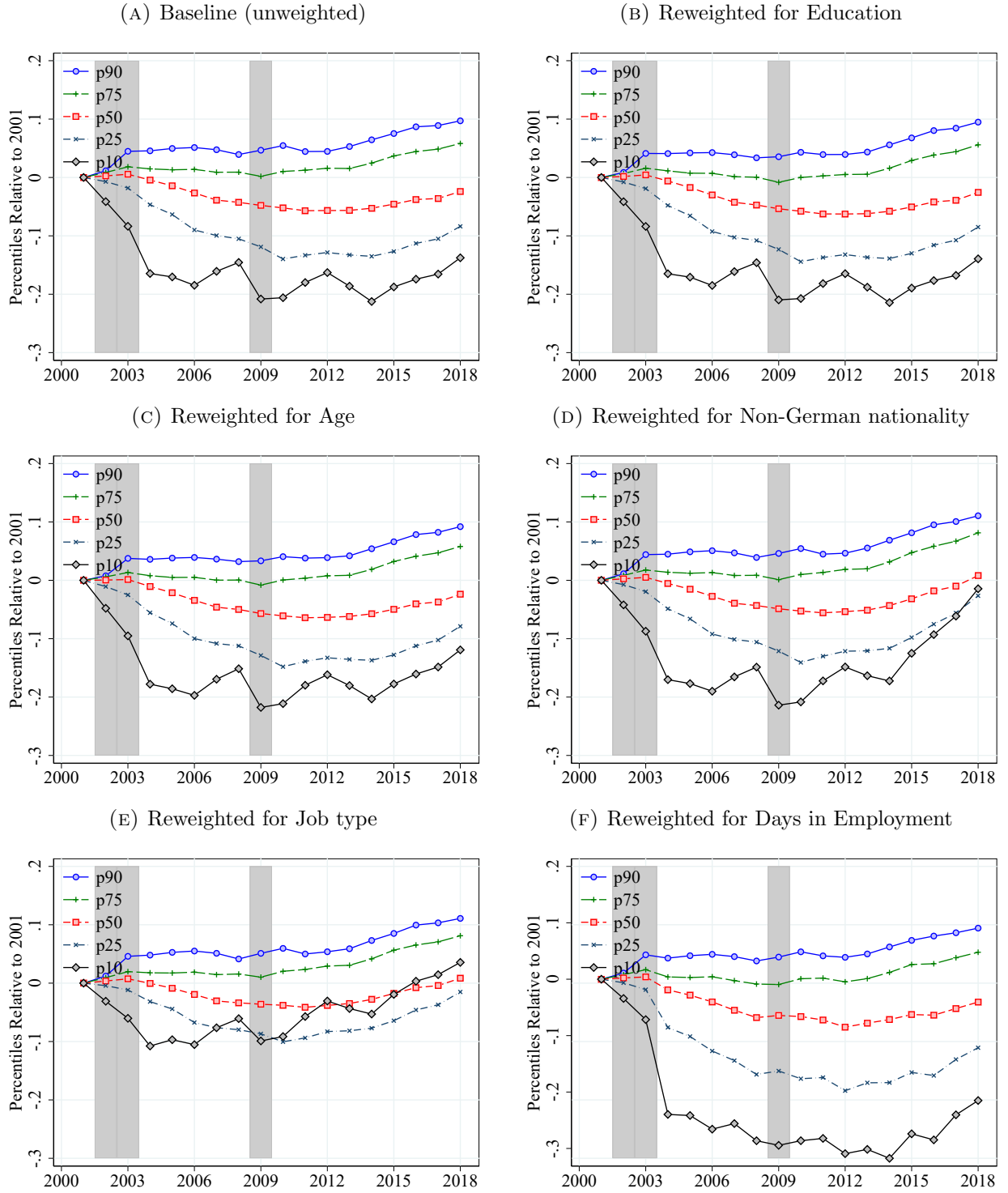


Notes: This figure plots the evolution of work characteristics in the IAB data (CS sample) before and after reweighting for men. For days in employment above and below median earnings, the earnings are weighted by $w = \frac{365}{\text{days in employment}}$ to account for the positive correlation of earnings and days in employment. Thereby, the median is applied to earnings as if every worker would have worked all days. Shaded areas indicate recessions.

In Figures E.17 (for men) and E.18 (for women), we show the evolution of log earnings percentiles before and after reweighting separately by certain demographic and work characteristics. Counterfactual percentiles are constructed by applying the weights obtained using the DFL approach as described above. These figures complement Figure 5 by plotting several percentiles for each reweighted observable in one single graph similarly to Figure 3. Holding age or education constant at their 2001 values appear not to affect percentile evolution patterns much. Keeping non-German nationality constant at initial values moves lower percentile patterns upwards in later years. Thus, earnings inequality would be lower if share of non-Germans would have stayed constant. This is in line with the share of non-Germans being almost constant until 2010 and increasing after 2010 (see Figure E.15). When holding job type (full-time, part-time or mini-job) or days in employment constant over time, we observe more notable changes to percentile evolution patterns.

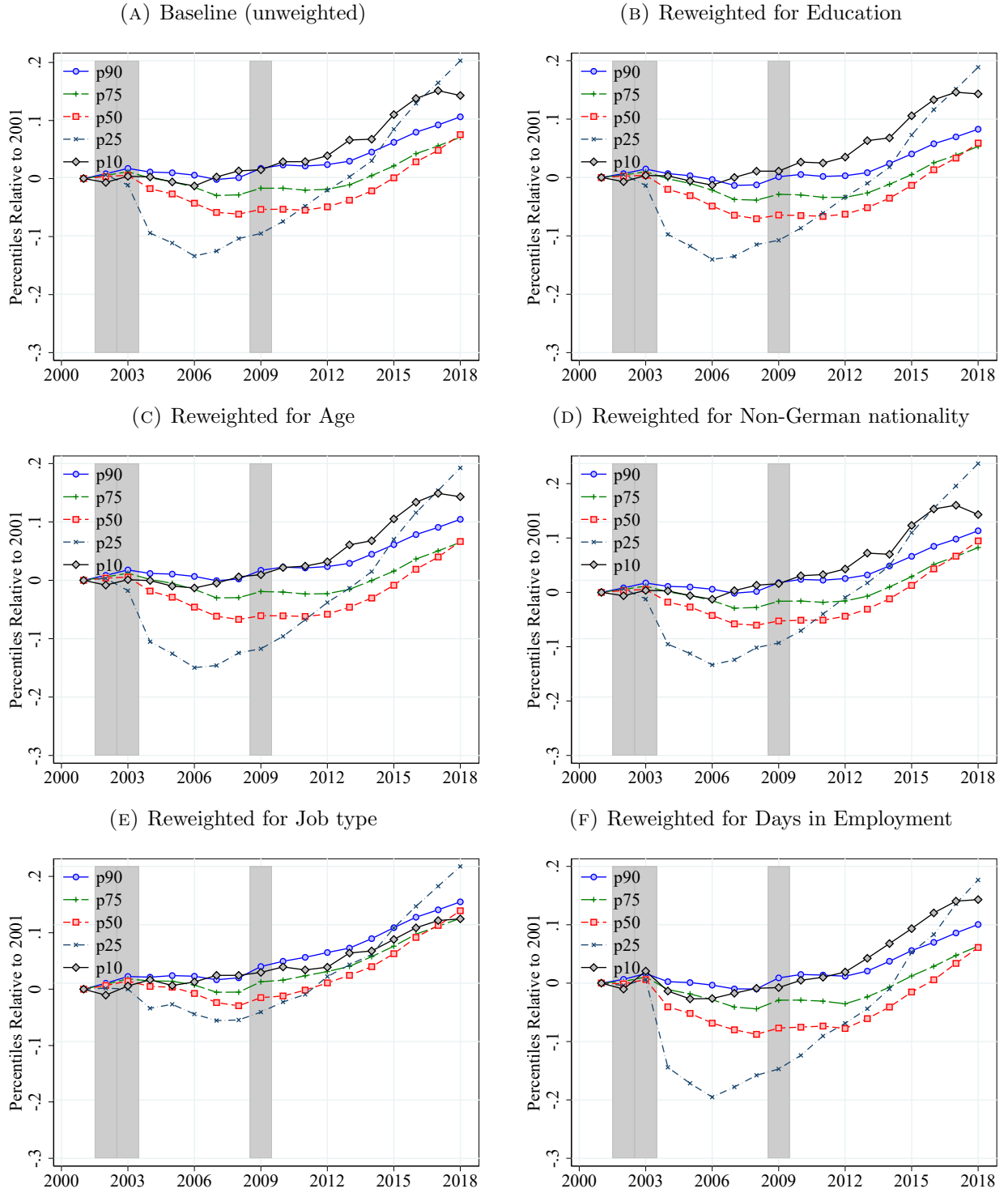
Those tend to affect lower percentiles more. For both, men and women, holding share of full-time, part-time and mini-job workers constant would have compressed the distribution such that percentile evolution appears more compressed. This would have resulted in a more constant evolution of real earnings inequality. The opposite is true for days in employment but almost solely for men. If days in employment would have been remained on (lower) 2001 values (see [E.16](#)), this would have resulted in a more spread evolution of real earnings percentiles and thus higher inequality. The result is in line with days in employment increasing by 15 days between 2001 and 2018 for men earning below-median but slightly decreased by 1 day for above-median earning men. The detailed percentile-wise results of the reweighting analyses are discussed in section [3.1](#).

FIGURE E.17: PERCENTILES OF THE LOG REAL ANNUAL EARNINGS BEFORE AND AFTER REWEIGHTING
– MEN



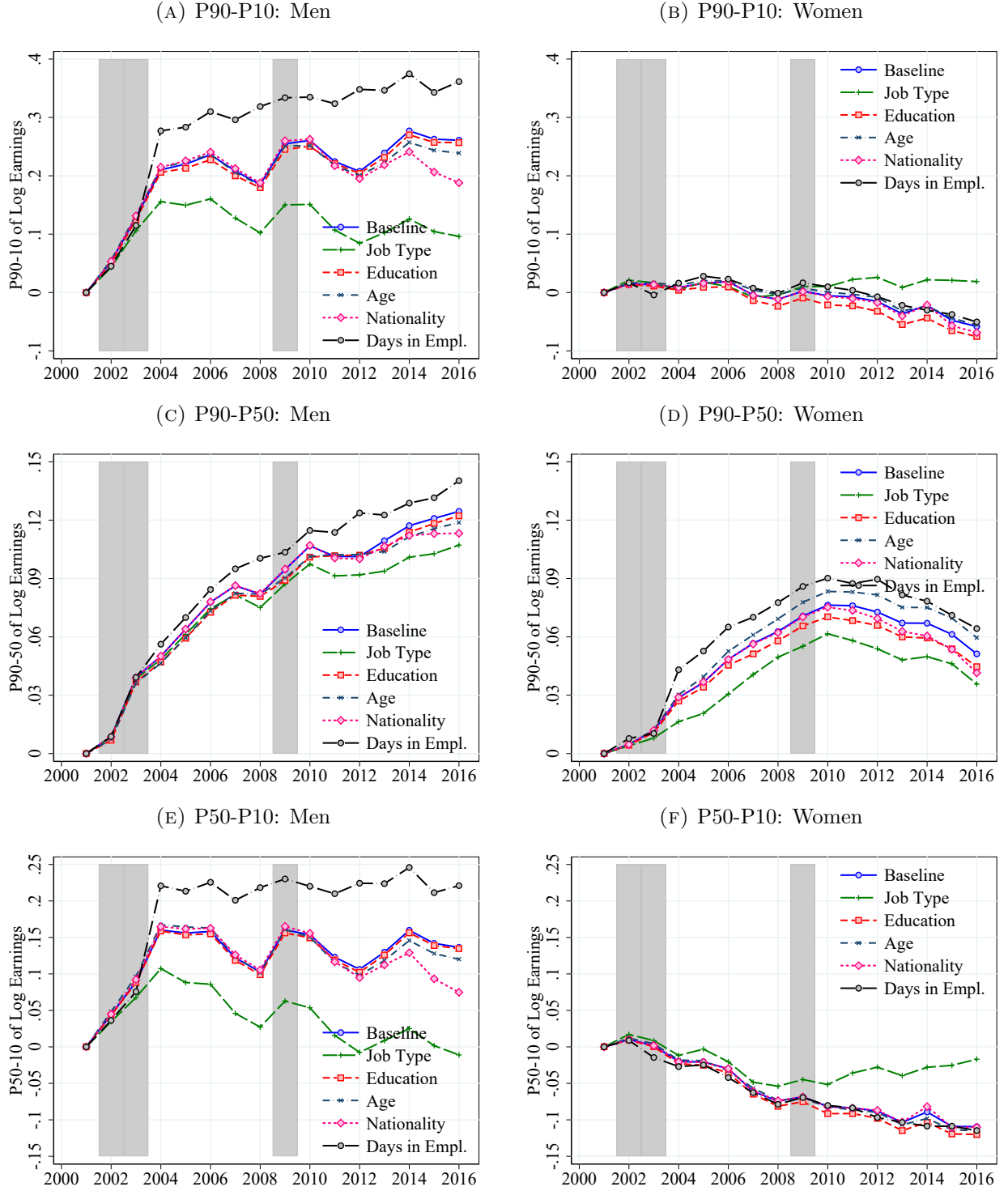
Notes: This figure shows the evolution of different counterfactual log real annual earnings percentiles in the IAB data (CS sample) for men. The counterfactual percentiles are constructed by reweighting the data such that observable dimensions are held constant at the 2001 level. Figure 5 in the main text includes the 10th, 50th and 90th percentile. Shaded areas indicate recessions.

FIGURE E.18: PERCENTILES OF THE LOG REAL ANNUAL EARNINGS BEFORE AND AFTER REWEIGHTING
– WOMEN



Notes: This figure shows the evolution of different counterfactual log real annual earnings percentiles in the IAB data (CS sample) for women. The counterfactual percentiles are constructed by reweighting the data such that observable dimensions are held constant at the 2001 level. Figure 5 in the main text includes the 10th, 50th and 90th percentile. Shaded areas indicate recessions.

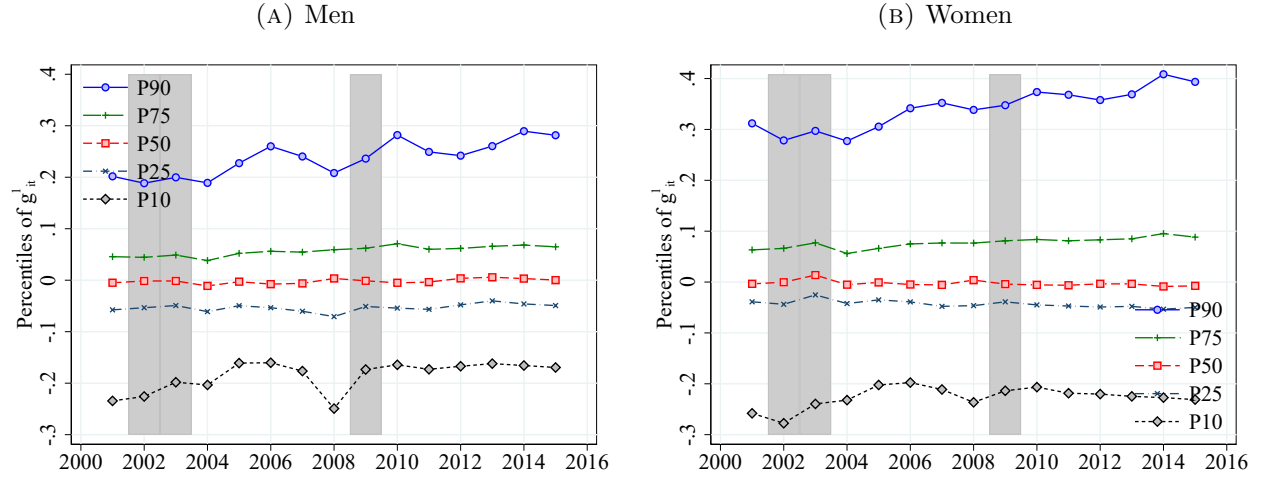
FIGURE E.19: COUNTERFACTUAL EVOLUTION OF LOG EARNINGS PERCENTILE DIFFERENTIALS (REWEIGHTING)



Notes: This figure shows the evolution of different counterfactual percentile differences of the log real annual earnings distribution over time in the IAB data (CS sample) separately for men and women. The counterfactual percentiles are constructed by reweighting the data such that observable dimensions are held constant at the 2001 level. Figure 5 in the main text includes the 10th, 50th and 90th percentile. Shaded areas indicate recessions.

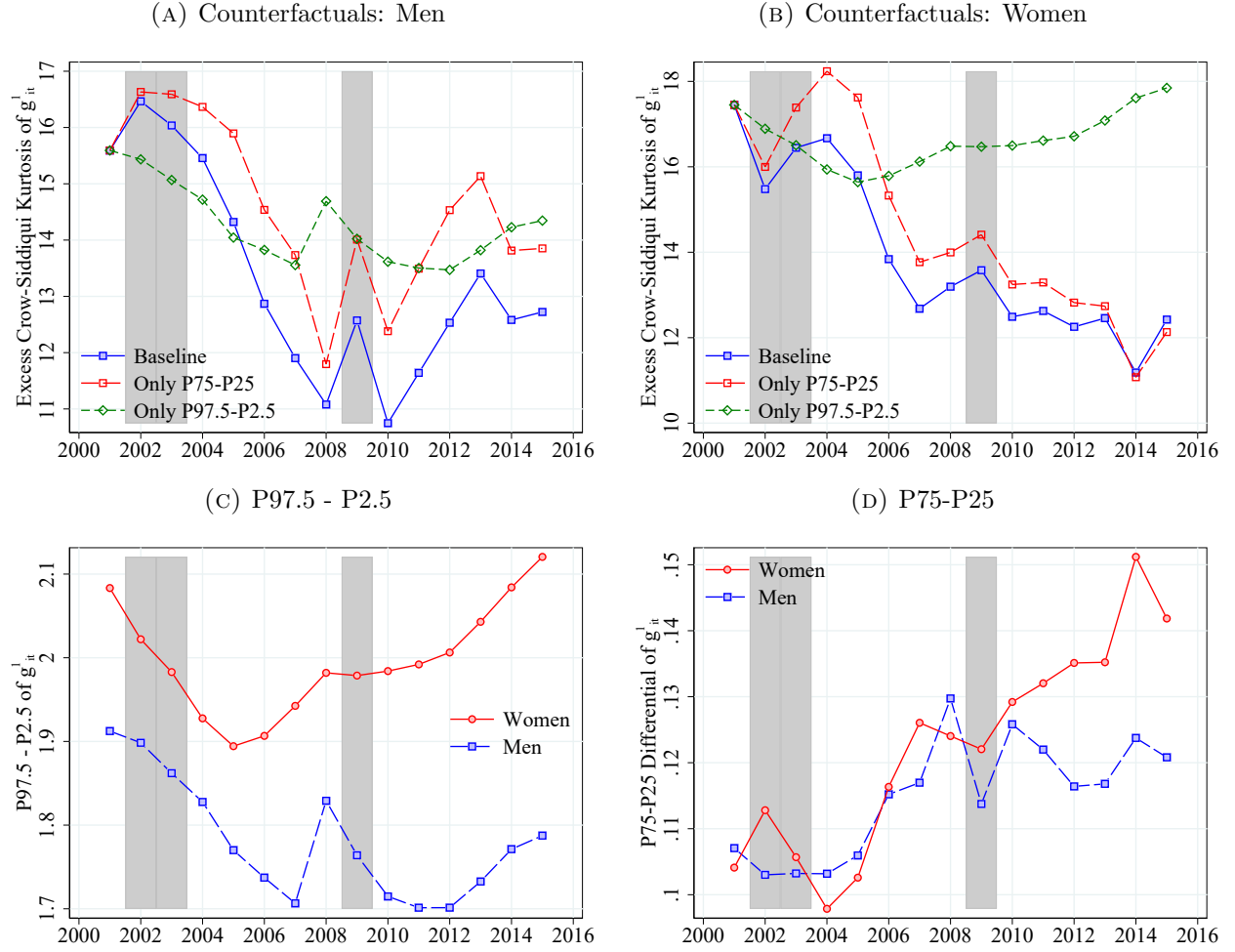
E.3 Additional Results for Earnings Dynamics (Section 3.2)

FIGURE E.20: PERCENTILES OF 1-YEAR LOG EARNINGS CHANGES



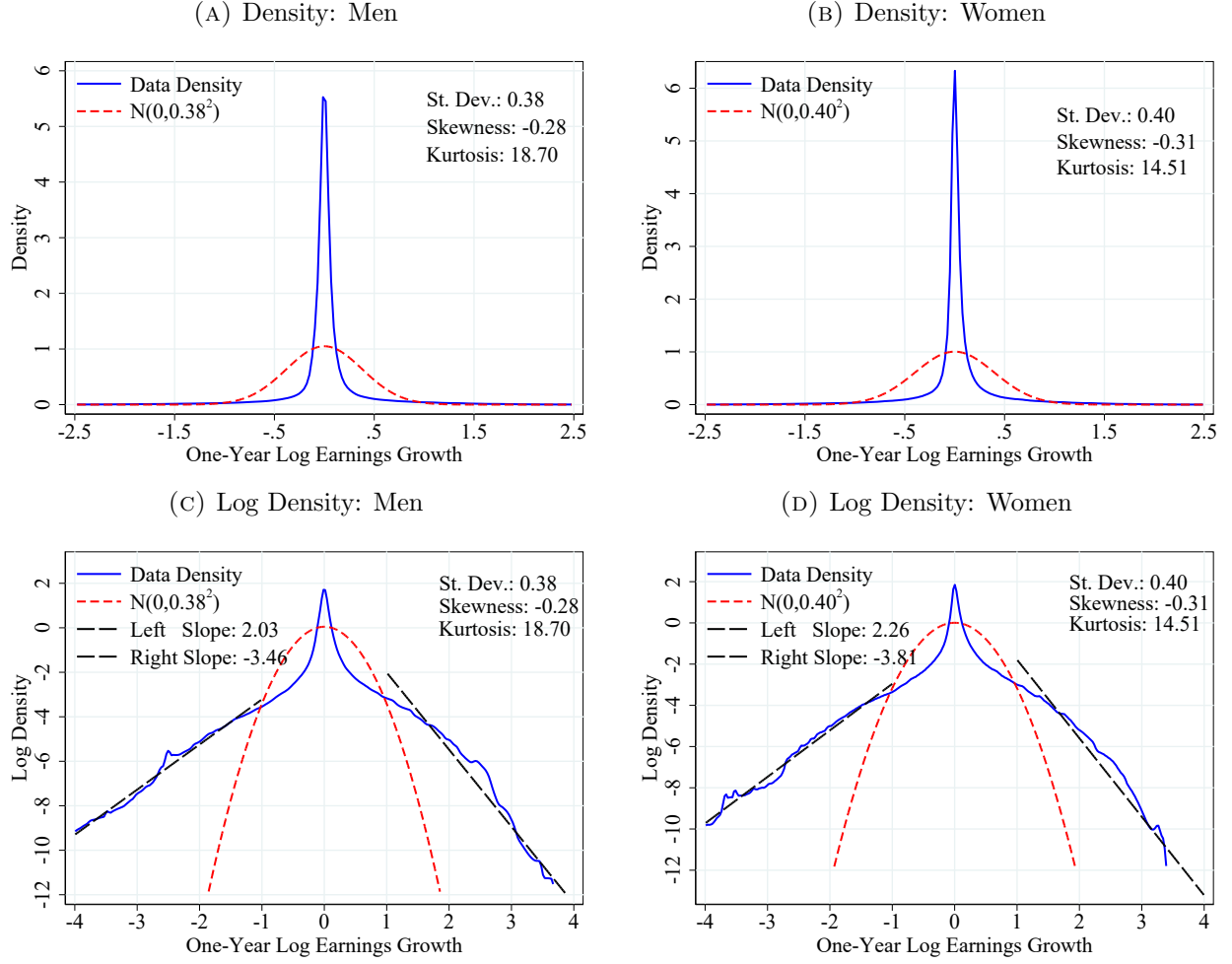
Notes: This figure shows selected percentiles of the distribution of 1-year changes in residualized log real annual earnings (from t to $t + 1$) in the combined IAB-TPP data (LS sample) separately for men and women. Shaded areas indicate recessions. See Appendix D.3 for details on how we construct the distribution of log earnings growth from IAB and TPP data.

FIGURE E.21: DECOMPOSITION OF EXCESS CROW-SIDDIQUI KURTOSIS OF 1-YEAR LOG EARNINGS CHANGES



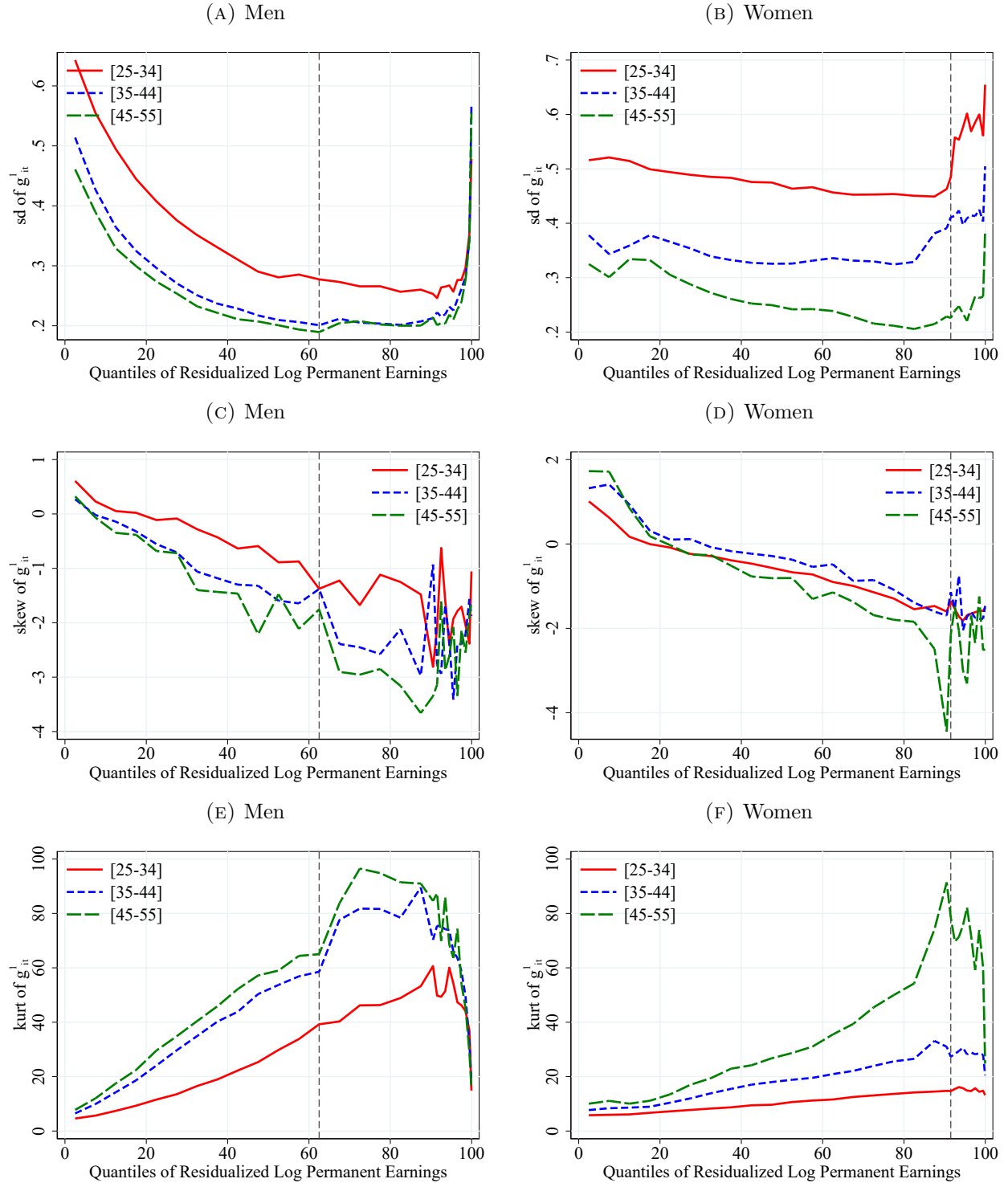
Notes: This figure shows decomposition analyses of the excess Crow-Siddiqui kurtosis in the combined IAB-TPP data (LS sample). Panels A and B show how the excess Crow-Siddiqui kurtosis of 1-year residualized log earnings changes (from t to $t + 1$) would have evolved if only the numerator (P97.5-P2.5) or only the denominator (P75-P25) of the excess Crow-Siddiqui kurtosis would have changed over time. Panels C and D show the evolution of these components. Excess Crow-Siddiqui kurtosis is calculated as $\frac{P_{97.5} - P_{2.5}}{P_{75} - P_{25}} - 2.91$ where the first term is the Crow-Siddiqui measure of kurtosis and 2.91 corresponds to the value of this measure for Normal distribution. Shaded areas indicate recessions. See Appendix D.3 for details on how we construct the distribution of log earnings growth from IAB and TPP data.

FIGURE E.22: DENSITIES OF 1-YEAR LOG EARNINGS CHANGES (YEAR 2005)



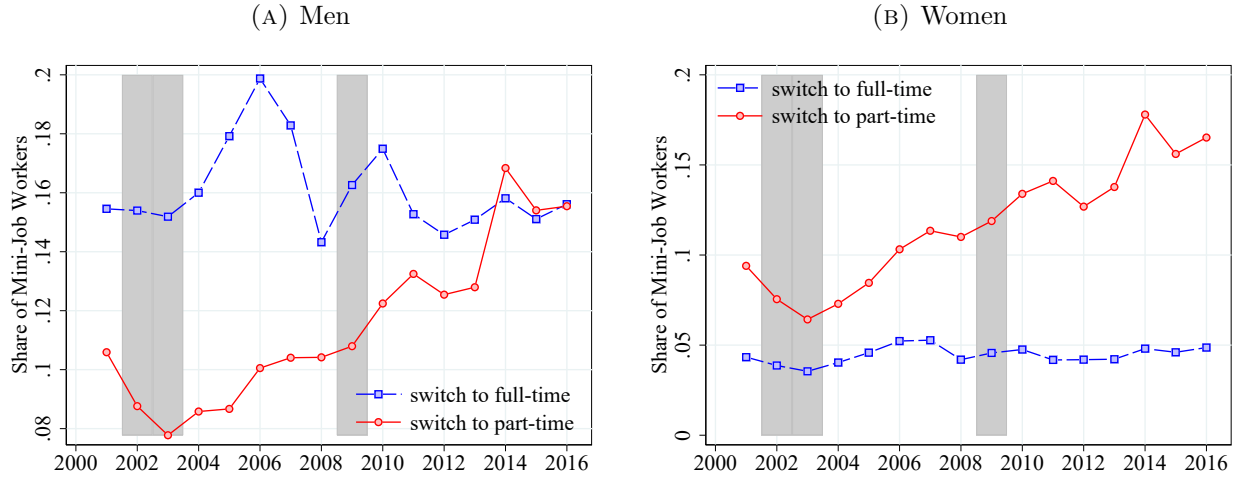
Notes: This figure shows Kernel density estimates of 1-year changes in residualized log earnings for the year 2005 and the respective density of a Normal distribution with zero mean and the same standard deviation as in the combined IAB-TPP data (LS sample).

FIGURE E.23: HETEROGENEITY IN STANDARDIZED MOMENTS OF 1-YEAR LOG EARNINGS CHANGES



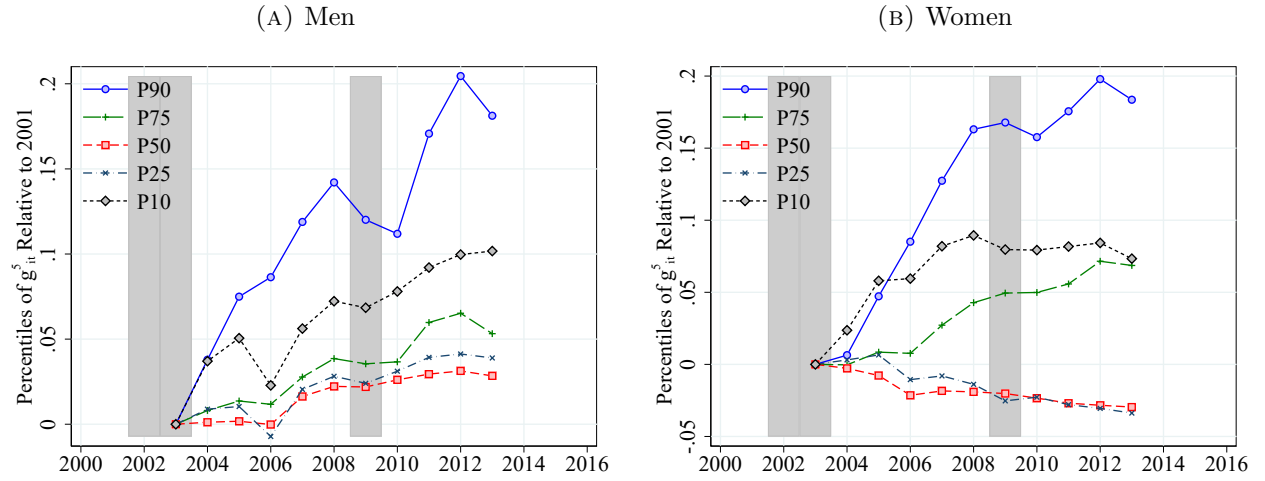
Notes: This figure shows the standard deviation, skewness and excess kurtosis (third and fourth standardized moments) of 1-year changes in residualized log real total income by quantiles of residualized permanent earnings and age groups in the combined IAB-TPP data (H sample) as averages from 2004 to 2011 and separately for men and women. Permanent earnings $P_{i,t-1}$ are defined as the residual (net of a full set of gender and year specific age dummies) of the log of average earnings between $t-3$ and $t-1$. See Footnote 24 definitions and interpretation of Kelley skewness and excess Crow-Siddiqui kurtosis. See Appendix Figures D.11, D.12 and D.13 for a comparison of the underlying data in both data sources.

FIGURE E.24: TRANSITIONS OUT OF MINI-JOBS



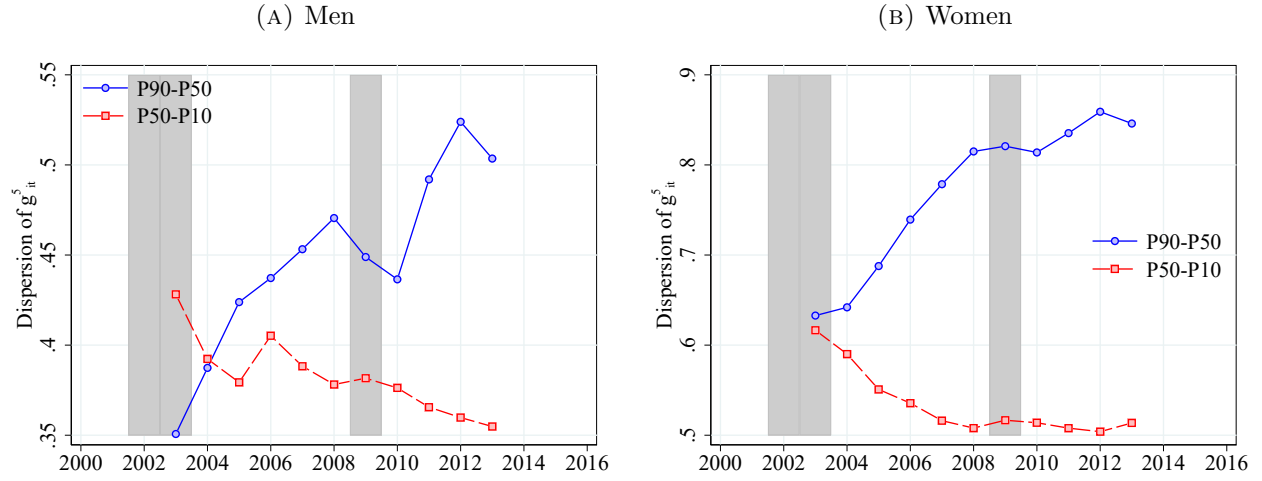
Notes: This figure shows the share of workers who transition from a mini-job to part-time and full-time employment (from t to $t + 1$) in the IAB data (CS sample).

FIGURE E.25: PERCENTILES OF 5-YEAR LOG EARNINGS CHANGES



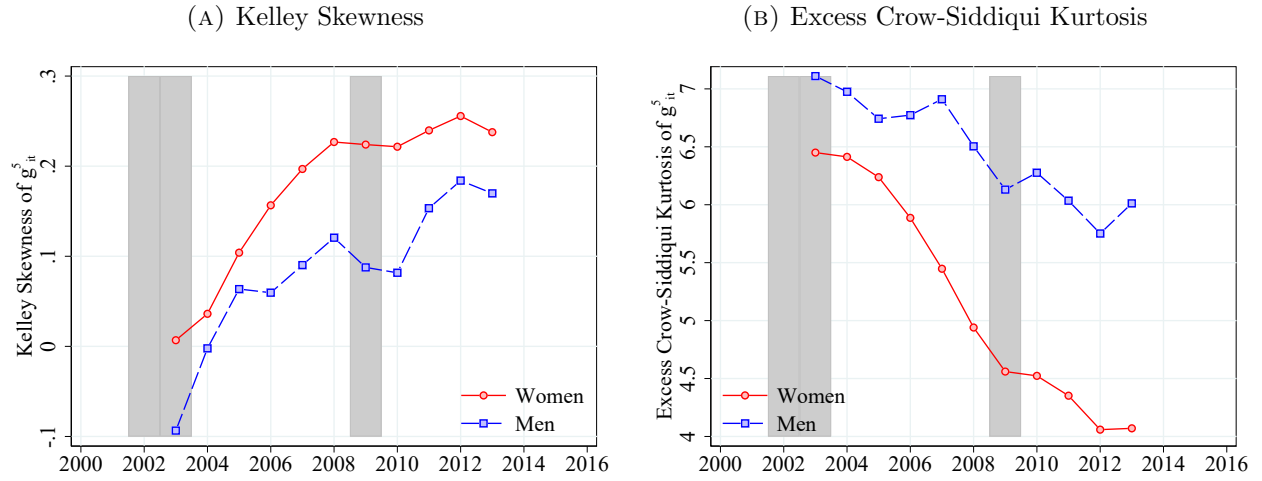
Notes: This figure shows selected percentiles of the distribution of 5-year changes in residualized log real annual earnings (from t to $t + 1$) in the combined IAB-TPP data (LS sample) separately for men and women. Shaded areas indicate recessions. See Appendix D.3 for details on how we construct the distribution of log earnings growth from IAB and TPP data.

FIGURE E.26: DISPERSION OF 5-YEAR LOG EARNINGS CHANGES



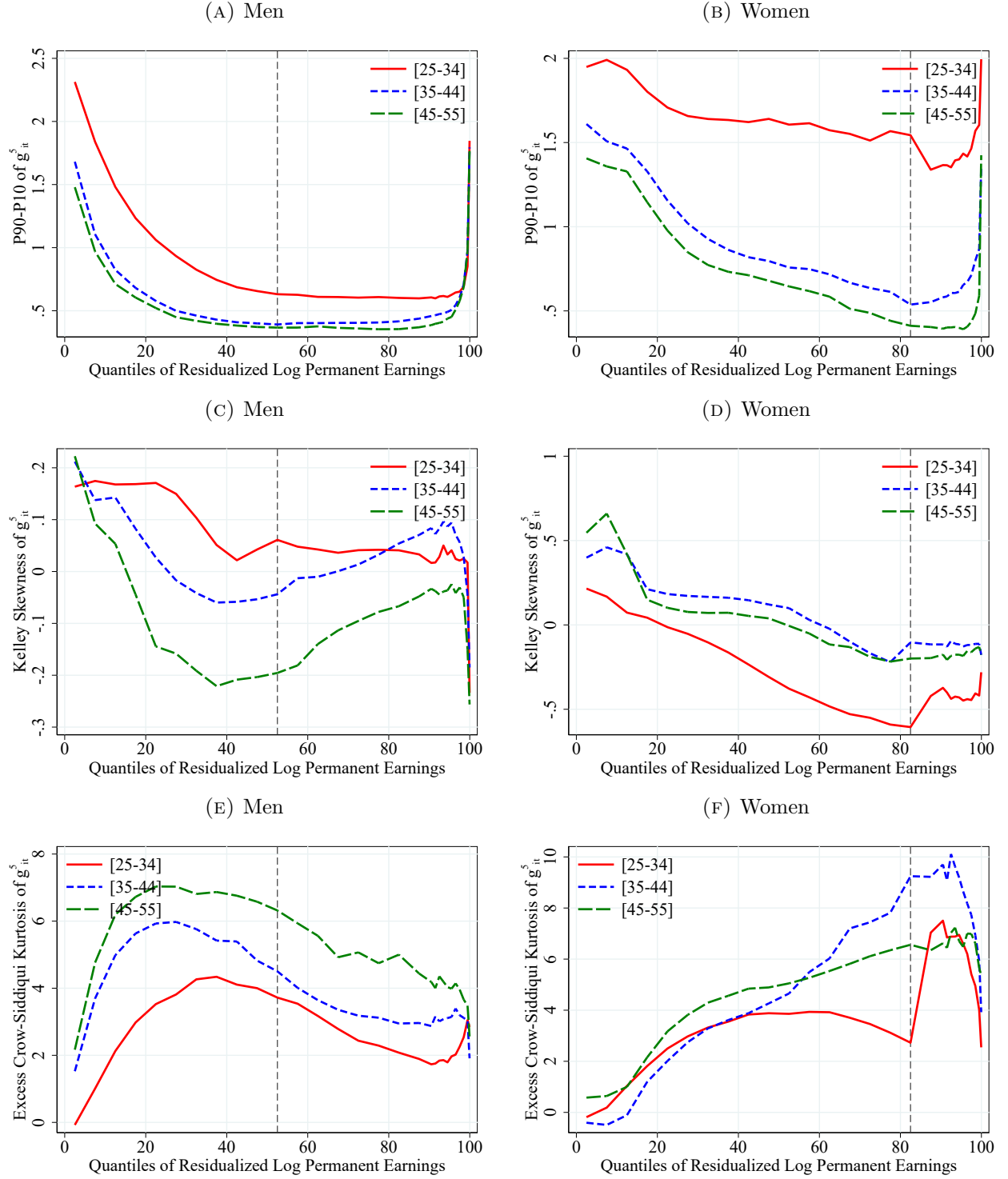
Notes: This figure shows 5-year changes in residualized log earnings (from $t - 2$ to $t + 3$) in the combined IAB-TPP data (LS sample). Shaded areas indicate recessions.

FIGURE E.27: SKEWNESS AND KURTOSIS OF 5-YEAR LOG EARNINGS CHANGES



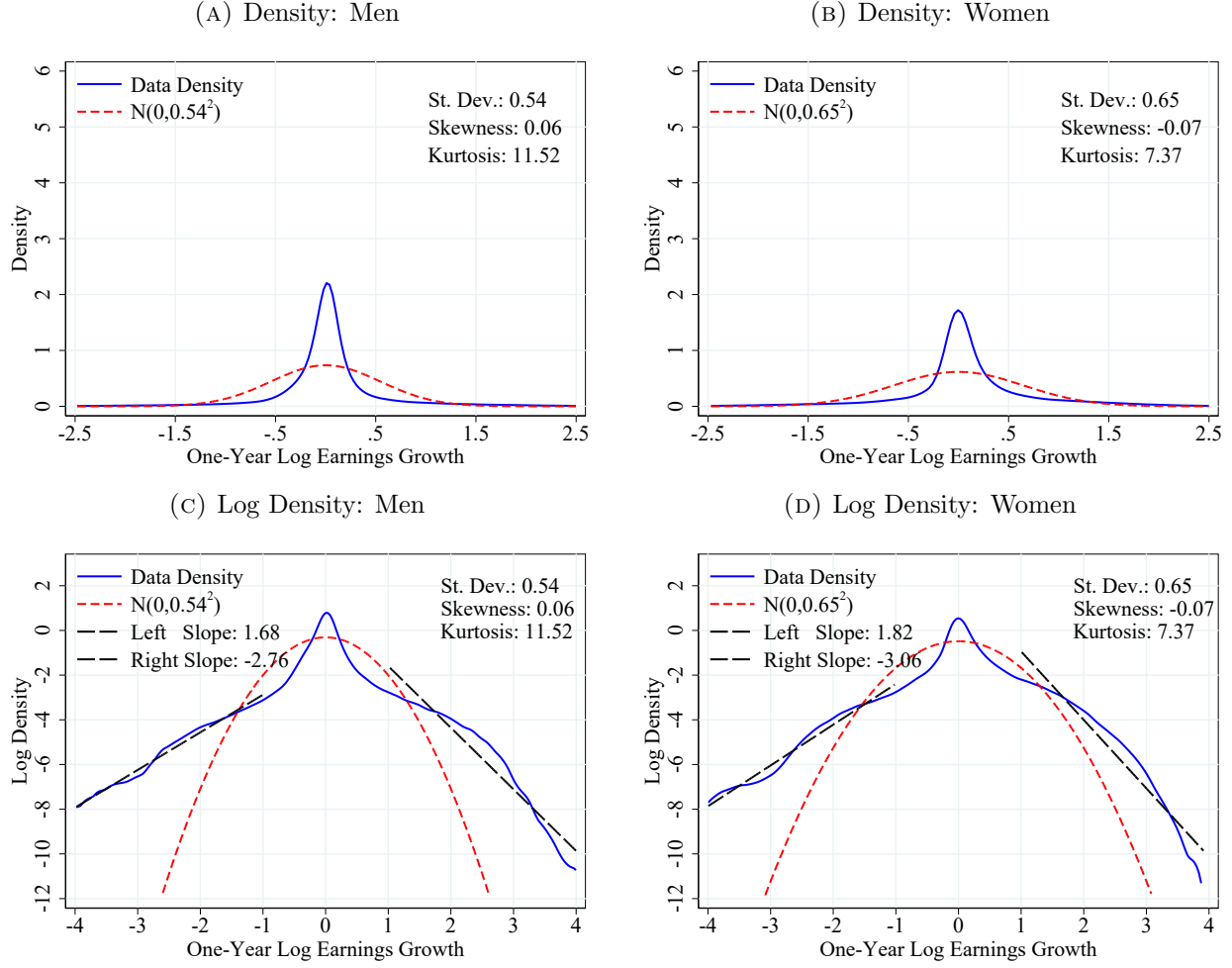
Notes: This figure shows 5-year changes in residualized log earnings (from $t - 2$ to $t + 3$) in the combined IAB-TPP data (LS sample). Kelley skewness is $\frac{P90 - 2P50 + P10}{P90 - P10}$. Excess Crow-Siddiqui kurtosis is calculated as $\frac{P97.5 - P2.5}{P75 - P25} - 2.91$ where the first term is the Crow-Siddiqui measure of kurtosis and 2.91 corresponds to the value of this measure for Normal distribution. Shaded areas indicate recessions.

FIGURE E.28: HETEROGENEITY IN DISPERSION, SKEWNESS AND KURTOSIS OF 5-YEAR LOG EARNINGS CHANGES



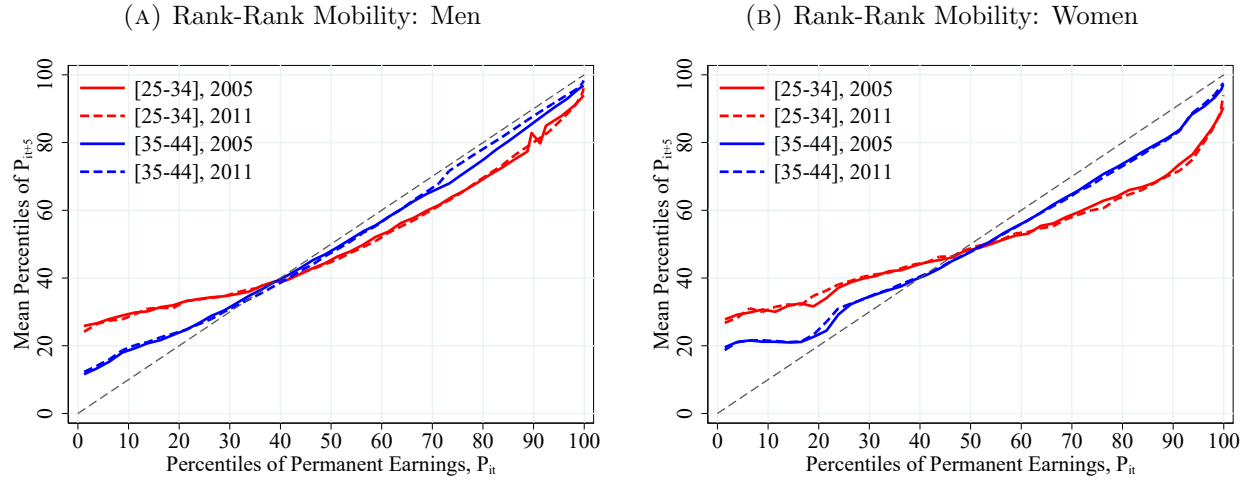
This figure shows the P90-P10 differential, Kelley skewness and excess Crow-Siddiqui kurtosis of 5-year changes in residualized log earnings (from $t - 2$ to $t + 3$) in the combined IAB-TPP data (H sample) as averages from 2004 to 2011 by quantiles of residualized permanent earnings and age groups. Kelley skewness is $\frac{P90 - 2P50 + P10}{P90 - P10}$. Excess Crow-Siddiqui kurtosis is calculated as $\frac{P97.5 - P2.5}{P75 - P25} - 2.91$ where the first term is the Crow-Siddiqui measure of kurtosis and 2.91 corresponds to the value of this measure for Normal distribution. Shaded areas indicate recessions.

FIGURE E.29: DENSITIES OF 5-YEAR LOG EARNINGS CHANGES (YEAR 2005)



Notes: This figure shows Kernel density estimates of 5-year changes in residualized log earnings (from $t-2$ to $t+3$) for the year 2005 and the respective density of a Normal distribution with zero mean and the same standard deviation as in the combined IAB-TPP data (LS sample).

FIGURE E.30: EVOLUTION OF 5-YEAR PERMANENT EARNINGS MOBILITY



Notes: This figure shows the evolution of average 5-year rank-rank mobility of permanent earnings in the combined IAB-TTP data (H sample) as averages from 2004 to 2011, separately for men and women and two different age groups. Permanent income calculated using earnings from $t - 1$, $t - 2$ and $t - 3$.

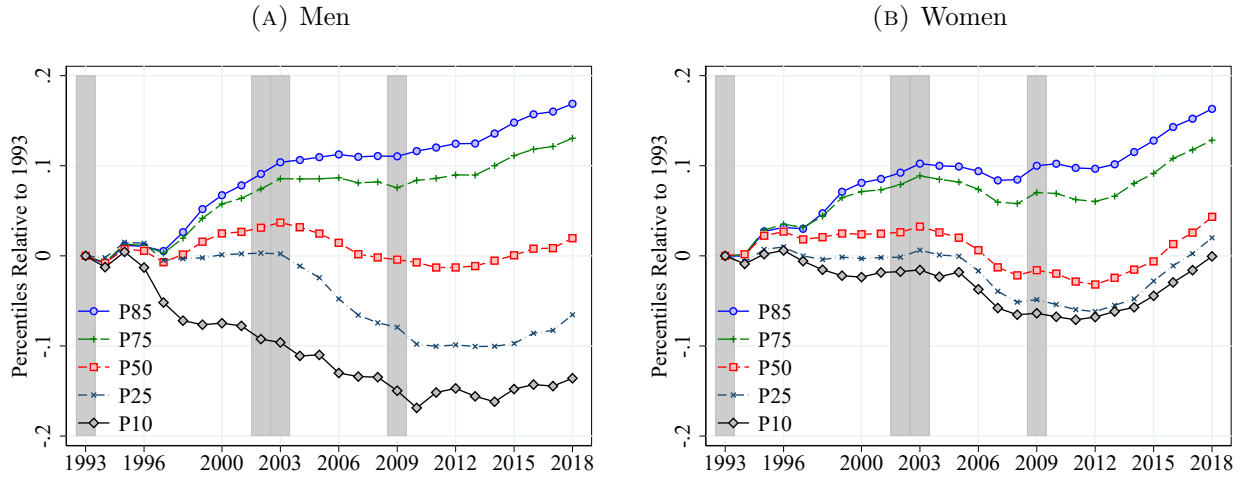
F Core Analysis of Earnings with Longer Samples

In this section, we present figures similar to those of the core analysis of this paper in Section 3 for longer samples based on IAB data only.

F.1 IAB Data 1993–2018

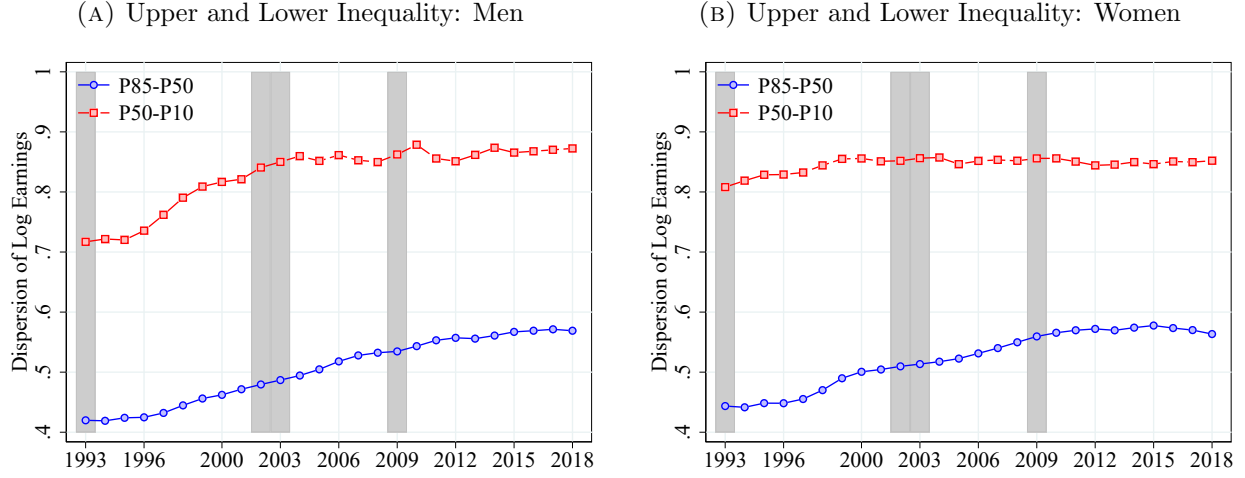
Using data from the IAB for the years 1993 to 2018, we extended the analysis by including several years prior to the sample used in the main section of this paper. To account for changes in mini-job regulations and workforce composition changes due to measurement changes in 1999, the minimum earnings threshold is set to 6,250 Euro annual earnings in 2018 to obtain a consistent sample over the whole time span, i.e. mini-jobs are not included in the longer sample. For men, the wages are imputed from around the P90 upwards, therefore we show the P85 here instead.

FIGURE F.1: EVOLUTION OF LOG EARNINGS PERCENTILES



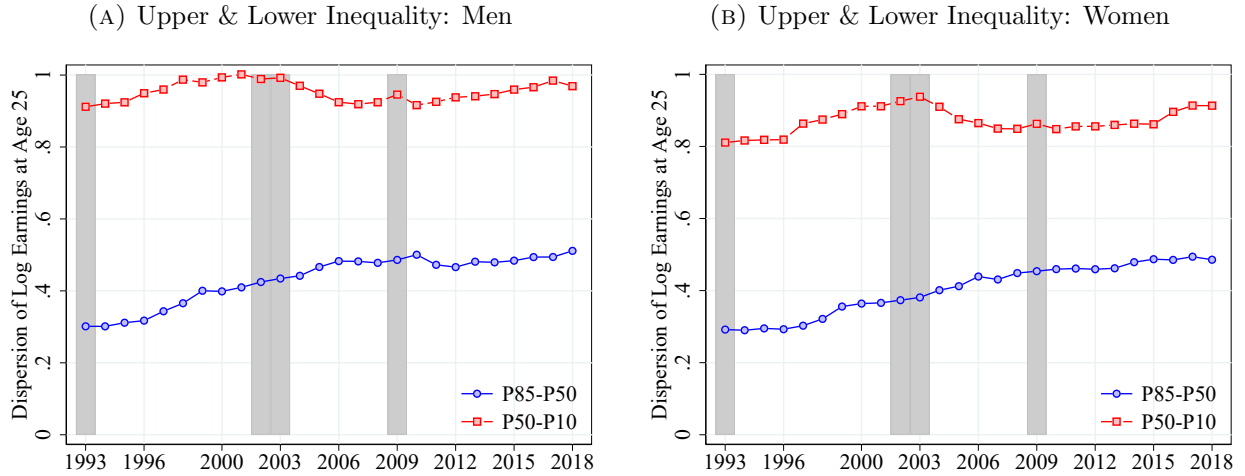
Notes: This figure shows the evolution of selected percentiles of log real earnings from 1993 to 2018 in the IAB data (CS sample, truncated as stated below). The P90 for men is above the top-coding threshold and therefore imputed. All percentiles are normalized to 0 in 1993. Shaded areas indicate recessions. CS sample with minimum income threshold of 6,250 Euro (2018 prices). The CS sample in the main text uses 2,300 Euro as cutoff to include mini-jobs. Shaded areas indicate recessions. The analysis for the core sample is in Figure 3.

FIGURE F.2: EARNINGS INEQUALITY: LOG PERCENTILE DIFFERENTIALS



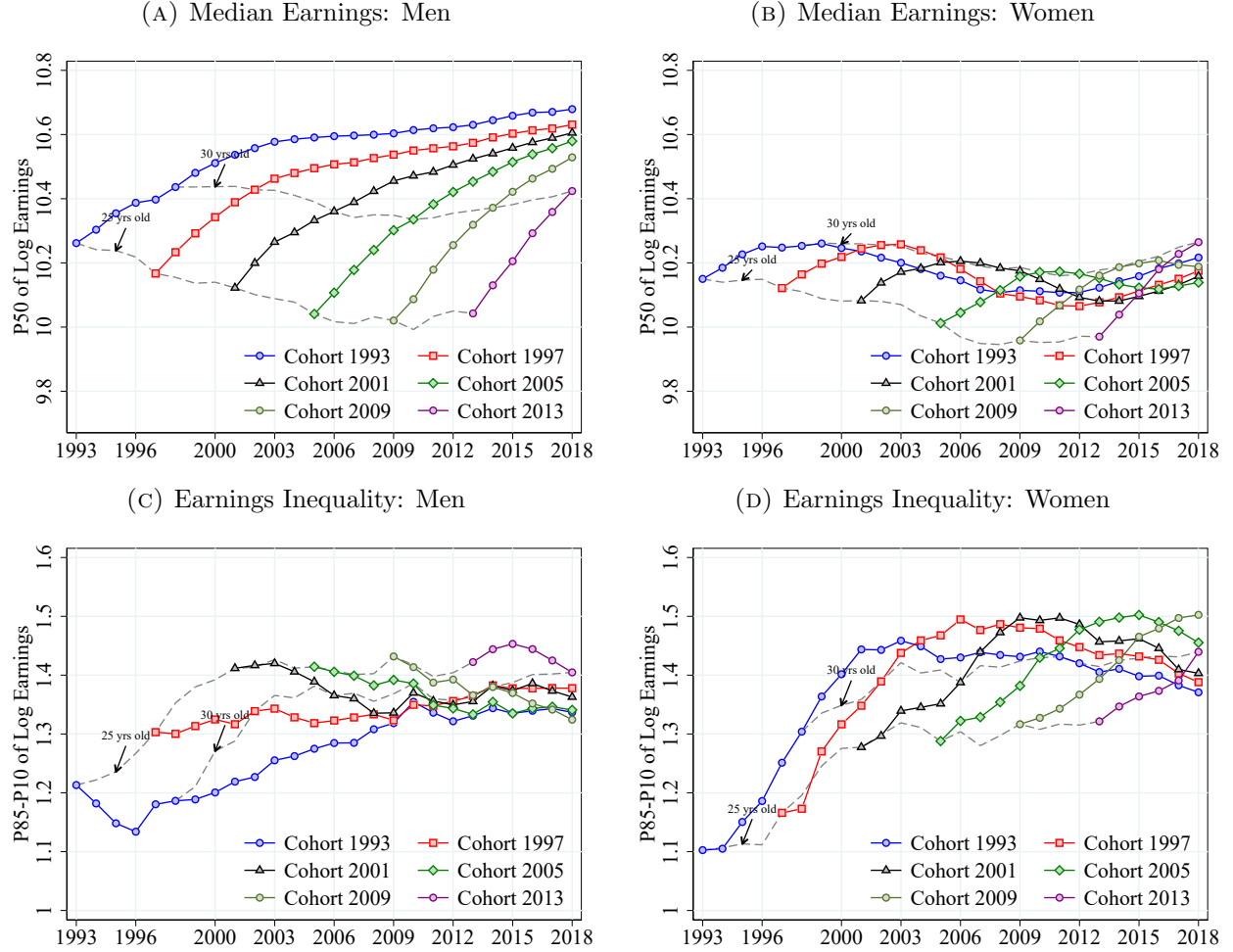
Notes: This figure shows percentile differentials of log real annual earnings in the IAB data (CS sample, truncated as stated below). The P90 for men is above the top-coding threshold and therefore imputed. CS sample with minimum income threshold of 6,250 Euro (2018 prices). The CS sample in the main text uses 2,300 Euro as cutoff to include mini-jobs. Shaded areas indicate recessions. The results of our main sample can be found in Figure 4.

FIGURE F.3: INITIAL INCOME INEQUALITY (AT AGE 25): LOG PERCENTILE DIFFERENTIALS



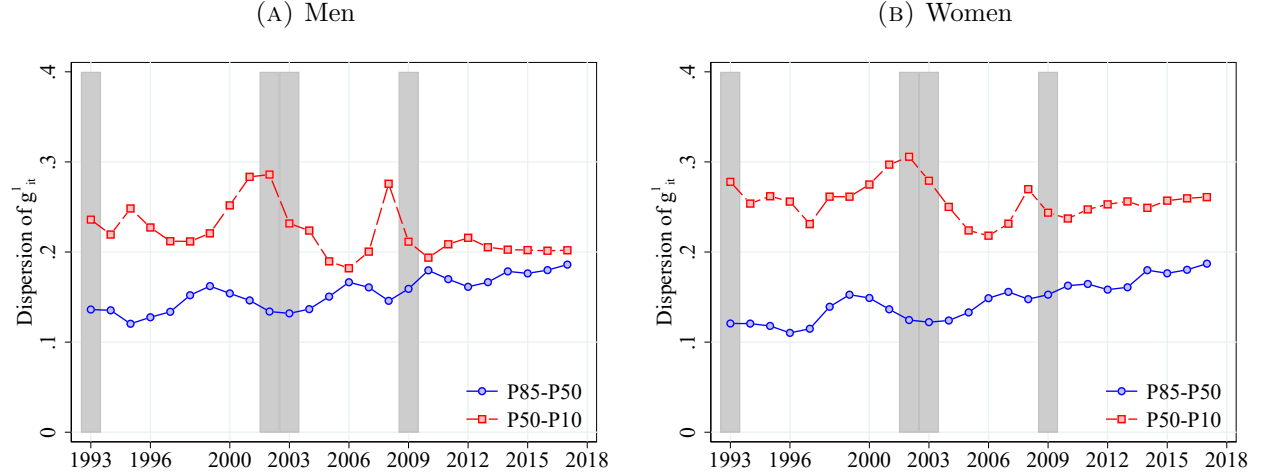
Notes: This figure shows initial inequality at age 25 in the IAB data (CS sample, truncated as stated below). CS sample with a minimum income threshold of 6,250 Euro (2018 prices). The CS sample in the main text uses 2,300 Euro as cutoff to include mini-jobs. The IAB data is top-coded and imputed above about 60,000 Euro, which is above the P90 here. Shaded areas indicate recessions.

FIGURE F.4: EARNINGS PROFILES AND INEQUALITY BY COHORT



Notes: This figure shows the evolution of the median (P50) as well as the P90-P10 differential of the log real annual earnings distribution over time in the IAB data (CS sample, truncated as stated below) separately for men and women. Each colored line corresponds to an individual cohort, where “cohort t ” represents the cohort aged 25 in year t . CS sample with minimum income threshold of 6,250 Euro (2018 prices). The CS sample in the main text uses 2,300 Euro as cutoff to include mini-jobs. Shaded areas indicate recessions. Results for our main sample can be found in Figure 6.

FIGURE F.5: DISPERSION OF 1-YEAR LOG EARNINGS CHANGES

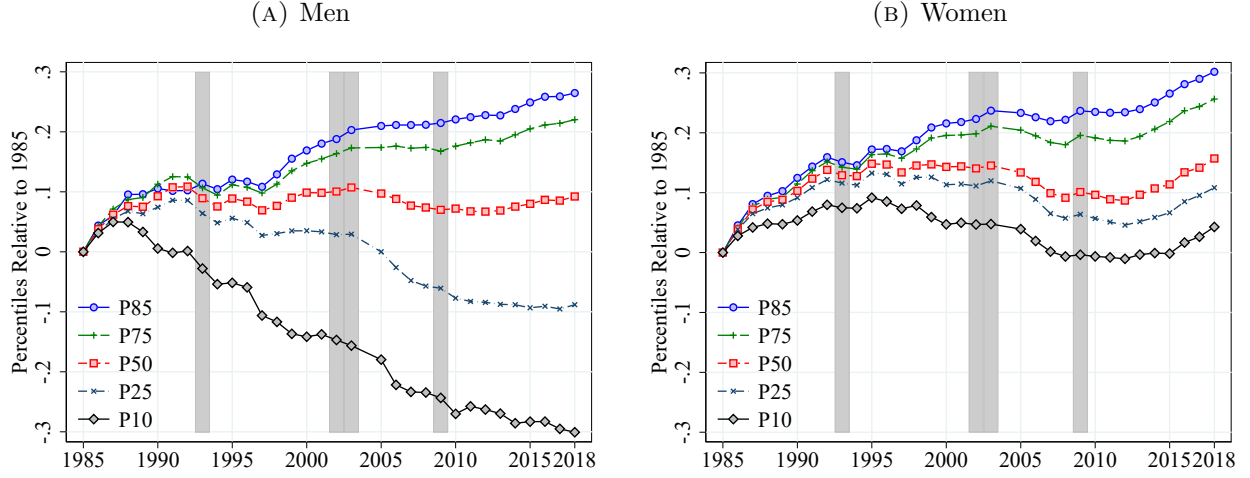


Notes: This figure shows the the P90-P50 and P50-P10 differentials of the distribution of 1-year changes in residualized log earnings (from $t - 1$ to t) in the IAB data (LS sample, truncated as stated below). The P90 for men is above the top-coding threshold and therefore imputed. LS sample with minimum income threshold of 6,250 Euro (2018 prices). The LS sample in the main text uses 2,300 Euro as cutoff to include mini-jobs. Shaded areas indicate recessions. The results for our core sample can be found in Figure 7.

F.2 IAB Data 1985–2018 (West Germany)

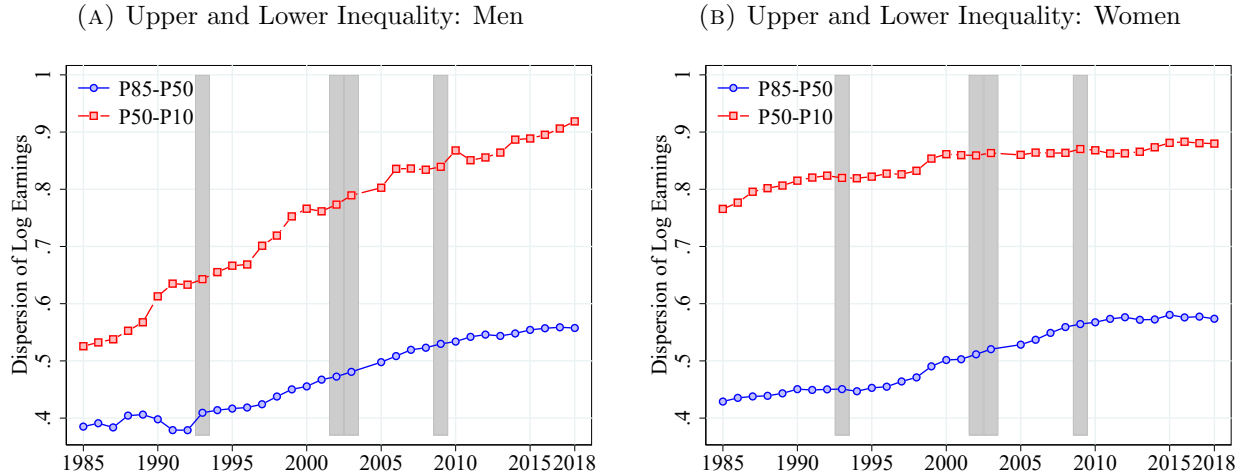
For our longest sample we use the SIAB 1975-2019 (Frodermann et al., 2021) for the years 1985 to 2018. We start in 1985 due to a structural break in the data in 1984. We apply the same minimum earnings threshold of 6,250 in 2018 Euro to exclude mini-jobs from the data as in the time sample 1993-2018. Furthermore, as data for East Germany is available from 1992 onward, we show the earnings development for West Germany only to avoid a structural break in the time series. For men, the wages are imputed from around the P90 upwards, therefore we show the P85 here instead. Moreover, we leave out the year 2004 as there is an unresolved data issue in the SIAB data affecting earnings in this year.

FIGURE F.6: EVOLUTION OF LOG EARNINGS PERCENTILES



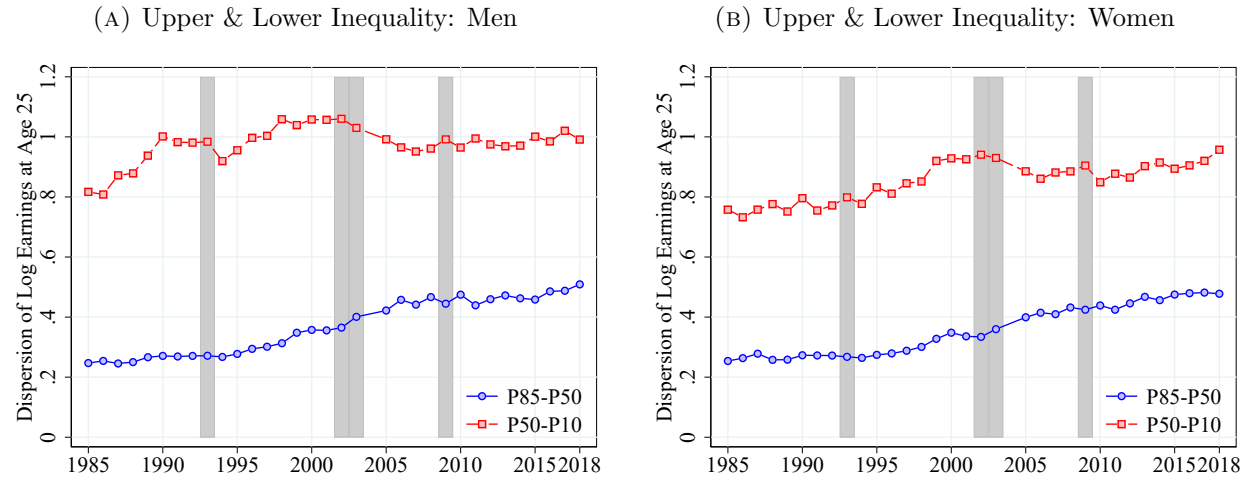
Notes: This figure shows the evolution of selected percentiles of log real earnings from 1985 to 2018 in the IAB data (CS sample, truncated as stated below). The P90 for men is above the top-coding threshold and therefore imputed. All percentiles are normalized to 0 in 1985. CS sample with minimum income threshold of 6,250 Euro (2018 prices). The CS sample in the main text uses 2,300 Euro as cutoff to include mini-jobs. Shaded areas indicate recessions. The results for our core sample can be found in Figure 7.

FIGURE F.7: EARNINGS INEQUALITY: LOG PERCENTILE DIFFERENTIALS



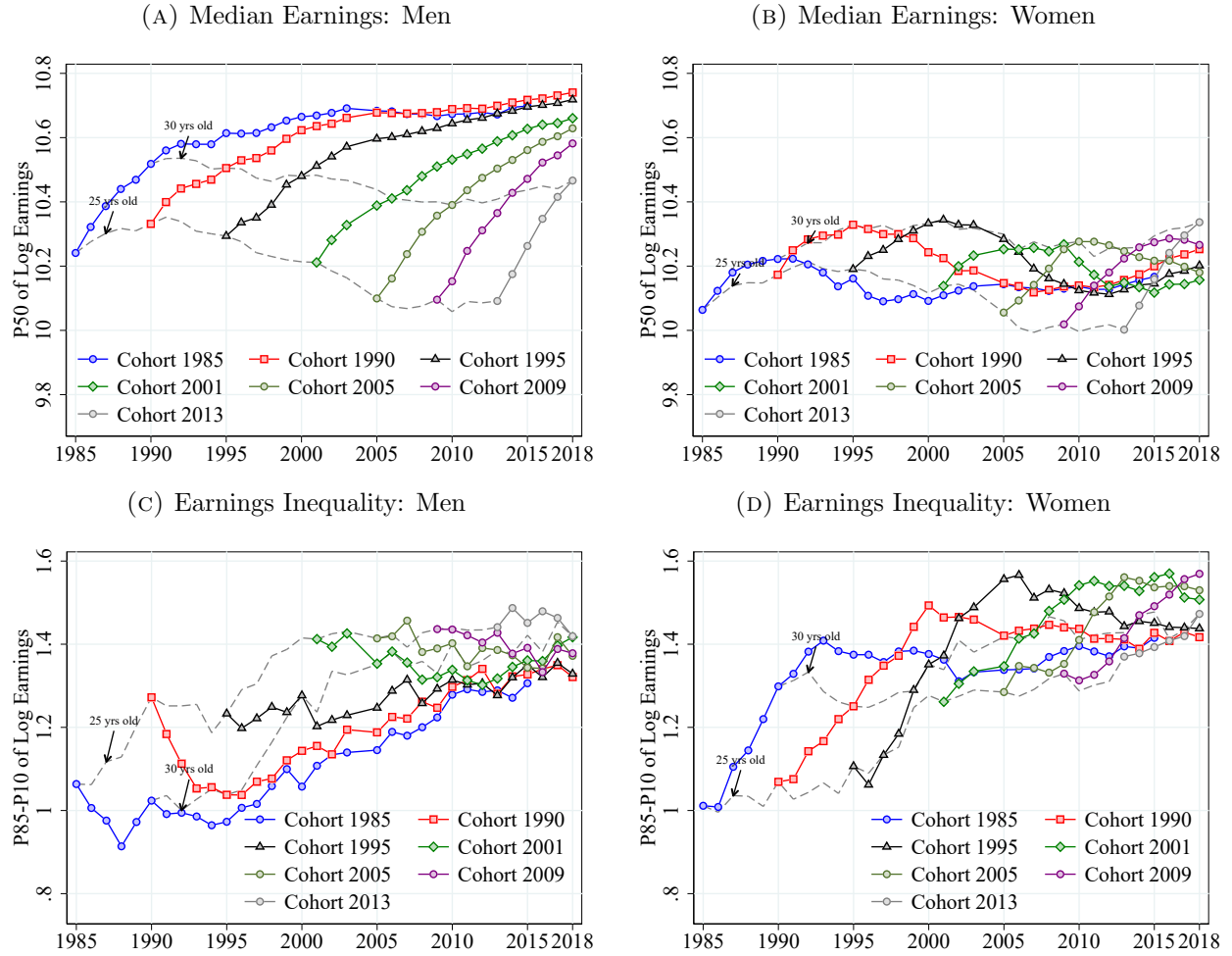
Notes: This figure shows percentile differentials of log real annual earnings in the IAB data (CS sample, truncated as stated below). The P90 for men is above the top-coding threshold and therefore imputed. CS sample with minimum income threshold of 6,250 Euro (2018 prices). The CS sample in the main text uses 2,300 Euro as cutoff to include mini-jobs. Shaded areas indicate recessions. The results for our core sample can be found in Figure 3.

FIGURE F.8: INITIAL INCOME INEQUALITY (AT AGE 25): LOG PERCENTILE DIFFERENTIALS



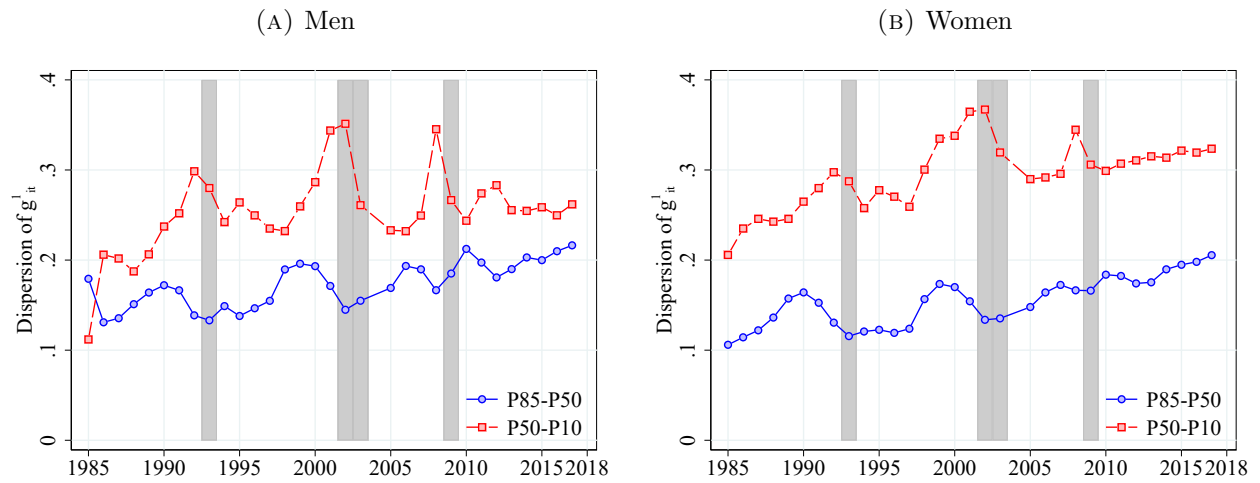
Notes: Shaded areas indicate recessions. CS sample with minimum income threshold of 6,250 Euro (2018 prices). The CS sample in the main text uses 2,300 Euro as cutoff to include mini-jobs. The IAB data is top-coded and imputed above about 60,000 Euro, which is above around the P90 here. The results for our main sample can be found in Figure 4.

FIGURE F.9: EARNINGS PROFILES AND INEQUALITY BY COHORT



Notes: This figure shows the evolution of the median (P50) as well as the P90-P10 differential of the log real annual earnings distribution over time in the IAB data (CS sample, truncated as stated below) separately for men and women. Each colored line corresponds to an individual cohort, where “cohort t ” represents the cohort aged 25 in year t . CS sample with minimum income threshold of 6,250 Euro (2018 prices). The CS sample in the main text uses 2,300 Euro as cutoff to include mini-jobs. Note that the year 2004 is omitted because of unresolved data issues. Shaded areas indicate recessions. Results for our main sample can be found in Figure 6.

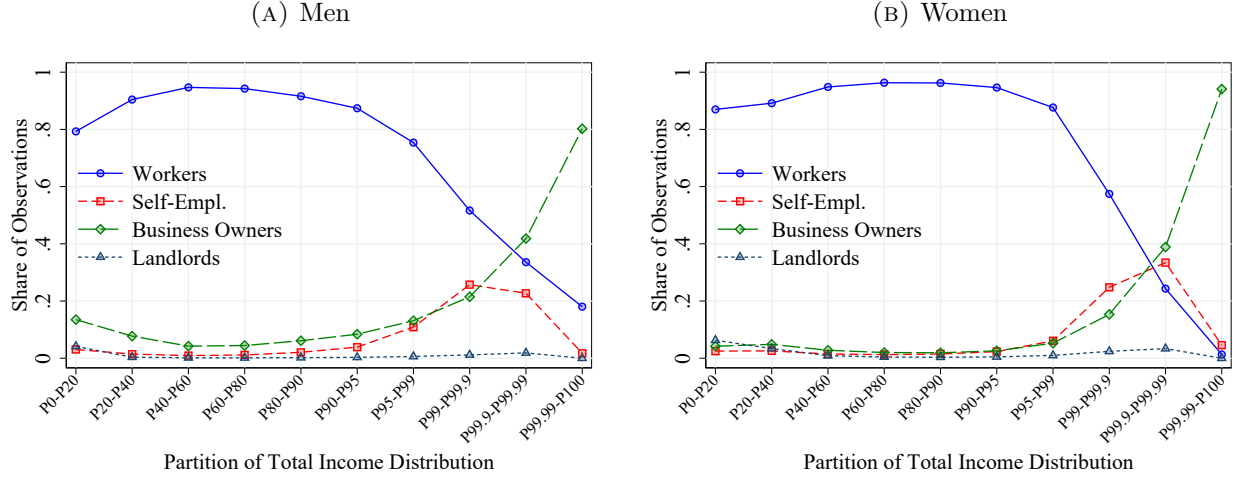
FIGURE F.10: DISPERSION OF 1-YEAR LOG EARNINGS CHANGES



Notes: This figure shows the the P90-P50 and P50-P10 differentials of the distribution of 1-year changes in residualized log earnings (from $t-1$ to t) in the IAB data (LS sample, truncated as stated below). LS sample with minimum income threshold of 6,250 Euro (2018 prices). The LS sample in the main text uses 2,300 Euro as cutoff to include mini-jobs. Shaded areas indicate recessions. The results for our core sample can be found in Figure 7.

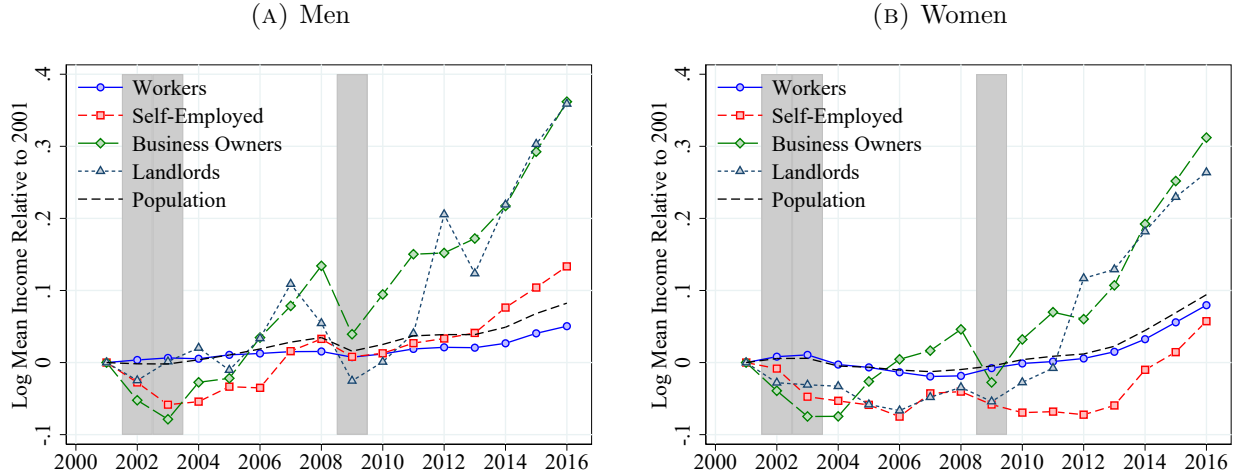
G Specific Analysis: Additional Figures and Tables on Total Income Inequality and Dynamics (Section 4)

FIGURE G.1: MAIN INCOME SOURCES ACROSS THE INCOME DISTRIBUTION



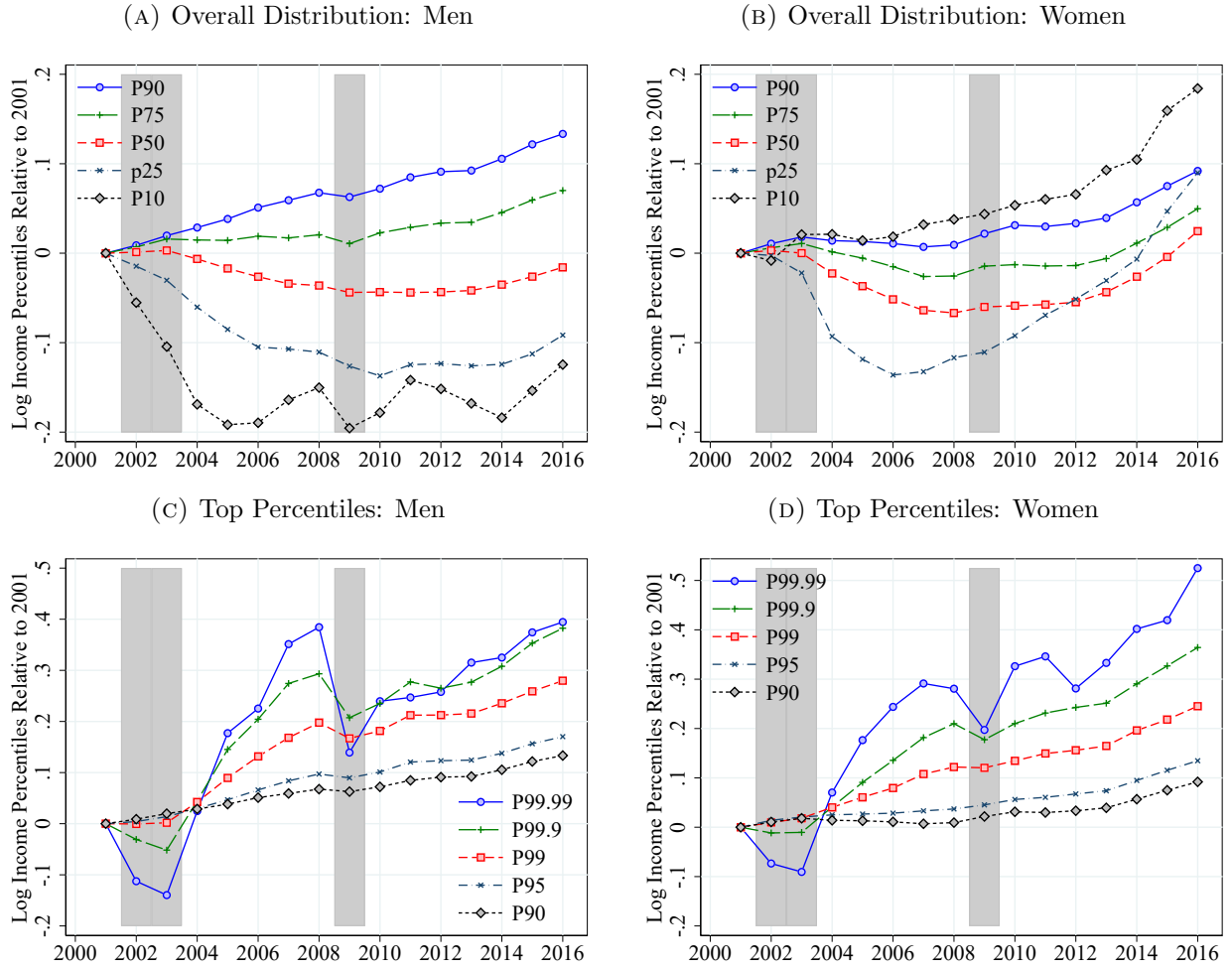
Notes: This figure shows the share of observations with different main income source for different groups of the total income distribution in the combined IAB-TPP data (CS analysis sample). The figure shows averages from 2001 to 2016.

FIGURE G.2: EVOLUTION OF LOG AVERAGE INCOME BY MAIN INCOME SOURCE



Notes: This figure shows the evolution the log of average real annual total income (relative to 2001) in the combined IAB-TPP data (CS analysis sample) by main income source separately for men and women. Shaded areas indicate recessions. See Figure 11 for corresponding levels.

FIGURE G.3: EVOLUTION OF LOG TOTAL INCOME PERCENTILES



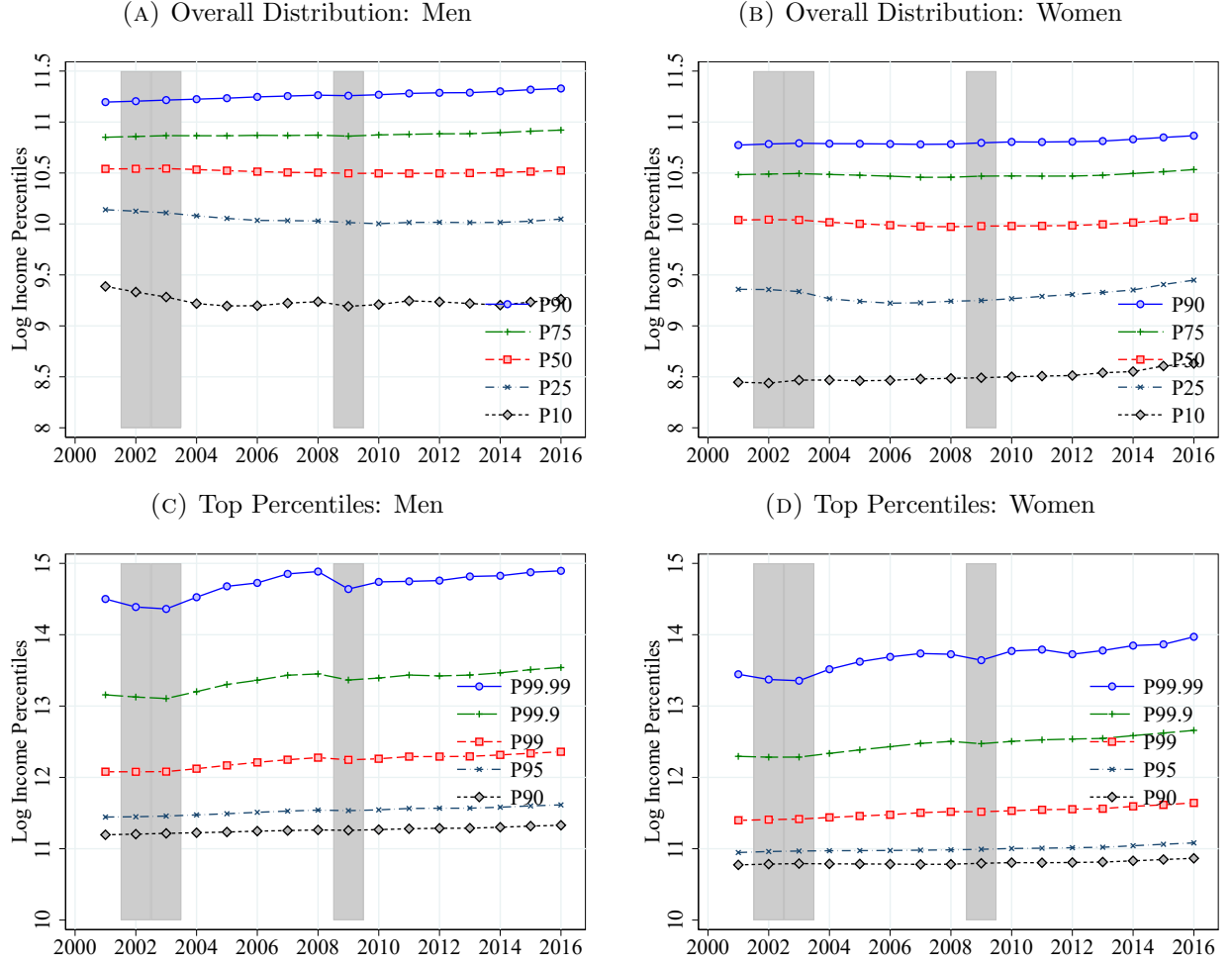
Notes: This figure shows the evolution of selected percentiles of log real annual total income (relative to 2001) in the combined IAB-TPP data (CS analysis sample) separately for men and women. Shaded areas indicate recessions. See Figure 3 for the same analysis of only labor earnings (albeit for a slightly different sample as discussed in the text).

TABLE G.1: PERCENTILES OF REAL ANNUAL TOTAL INCOME (ANALYSIS SAMPLE)

| Year | N | Mean | P5 | P10 | P25 | P50 | P75 | P90 | P95 | P99 | P99.9 | P99.99 |
|-------------------|--------|--------|-------|--------|--------|--------|--------|--------|---------|---------|---------|-----------|
| Men | | | | | | | | | | | | |
| 2001 | 15.373 | 43,989 | 6,654 | 11,937 | 25,329 | 37,808 | 51,579 | 72,826 | 93,353 | 176,187 | 517,722 | 1,984,431 |
| 2002 | 15.127 | 43,912 | 6,275 | 11,295 | 24,964 | 37,853 | 51,937 | 73,469 | 93,746 | 176,093 | 502,011 | 1,773,271 |
| 2003 | 14.866 | 43,903 | 5,889 | 10,754 | 24,574 | 37,924 | 52,403 | 74,273 | 94,649 | 176,508 | 491,658 | 1,725,813 |
| 2004 | 14.741 | 44,157 | 5,598 | 10,081 | 23,848 | 37,564 | 52,348 | 74,945 | 96,180 | 183,802 | 541,297 | 2,033,983 |
| 2005 | 14.565 | 44,453 | 5,523 | 9,854 | 23,259 | 37,169 | 52,323 | 75,668 | 97,818 | 192,677 | 598,642 | 2,369,003 |
| 2006 | 14.621 | 44,826 | 5,558 | 9,875 | 22,806 | 36,824 | 52,567 | 76,633 | 99,720 | 200,946 | 635,102 | 2,485,810 |
| 2007 | 14.758 | 45,253 | 5,674 | 10,132 | 22,757 | 36,540 | 52,468 | 77,266 | 101,524 | 208,477 | 681,134 | 2,819,799 |
| 2008 | 14.768 | 45,542 | 5,650 | 10,272 | 22,678 | 36,465 | 52,652 | 77,918 | 102,878 | 214,685 | 694,181 | 2,914,545 |
| 2009 | 14.498 | 44,690 | 5,527 | 9,817 | 22,324 | 36,184 | 52,133 | 77,544 | 102,111 | 208,188 | 637,002 | 2,280,813 |
| 2010 | 14.630 | 45,107 | 5,565 | 9,988 | 22,084 | 36,203 | 52,765 | 78,271 | 103,285 | 211,273 | 655,038 | 2,521,691 |
| 2011 | 14.796 | 45,650 | 5,697 | 10,358 | 22,360 | 36,182 | 53,091 | 79,270 | 105,317 | 217,863 | 683,204 | 2,540,327 |
| 2012 | 14.854 | 45,730 | 5,574 | 10,257 | 22,388 | 36,202 | 53,345 | 79,771 | 105,600 | 217,914 | 674,841 | 2,568,046 |
| 2013 | 14.892 | 45,729 | 5,550 | 10,092 | 22,332 | 36,265 | 53,393 | 79,874 | 105,720 | 218,549 | 682,847 | 2,719,594 |
| 2014 | 14.974 | 46,199 | 5,523 | 9,932 | 22,371 | 36,501 | 53,977 | 80,933 | 107,136 | 223,002 | 704,432 | 2,746,560 |
| 2015 | 15.054 | 47,085 | 5,663 | 10,237 | 22,632 | 36,829 | 54,736 | 82,248 | 109,164 | 228,277 | 736,971 | 2,885,246 |
| 2016 | 15.079 | 47,768 | 5,783 | 10,540 | 23,111 | 37,216 | 55,323 | 83,213 | 110,686 | 233,049 | 758,905 | 2,944,148 |
| Women | | | | | | | | | | | | |
| 2001 | 12.558 | 26,126 | 3,704 | 4,662 | 11,602 | 22,908 | 35,732 | 47,757 | 56,790 | 89,124 | 218,577 | 690,697 |
| 2002 | 12.531 | 26,274 | 3,685 | 4,624 | 11,573 | 22,980 | 35,949 | 48,267 | 57,596 | 89,996 | 216,063 | 641,789 |
| 2003 | 12.363 | 26,280 | 3,652 | 4,760 | 11,347 | 22,910 | 36,120 | 48,631 | 57,970 | 90,759 | 216,306 | 630,920 |
| 2004 | 12.345 | 26,001 | 3,628 | 4,761 | 10,571 | 22,396 | 35,786 | 48,433 | 58,234 | 92,789 | 227,721 | 741,022 |
| 2005 | 12.294 | 25,957 | 3,600 | 4,729 | 10,304 | 22,081 | 35,533 | 48,386 | 58,325 | 94,709 | 239,373 | 823,981 |
| 2006 | 12.330 | 25,855 | 3,597 | 4,749 | 10,125 | 21,756 | 35,200 | 48,276 | 58,432 | 96,490 | 250,291 | 881,431 |
| 2007 | 12.486 | 25,799 | 3,691 | 4,813 | 10,163 | 21,492 | 34,812 | 48,094 | 58,703 | 99,261 | 262,099 | 924,239 |
| 2008 | 12.522 | 25,876 | 3,723 | 4,841 | 10,322 | 21,429 | 34,832 | 48,205 | 58,929 | 100,686 | 269,619 | 914,497 |
| 2009 | 12.544 | 26,000 | 3,734 | 4,870 | 10,386 | 21,573 | 35,217 | 48,808 | 59,419 | 100,528 | 261,006 | 841,224 |
| 2010 | 12.644 | 26,232 | 3,775 | 4,918 | 10,580 | 21,601 | 35,279 | 49,274 | 60,077 | 101,969 | 269,650 | 957,126 |
| 2011 | 12.774 | 26,355 | 3,814 | 4,951 | 10,824 | 21,629 | 35,224 | 49,202 | 60,336 | 103,478 | 275,482 | 976,390 |
| 2012 | 12.905 | 26,440 | 3,810 | 4,979 | 11,020 | 21,688 | 35,242 | 49,378 | 60,753 | 104,167 | 278,631 | 914,946 |
| 2013 | 12.962 | 26,723 | 3,891 | 5,115 | 11,253 | 21,933 | 35,518 | 49,673 | 61,141 | 105,084 | 281,024 | 963,589 |
| 2014 | 13.028 | 27,315 | 3,964 | 5,176 | 11,525 | 22,313 | 36,135 | 50,550 | 62,439 | 108,432 | 292,271 | 1,032,300 |
| 2015 | 13.092 | 27,996 | 4,093 | 5,467 | 12,158 | 22,811 | 36,777 | 51,477 | 63,736 | 110,838 | 303,084 | 1,050,728 |
| 2016 | 13.079 | 28,707 | 4,179 | 5,604 | 12,693 | 23,477 | 37,553 | 52,357 | 64,982 | 113,904 | 314,598 | 1,167,646 |
| Population | | | | | | | | | | | | |
| 2001 | 27.930 | 35,958 | 4,353 | 6,679 | 17,265 | 31,617 | 44,775 | 62,189 | 79,289 | 144,373 | 409,211 | 1,504,645 |
| 2002 | 27.658 | 35,921 | 4,295 | 6,483 | 17,004 | 31,554 | 45,005 | 62,730 | 79,860 | 144,218 | 397,504 | 1,374,766 |
| 2003 | 27.230 | 35,902 | 4,279 | 6,114 | 16,735 | 31,500 | 45,295 | 63,238 | 80,655 | 144,970 | 391,890 | 1,327,219 |
| 2004 | 27.086 | 35,882 | 4,214 | 5,838 | 16,057 | 30,976 | 45,033 | 63,502 | 81,621 | 149,589 | 424,043 | 1,530,896 |
| 2005 | 26.859 | 35,987 | 4,187 | 5,775 | 15,713 | 30,490 | 44,868 | 63,751 | 82,548 | 155,180 | 464,490 | 1,771,740 |
| 2006 | 26.951 | 36,147 | 4,206 | 5,763 | 15,491 | 30,060 | 44,785 | 64,255 | 83,873 | 160,990 | 492,269 | 1,935,221 |
| 2007 | 27.244 | 36,338 | 4,280 | 5,799 | 15,450 | 29,727 | 44,556 | 64,505 | 84,916 | 166,350 | 520,798 | 2,096,380 |
| 2008 | 27.291 | 36,518 | 4,326 | 5,849 | 15,483 | 29,599 | 44,579 | 64,881 | 85,714 | 170,816 | 534,067 | 2,142,705 |
| 2009 | 27.042 | 36,020 | 4,299 | 5,806 | 15,274 | 29,446 | 44,497 | 64,623 | 85,083 | 166,595 | 496,910 | 1,745,501 |
| 2010 | 27.274 | 36,357 | 4,335 | 5,901 | 15,325 | 29,369 | 44,840 | 65,310 | 86,026 | 169,209 | 510,570 | 1,892,326 |
| 2011 | 27.570 | 36,710 | 4,360 | 6,085 | 15,636 | 29,353 | 44,896 | 65,835 | 87,229 | 173,786 | 530,741 | 1,966,156 |
| 2012 | 27.760 | 36,762 | 4,346 | 6,154 | 15,641 | 29,342 | 44,938 | 66,273 | 87,759 | 173,904 | 522,565 | 1,887,114 |
| 2013 | 27.854 | 36,884 | 4,421 | 6,218 | 15,735 | 29,419 | 45,109 | 66,442 | 87,958 | 174,208 | 527,232 | 1,967,935 |
| 2014 | 28.002 | 37,413 | 4,450 | 6,287 | 15,897 | 29,693 | 45,663 | 67,384 | 89,260 | 177,867 | 542,424 | 2,047,636 |
| 2015 | 28.146 | 38,205 | 4,602 | 6,718 | 16,476 | 30,070 | 46,318 | 68,547 | 90,879 | 182,272 | 564,562 | 2,194,591 |
| 2016 | 28.158 | 38,914 | 4,719 | 7,120 | 17,013 | 30,659 | 46,947 | 69,538 | 92,185 | 185,773 | 582,065 | 2,291,329 |

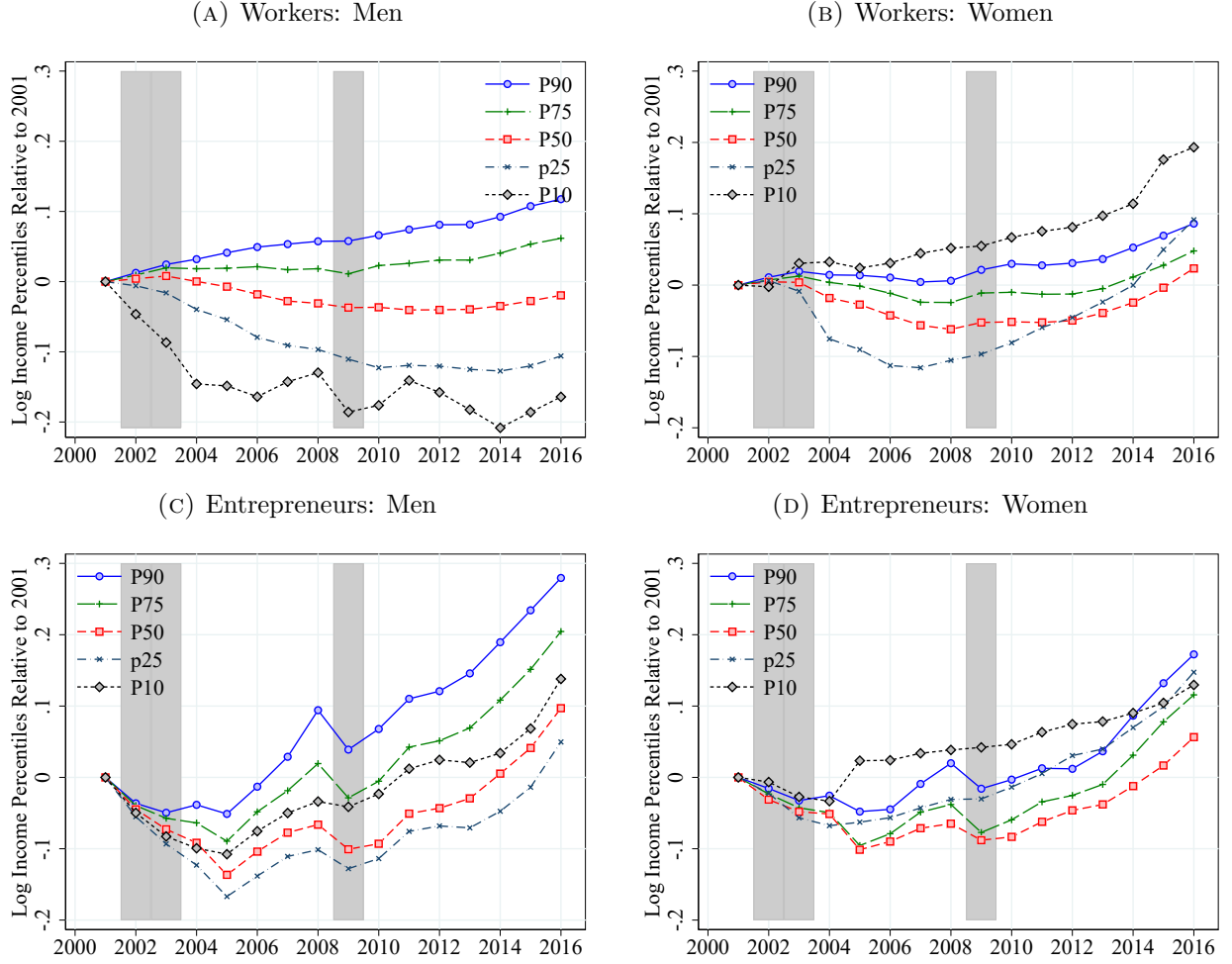
Notes: This table shows the number of observations (in millions) and selected percentiles of real annual total income (in 2018 Euro) in the combined IAB-TPP data (CS analysis sample) separately for men and women and in the population. See Table 1 for the percentiles of labor earnings (albeit for a slightly different sample, as discussed in the text).

FIGURE G.4: PERCENTILES OF LOG INCOME



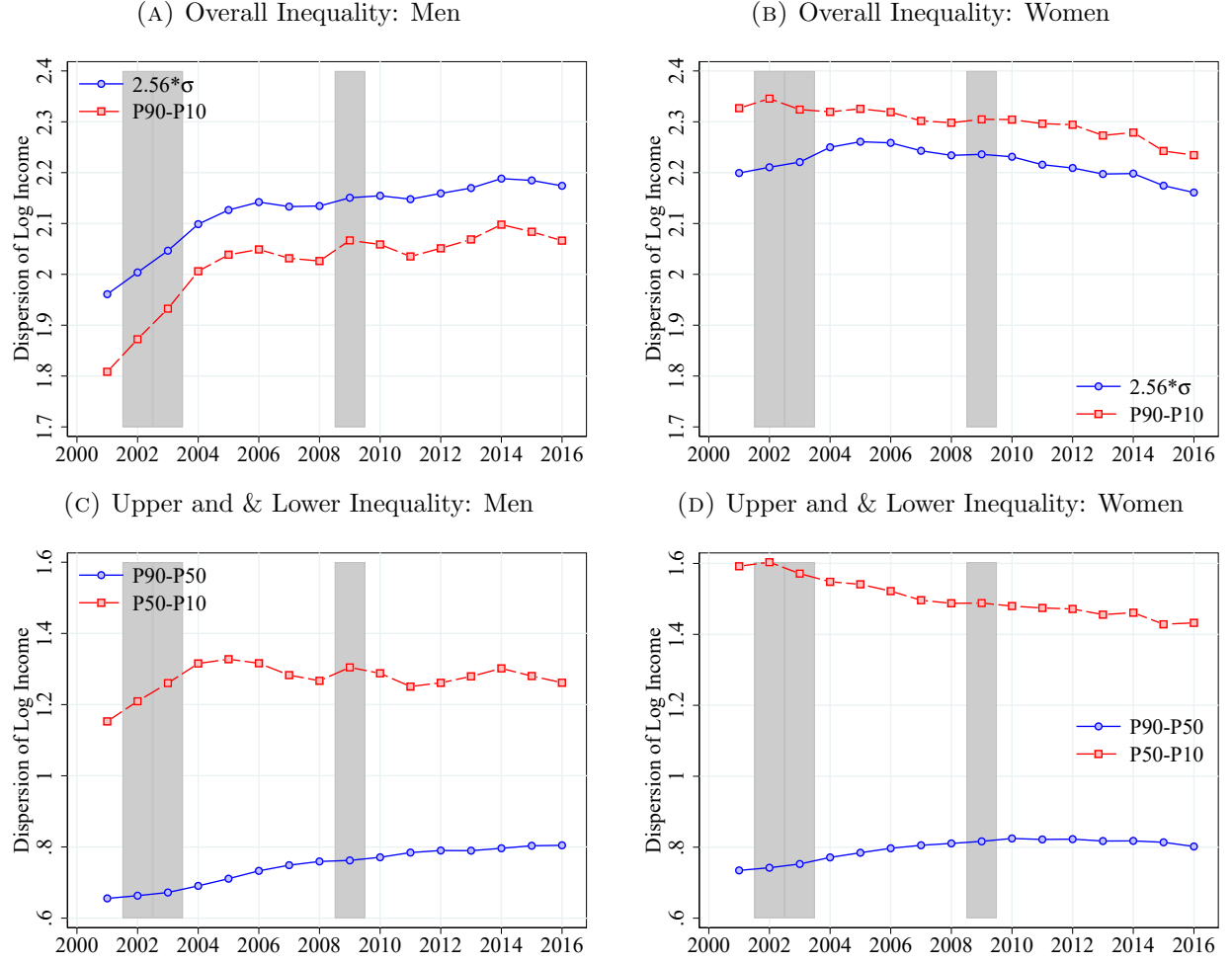
Notes: This figure shows the evolution of absolute log real annual total income percentiles in the combined IAB-TPP data (CS sample) separately for men and women. Shaded areas indicate recessions.

FIGURE G.5: EVOLUTION OF LOG INCOME PERCENTILES BY MAIN INCOME SOURCE



Notes: This figure shows the evolution of different percentiles of log total income among workers and entrepreneurs in the combined IAB-TTP data (CS sample). Workers receive at least half of their income from labor earnings. The jump in the P10 for entrepreneurs (while it is obvious for women, it is hidden for men) is related to a similar jump in the number of observations classified as landlords from 2004 to 2005 which is plausibly related to a reform in the taxation of pensions. In line with this, the jump is entirely driven by landlords (as opposed to self-employed or business owners). Shaded areas indicate recessions.

FIGURE G.6: DISPERSION OF LOG REAL INCOME DISTRIBUTION



Notes: This figure shows the evolution of different log percentile differentials as well as the (rescaled) standard deviation of the log real annual total income distribution over time in the combined IAB-TPP data (CS sample) separately for men and women. The standard deviation σ is rescaled as $2.56 * \sigma$ corresponds to P90-P10 differential for a Gaussian distribution. Shaded areas indicate recessions.

TABLE G.2: INCOME SHARES – MEN

| Year | Q1 | Q2 | Q3 | Q4 | Q5 | Bot 50 | Bot 90 | Mid 40 | Top 10 | Top 5 | Top 1 | Top 0.1 | Top 0.01 |
|------|------|-------|-------|-------|-------|--------|--------|--------|--------|-------|-------|---------|----------|
| 2001 | 5.44 | 12.83 | 17.20 | 21.98 | 42.57 | 26.38 | 71.82 | 45.44 | 28.18 | 18.91 | 8.11 | 2.85 | 1.14 |
| 2002 | 5.21 | 12.73 | 17.25 | 22.15 | 42.66 | 26.07 | 71.88 | 45.81 | 28.12 | 18.77 | 7.94 | 2.73 | 1.11 |
| 2003 | 5.01 | 12.59 | 17.29 | 22.33 | 42.77 | 25.74 | 71.91 | 46.17 | 28.09 | 18.64 | 7.72 | 2.55 | 0.98 |
| 2004 | 4.74 | 12.23 | 17.02 | 22.14 | 43.87 | 24.95 | 70.80 | 45.85 | 29.20 | 19.69 | 8.56 | 3.10 | 1.37 |
| 2005 | 4.60 | 11.89 | 16.73 | 21.95 | 44.82 | 24.32 | 69.84 | 45.52 | 30.16 | 20.59 | 9.22 | 3.43 | 1.46 |
| 2006 | 4.53 | 11.58 | 16.45 | 21.82 | 45.62 | 23.78 | 69.04 | 45.26 | 30.96 | 21.31 | 9.71 | 3.66 | 1.55 |
| 2007 | 4.58 | 11.40 | 16.18 | 21.55 | 46.30 | 23.51 | 68.28 | 44.77 | 31.72 | 22.04 | 10.23 | 3.96 | 1.70 |
| 2008 | 4.58 | 11.29 | 16.03 | 21.45 | 46.65 | 23.33 | 67.93 | 44.60 | 32.07 | 22.35 | 10.36 | 3.93 | 1.65 |
| 2009 | 4.51 | 11.35 | 16.21 | 21.68 | 46.25 | 23.40 | 68.53 | 45.14 | 31.47 | 21.62 | 9.60 | 3.37 | 1.36 |
| 2010 | 4.49 | 11.14 | 16.08 | 21.70 | 46.59 | 23.09 | 68.21 | 45.12 | 31.79 | 21.94 | 9.88 | 3.58 | 1.48 |
| 2011 | 4.57 | 11.09 | 15.88 | 21.54 | 46.93 | 23.02 | 67.84 | 44.81 | 32.16 | 22.28 | 10.05 | 3.61 | 1.46 |
| 2012 | 4.53 | 11.07 | 15.86 | 21.57 | 46.96 | 22.96 | 67.87 | 44.91 | 32.13 | 22.21 | 9.98 | 3.60 | 1.49 |
| 2013 | 4.48 | 11.05 | 15.88 | 21.62 | 46.97 | 22.88 | 67.88 | 45.00 | 32.12 | 22.19 | 9.94 | 3.51 | 1.35 |
| 2014 | 4.40 | 10.96 | 15.82 | 21.61 | 47.20 | 22.68 | 67.68 | 45.00 | 32.32 | 22.36 | 10.05 | 3.54 | 1.34 |
| 2015 | 4.44 | 10.86 | 15.67 | 21.49 | 47.54 | 22.54 | 67.30 | 44.76 | 32.70 | 22.76 | 10.44 | 3.84 | 1.54 |
| 2016 | 4.50 | 10.89 | 15.62 | 21.41 | 47.58 | 22.61 | 67.22 | 44.61 | 32.78 | 22.86 | 10.52 | 3.88 | 1.55 |

Notes: This table shows the share of (total) income that goes to selected parts of the income distribution of men in the combined IAB-TPP data (CS sample). Q1 to Q5 refer to the five quintiles where Q1 (Q5) stands for the bottom (top) 20% of the income distribution. The quintile shares sum to one. Bot 50, Bot 90 and Mid 40 refer to observations in the bottom 50%, the bottom 90% and between the median and the 90th percentile of the income distribution. Top x refers to the top $x\%$ of the income distribution.

TABLE G.3: INCOME SHARES – WOMEN

| Year | Q1 | Q2 | Q3 | Q4 | Q5 | Bot 50 | Bot 90 | Mid 40 | Top 10 | Top 5 | Top 1 | Top 0.1 | Top 0.01 |
|------|------|-------|-------|-------|-------|--------|--------|--------|--------|-------|-------|---------|----------|
| 2001 | 3.80 | 10.66 | 17.61 | 25.35 | 42.58 | 22.38 | 73.79 | 51.42 | 26.21 | 16.31 | 6.06 | 1.94 | 0.83 |
| 2002 | 3.77 | 10.62 | 17.56 | 25.35 | 42.71 | 22.27 | 73.72 | 51.45 | 26.28 | 16.32 | 5.99 | 1.88 | 0.82 |
| 2003 | 3.73 | 10.48 | 17.51 | 25.43 | 42.85 | 22.07 | 73.69 | 51.62 | 26.31 | 16.28 | 5.89 | 1.74 | 0.68 |
| 2004 | 3.63 | 10.07 | 17.30 | 25.39 | 43.61 | 21.44 | 73.01 | 51.56 | 26.99 | 16.86 | 6.22 | 1.89 | 0.73 |
| 2005 | 3.59 | 9.88 | 17.09 | 25.21 | 44.22 | 21.12 | 72.38 | 51.25 | 27.62 | 17.47 | 6.73 | 2.26 | 0.99 |
| 2006 | 3.62 | 9.74 | 16.91 | 25.04 | 44.69 | 20.92 | 71.88 | 50.96 | 28.12 | 17.93 | 7.05 | 2.40 | 1.08 |
| 2007 | 3.69 | 9.72 | 16.74 | 24.81 | 45.05 | 20.88 | 71.45 | 50.56 | 28.55 | 18.35 | 7.30 | 2.47 | 1.06 |
| 2008 | 3.72 | 9.75 | 16.63 | 24.70 | 45.19 | 20.92 | 71.30 | 50.38 | 28.70 | 18.49 | 7.38 | 2.46 | 1.03 |
| 2009 | 3.73 | 9.76 | 16.65 | 24.81 | 45.05 | 20.93 | 71.56 | 50.63 | 28.44 | 18.18 | 7.07 | 2.22 | 0.88 |
| 2010 | 3.75 | 9.77 | 16.54 | 24.64 | 45.30 | 20.93 | 71.27 | 50.34 | 28.73 | 18.44 | 7.29 | 2.39 | 0.97 |
| 2011 | 3.82 | 9.89 | 16.48 | 24.47 | 45.34 | 21.09 | 71.12 | 50.03 | 28.88 | 18.63 | 7.44 | 2.50 | 1.05 |
| 2012 | 3.86 | 9.97 | 16.47 | 24.40 | 45.30 | 21.22 | 71.13 | 49.91 | 28.87 | 18.60 | 7.35 | 2.38 | 0.97 |
| 2013 | 3.89 | 10.05 | 16.48 | 24.35 | 45.23 | 21.33 | 71.14 | 49.81 | 28.86 | 18.65 | 7.43 | 2.46 | 1.03 |
| 2014 | 3.91 | 10.04 | 16.40 | 24.22 | 45.43 | 21.31 | 70.85 | 49.54 | 29.15 | 18.97 | 7.73 | 2.68 | 1.23 |
| 2015 | 4.02 | 10.24 | 16.38 | 24.07 | 45.29 | 21.64 | 70.88 | 49.24 | 29.12 | 18.99 | 7.77 | 2.70 | 1.22 |
| 2016 | 4.10 | 10.37 | 16.43 | 24.00 | 45.10 | 21.88 | 70.96 | 49.08 | 29.04 | 18.98 | 7.81 | 2.74 | 1.21 |

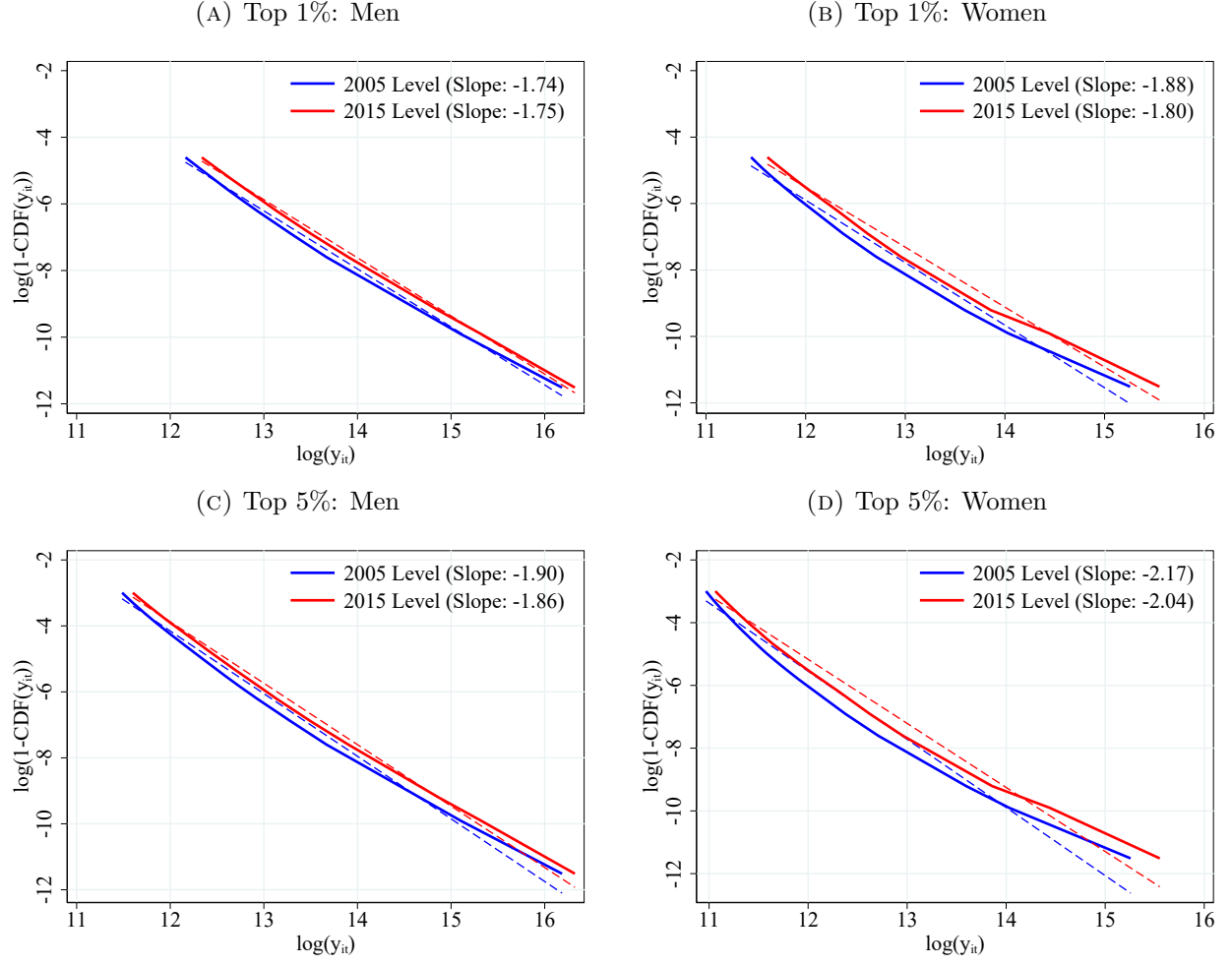
Notes: This table shows the share of (total) income that goes to selected parts of the income distribution of women in the combined IAB-TPP data (CS sample). Q1 to Q5 refer to the five quintiles where Q1 (Q5) stands for the bottom (top) 20% of the income distribution. The quintile shares sum to one. Bot 50, Bot 90 and Mid 40 refer to observations in the bottom 50%, the bottom 90% and between the median and the 90th percentile of the income distribution. Top x refers to the top $x\%$ of the income distribution.

TABLE G.4: INCOME SHARES – POPULATION

| Year | Q1 | Q2 | Q3 | Q4 | Q5 | Bot 50 | Bot 90 | Mid 40 | Top 10 | Top 5 | Top 1 | Top 0.1 | Top 0.01 |
|------|------|-------|-------|-------|-------|--------|--------|--------|--------|-------|-------|---------|----------|
| 2001 | 4.10 | 11.27 | 17.53 | 23.32 | 43.79 | 23.43 | 71.36 | 47.93 | 28.64 | 18.98 | 7.91 | 2.72 | 1.10 |
| 2002 | 4.01 | 11.14 | 17.52 | 23.44 | 43.90 | 23.19 | 71.39 | 48.20 | 28.61 | 18.87 | 7.75 | 2.60 | 1.07 |
| 2003 | 3.89 | 11.01 | 17.49 | 23.57 | 44.04 | 22.91 | 71.37 | 48.47 | 28.63 | 18.79 | 7.57 | 2.43 | 0.94 |
| 2004 | 3.68 | 10.65 | 17.19 | 23.42 | 45.06 | 22.16 | 70.37 | 48.20 | 29.63 | 19.71 | 8.27 | 2.87 | 1.22 |
| 2005 | 3.59 | 10.40 | 16.89 | 23.21 | 45.91 | 21.67 | 69.48 | 47.81 | 30.52 | 20.56 | 8.92 | 3.24 | 1.38 |
| 2006 | 3.56 | 10.19 | 16.59 | 23.00 | 46.66 | 21.28 | 68.72 | 47.45 | 31.28 | 21.24 | 9.39 | 3.46 | 1.47 |
| 2007 | 3.60 | 10.08 | 16.32 | 22.74 | 47.26 | 21.09 | 68.04 | 46.95 | 31.96 | 21.90 | 9.84 | 3.70 | 1.59 |
| 2008 | 3.63 | 10.03 | 16.17 | 22.62 | 47.54 | 20.99 | 67.74 | 46.75 | 32.26 | 22.17 | 9.98 | 3.69 | 1.55 |
| 2009 | 3.63 | 10.06 | 16.32 | 22.89 | 47.10 | 21.08 | 68.31 | 47.23 | 31.69 | 21.52 | 9.32 | 3.20 | 1.27 |
| 2010 | 3.66 | 9.96 | 16.14 | 22.80 | 47.45 | 20.91 | 68.00 | 47.09 | 32.00 | 21.82 | 9.57 | 3.39 | 1.37 |
| 2011 | 3.72 | 9.99 | 15.98 | 22.58 | 47.73 | 20.94 | 67.66 | 46.72 | 32.34 | 22.15 | 9.78 | 3.46 | 1.39 |
| 2012 | 3.73 | 9.99 | 15.96 | 22.56 | 47.76 | 20.94 | 67.68 | 46.74 | 32.32 | 22.08 | 9.69 | 3.40 | 1.39 |
| 2013 | 3.75 | 10.00 | 15.95 | 22.58 | 47.71 | 20.98 | 67.71 | 46.73 | 32.29 | 22.05 | 9.67 | 3.37 | 1.32 |
| 2014 | 3.74 | 9.97 | 15.88 | 22.52 | 47.89 | 20.90 | 67.52 | 46.62 | 32.48 | 22.23 | 9.82 | 3.46 | 1.37 |
| 2015 | 3.83 | 10.03 | 15.75 | 22.36 | 48.03 | 20.99 | 67.31 | 46.32 | 32.69 | 22.48 | 10.09 | 3.67 | 1.50 |
| 2016 | 3.91 | 10.12 | 15.76 | 22.26 | 47.94 | 21.19 | 67.33 | 46.13 | 32.67 | 22.51 | 10.14 | 3.70 | 1.50 |

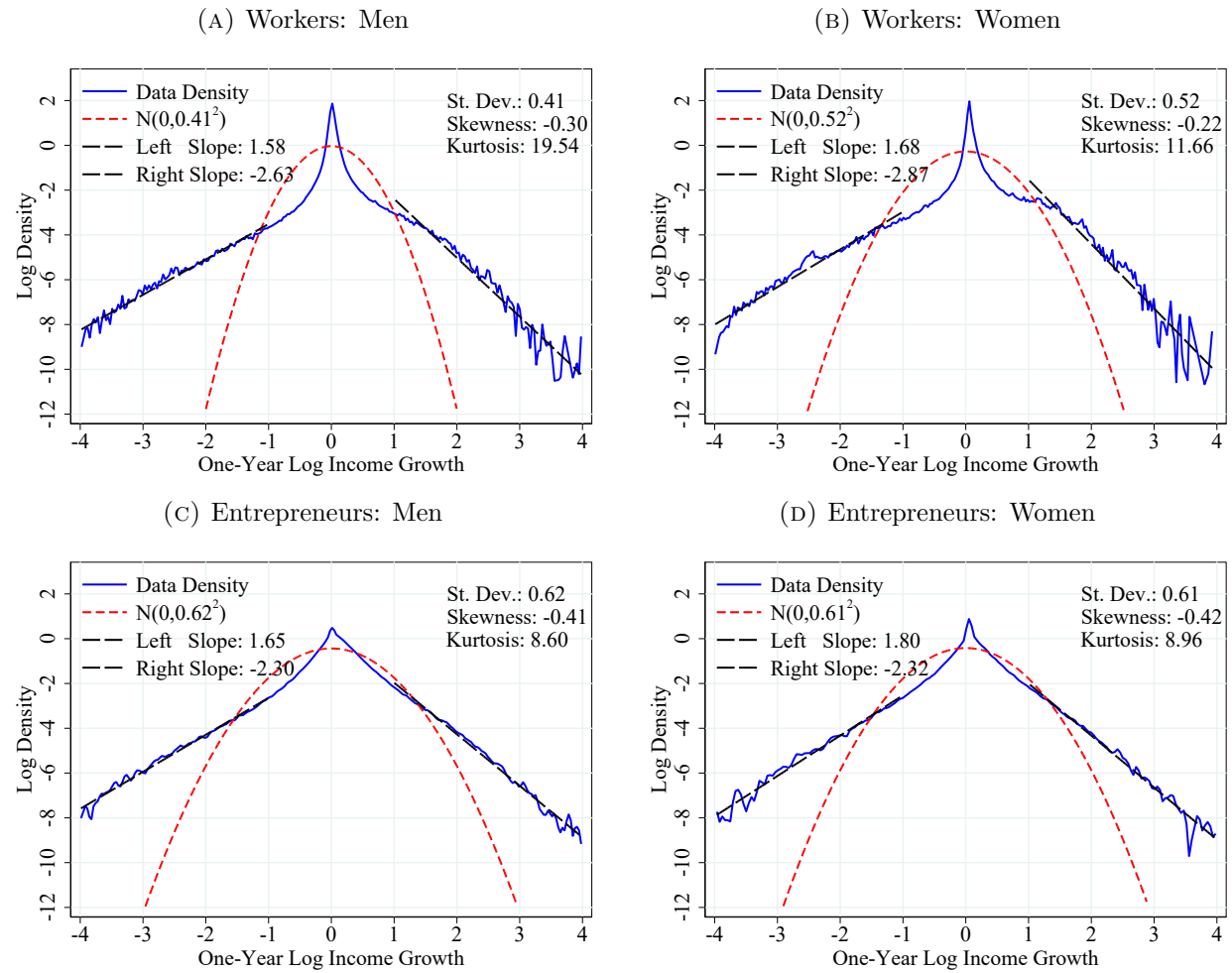
Notes: This table shows the share of (total) income that goes to selected parts of the income distribution in the combined IAB-TPP data (CS sample). Q1 to Q5 refer to the five quintiles where Q1 (Q5) stands for the bottom (top) 20% of the income distribution. The quintile shares sum to one. Bot 50, Bot 90 and Mid 40 refer to observations in the bottom 50%, the bottom 90% and between the median and the 90th percentile of the income distribution. Top x refers to the top $x\%$ of the income distribution.

FIGURE G.7: TOP INCOME INEQUALITY: PARETO TAIL AT TOP 1% AND TOP 5%



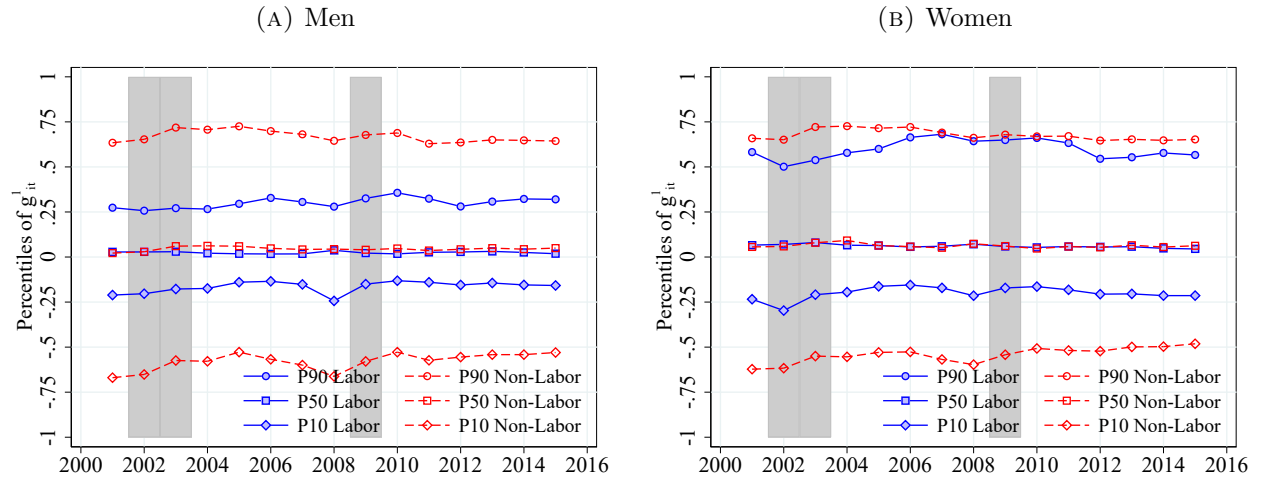
Notes: This figure shows the log of the inverse empirical CDF of log total income and a fitted linear regression line for observations with income in the top 1% and top 5% in the combined IAB-TPP data (CS sample). The absolute value of the slope of the regression line is the Pareto parameter above the respective cutoff.

FIGURE G.8: LOG DENSITY OF 1-YEAR INCOME GROWTH BY MAIN INCOME SOURCE (YEAR 2005)



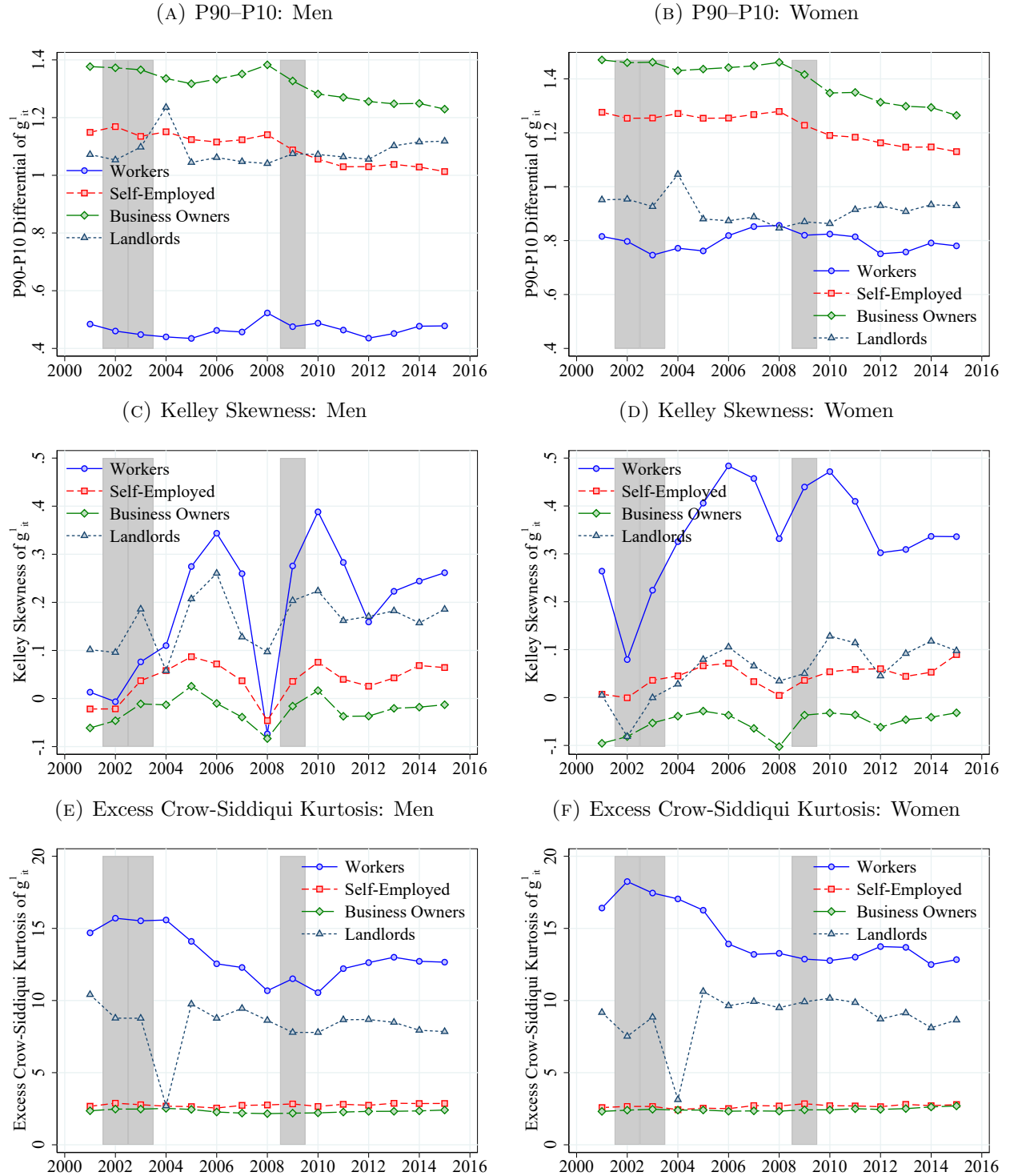
Notes: This figure shows the log density of 1-year changes of residualized log total income separately for workers (labor income as main income source) and entrepreneurs (non-labor income as main income source) and for men and women in the year 2005. LS sample of the combined IAB-TPP data. The dashed line corresponds to the log density of a Normal distribution with the same variance.

FIGURE G.9: PERCENTILES OF 1-YEAR INCOME GROWTH BY MAIN INCOME SOURCE



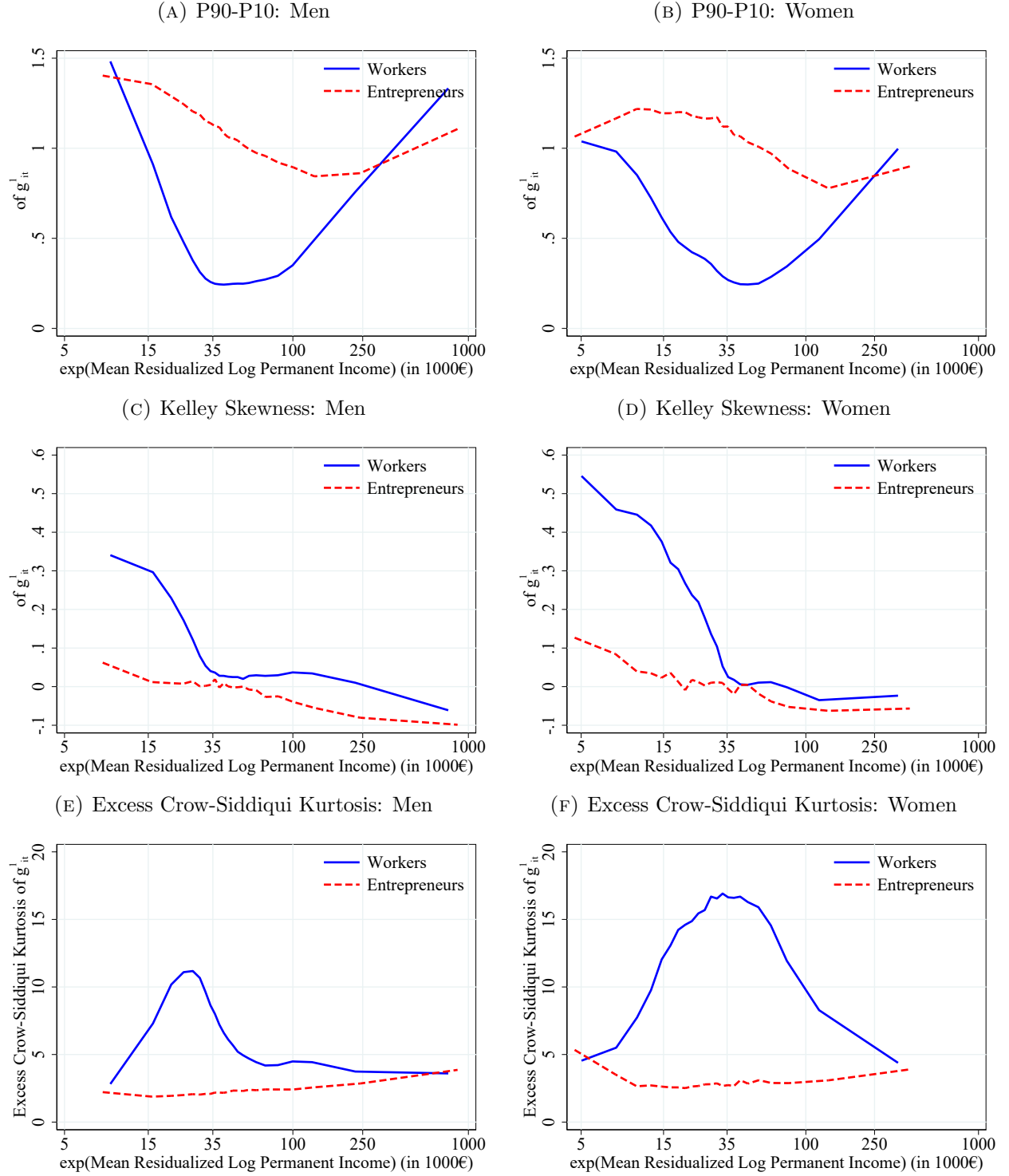
Notes: This figure the 90th, 50th and 10th percentiles of the distribution of 1-year changes in residualized log income (from $t - 1$ to t) by main income source (workers vs. entrepreneurs) using the combined IAB-TPP data (LS sample).

FIGURE G.10: DISPERSION, SKEWNESS AND KURTOSIS OF 1-YEAR LOG INCOME CHANGES



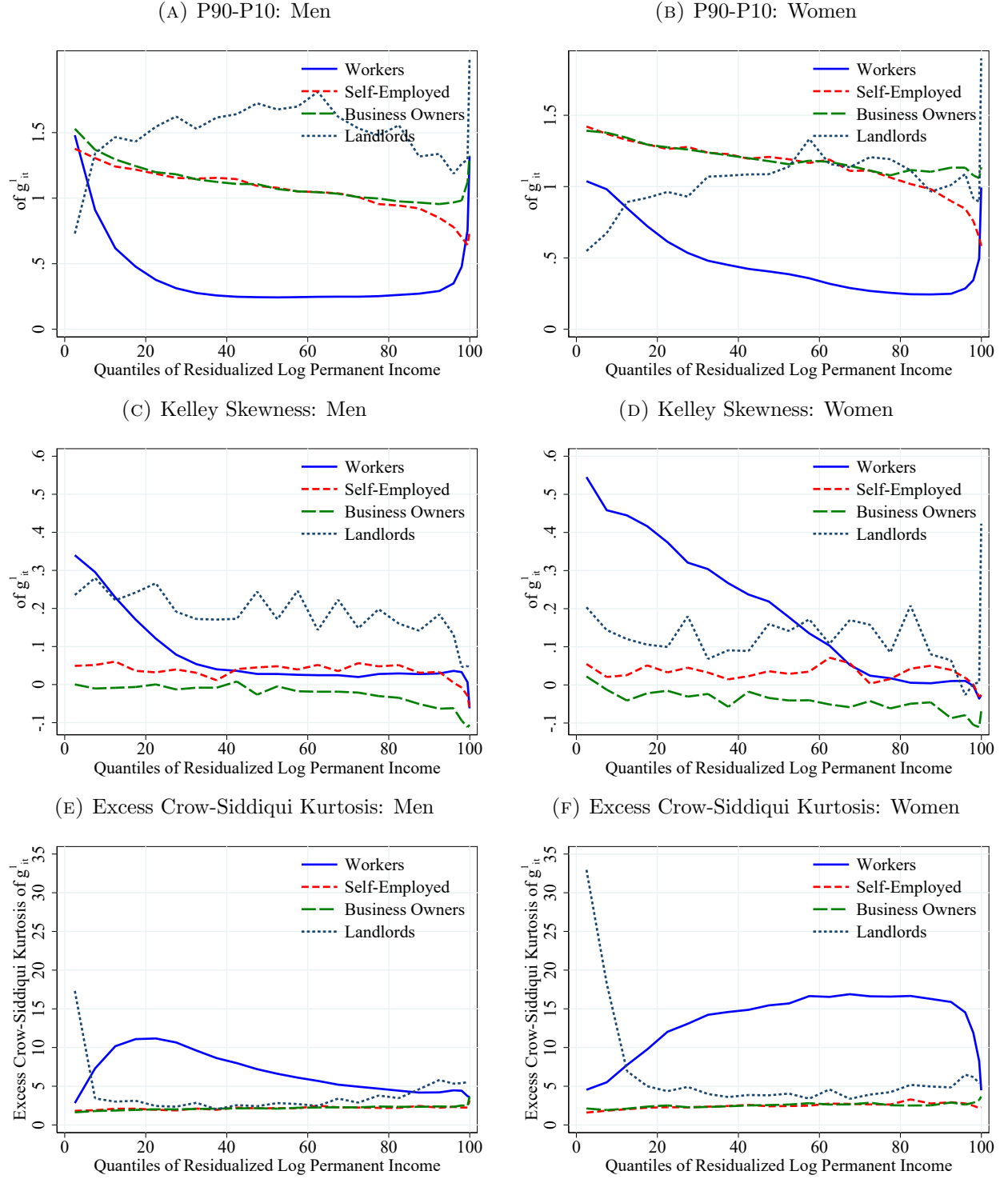
Notes: This figure shows the evolution of the P90-P10 differential, Kelley skewness and excess Crow-Siddiqui kurtosis of 1-year changes in residualized log real annual total income (from $t - 1$ to t) in the combined IAB-TPP data (LS sample) separately for men and women by main income source (workers, self-employment, business owners, landlords). See Footnote 24 definitions and interpretation of Kelley skewness and excess Crow-Siddiqui kurtosis. Shaded areas indicate recessions.

FIGURE G.11: HETEROGENEITY IN DISPERSION, SKEWNESS AND KURTOSIS OF 1-YEAR LOG INCOME GROWTH BY MAIN INCOME SOURCE



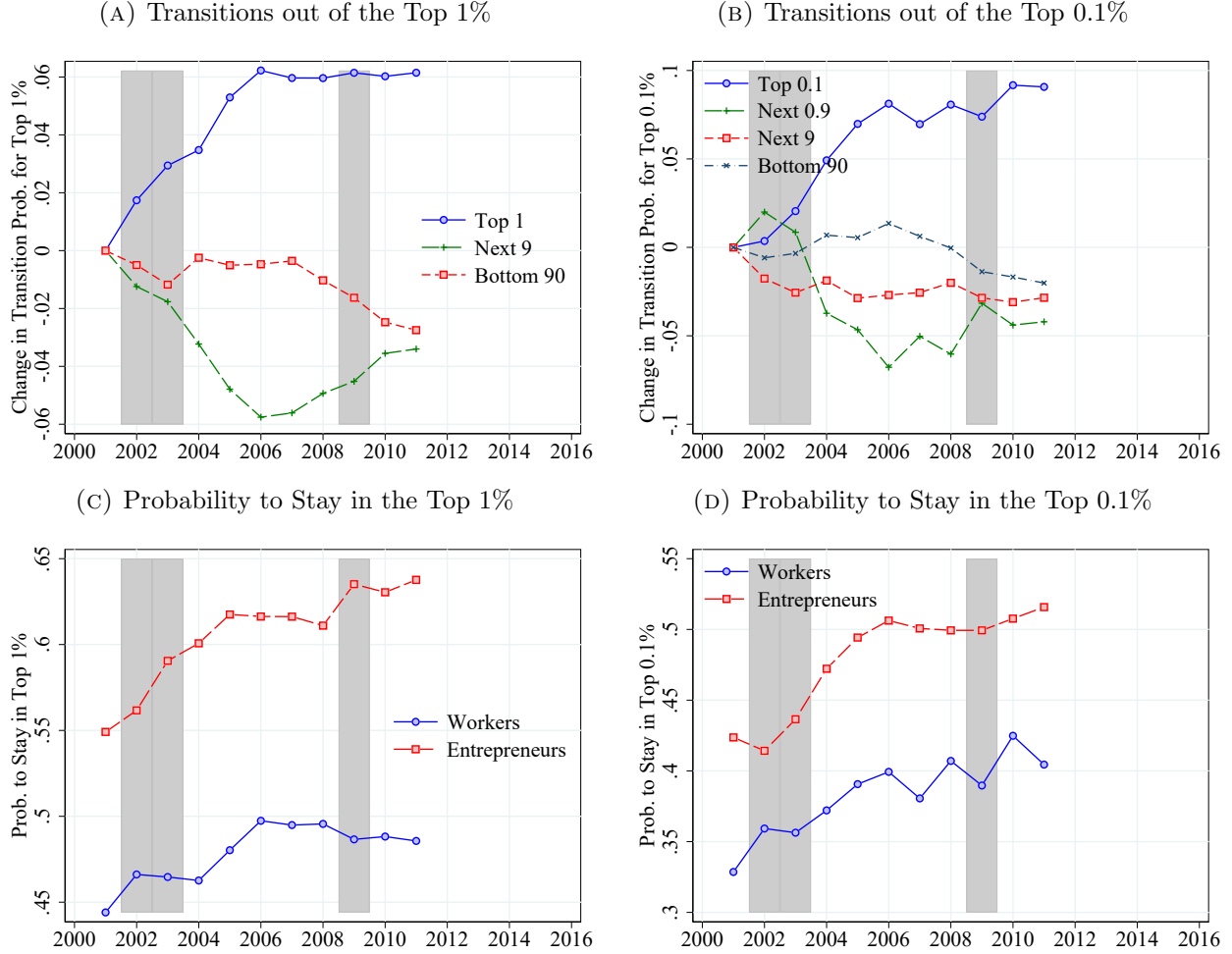
Notes: This figure shows the P90-P10 differential, Kelley skewness and excess Crow-Siddiqui kurtosis of 1-year changes in residualized log real total income by permanent total income (from $t - 1$ to t) in the combined IAB-TTP data (H Sample) as averages from 2004 to 2011 and separately for men and women by main income source (workers, self-employment, business owners, landlords). The horizontal axis plots the exponential of mean permanent income in 1,000 Euro. See Footnote 24 definitions and interpretation of Kelley skewness and excess Crow-Siddiqui kurtosis. Shaded areas indicate recessions.

FIGURE G.12: HETEROGENEITY IN DISPERSION, SKEWNESS AND KURTOSIS OF 1-YEAR LOG INCOME GROWTH BY MAIN INCOME SOURCE



Notes: This figure shows the P90-P10 differential, Kelley skewness and excess Crow-Siddiqui kurtosis of 1-year changes in residualized log real total income by quantiles of the distribution of permanent total income (from $t - 1$ to t) in the combined IAB-TPP data (H Sample) as averages from 2004 to 2011 and separately for men and women by main income source (workers, self-employment, business owners, landlords). The (gender-specific) ranking of permanent income is based on the distribution of total income of all taxpayers. See Footnote 24 definitions and interpretation of Kelley skewness and excess Crow-Siddiqui kurtosis. Shaded areas indicate recessions.

FIGURE G.13: TOP INCOME MOBILITY – 5-YEAR TRANSITION PROBABILITIES



Notes: This figure plots transition probabilities from top income using the combined IAB-TPP data (LS sample). Panels A and B show the evolution of 5-year transition probabilities out of the top 1% and top 0.1% of the income distribution into selected parts of the income distribution from one year to the next. The “Next 9” is the part of the distribution between the P90 and P99 and the “Next 0.9” is the part between the P99 and the P99.9. The lines sum to zero. Panels C and D show the 5-year probability of staying in the top 1% or top 0.1% for workers and entrepreneurs. The ranking is based on the total income distribution and not conditional on the main income source. Shaded areas indicate recessions.

References

- BRADLEY, J. AND A. KÜGLER (2019): “Labor market reforms: An evaluation of the Hartz policies in Germany,” *European Economic Review*, 113, 108–135.
- BRENKE, K., U. RINNE, AND K. F. ZIMMERMANN (2013): “Short-time work: The German answer to the Great Recession,” *International Labour Review*, 152, 287–305.
- CARD, D., J. HEINING, AND P. KLINE (2013): “Workplace Heterogeneity and the Rise of West German Wage Inequality,” *The Quarterly Journal of Economics*, 128, 967–1015.
- DAUTH, W. AND J. EPPELSHEIMER (2020): “Preparing the sample of integrated labour market biographies (SIAB) for scientific analysis: a guide,” *Journal for Labour Market Research*, 54.

- DiNARDO, J., N. FORTIN, AND T. LEMIEUX (1996): “Labor Market Institution and the Distribution of Wages,” *Econometrica*, 1001–1044.
- DOERRENBERG, P., A. PEICHL, AND S. SIEGLOCH (2017): “The elasticity of taxable income in the presence of deduction possibilities,” *Journal of Public Economics*, 151, 41–55.
- DOLLS, M., P. DOERRENBERG, A. PEICHL, AND H. STICHNOTH (2018): “Do retirement savings increase in response to information about retirement and expected pensions?” *Journal of Public Economics*, 158, 168–179.
- DUSTMANN, C., B. FITZENBERGER, U. SCHÖNBERG, AND A. SPITZ-OENER (2014): “From Sick Man of Europe to Economic Superstar: Germany’s Resurgent Economy,” *Journal of Economic Perspectives*, 28, 167–88.
- DUSTMANN, C., A. LINDNER, U. SCHÖNBERG, M. UMKEHRER, AND P. VOM BERGE (2022): “Reallocation Effects of the Minimum Wage,” *The Quarterly Journal of Economics*, 137, 267–328.
- ELLGUTH, P. AND S. KOHAUT (2019): “Tarifbindung und betriebliche Interessenvertretung. Aktuelle Ergebnisse aus dem IAB-Betriebspanel 2018,” *WSI-Mitteilungen*, 72, 290–297.
- (2020): “Tarifbindung und betriebliche Interessenvertretung: Aktuelle Ergebnisse aus dem IAB-Betriebspanel 2019,” *WSI-Mitteilungen*, 73.
- FITZENBERGER, B. AND A. SEIDLITZ (2020): “The 2011 break in the part-time indicator and the evolution of wage inequality in Germany,” *Journal for Labour Market Research*, 54.
- FRODERMANN, C., A. SCHMUCKER, S. SETH, AND P. VOM BERGE (2021): “Sample of Integrated Labour Market Biographies (SIAB) 1975-2019,” FDZ-Datenreport 01, Nuremberg.
- GUDGEON, M. AND S. TRENKLE (2020): “The speed of earnings responses to taxation and the role of firm labor demand,” IZA Discussion Paper 13931.
- HARTUNG, B., P. JUNG, AND M. KUHN (2018): “What Hides Behind the German Labor Market Miracle? Unemployment Insurance Reforms and Labor Market Dynamics,” CESifo Working Paper Series 7379, CESifo.
- HOCHMUTH, B., B. KOHLBRECHER, C. MERKL, AND H. GARTNER (2021): “Hartz IV and the decline of German unemployment: A macroeconomic evaluation,” *Journal of Economic Dynamics and Control*, 127, 104114.
- HOHENDANNER, C. AND J. STEGMAIER (2012): “Geringfügig Beschäftigte in deutschen Betrieben: Umstrittene Minijobs,” *IAB-Kurzbericht*.
- KREBS, T. AND M. SCHEFFEL (2013): “Macroeconomic Evaluation of Labor Market Reform in Germany,” *IMF Economic Review*, 61, 664–701.
- (2017): “Labor Market Institutions and the Cost of Recessions,” IMF Working Papers 2017/087, International Monetary Fund.
- KRIETE-DODDS, S. AND D. VORGRIMLER (2007): “The German Taxpayer-Panel,” *Schmollers Jahrbuch*, 127, 497–509.

- KROLAGE, C., A. PEICHL, AND D. WALDENSTRÖM (forthcoming): “Long Run Trends in Top Income Shares: the Role of Income and Population Growth,” *Journal of Economic Inequality*.
- LAUNOV, A. AND K. WÄLDE (2013): “Estimating Incentive and Welfare Effects of Nonstationary Unemployment Benefits,” *International Economic Review*, 54, 1159–1198.
- RIEBESELL, P. (1922): *Steuer-Mathematik. Die Fehler in den Reichssteuertarifen.*, Hamburg: Henri Grand.
- THOMSEN, U., J. LUDSTECK, AND A. SCHMUCKER (2018): “Skilled or unskilled – Improving the information on qualification for employee data in the IAB Employee Biography,” FDZ-Methodenreport 09, Nuremberg.